

Probability & Computing

Continuous Probability Spaces & Random Geometric Graphs

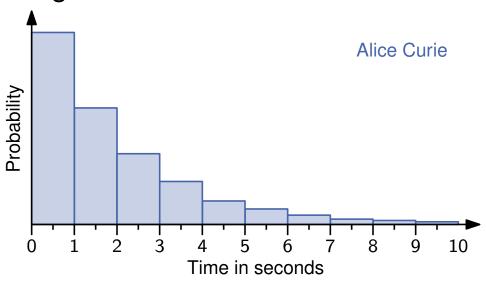




- Two physicists study radioactive material that emits particles every now and then
- Both compete to get the most accurate model describing the emission

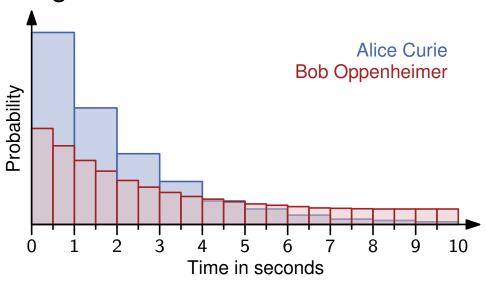


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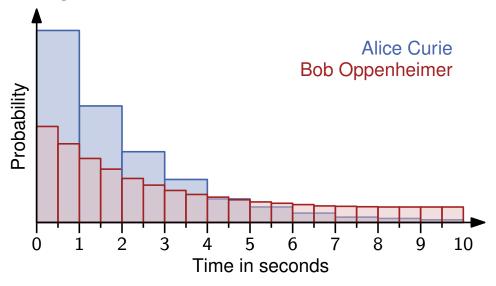


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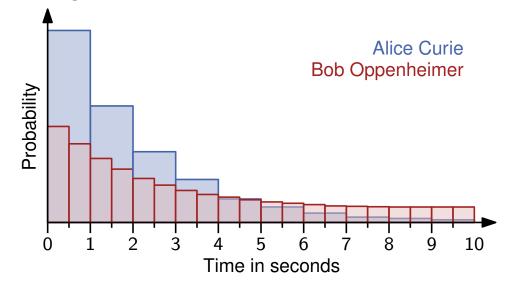


- Two physicists study radioactive material that emits particles every now and then
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- "We could do this forever!" Could they really?





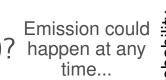
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- "We could do this forever!" Could they really?
- They measure with infinite precision...
 - What is Pr[X = 2.71828182846]?
 - What is Pr[X = 2.71828182847]?

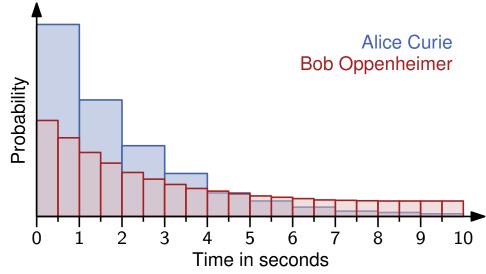




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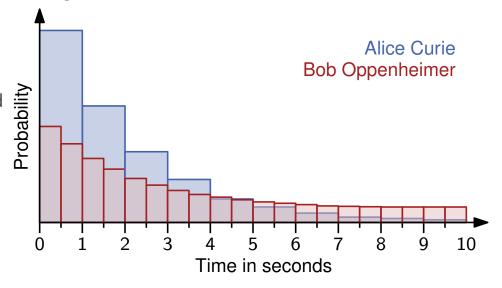






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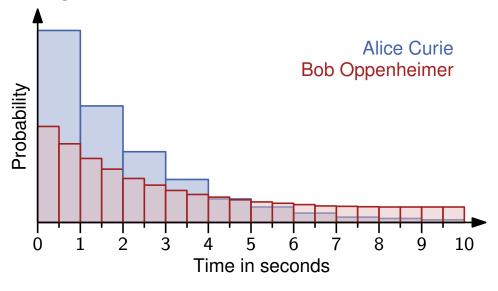
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- For continuous spaces we need to adjust how we measure probabilities



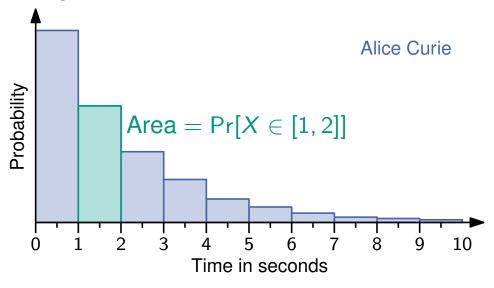


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We assign probabilities to *intervals* instead of individual values! The probability is the *area* of the bar, *not* the height

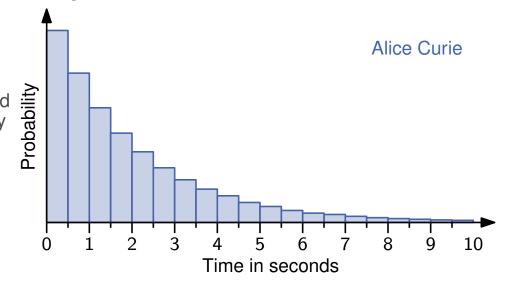




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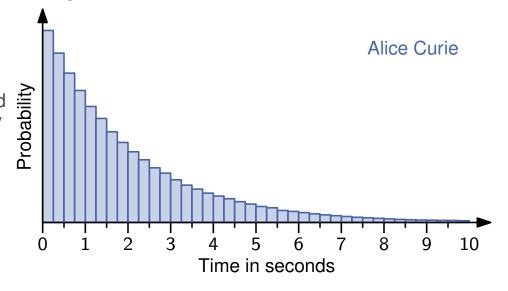




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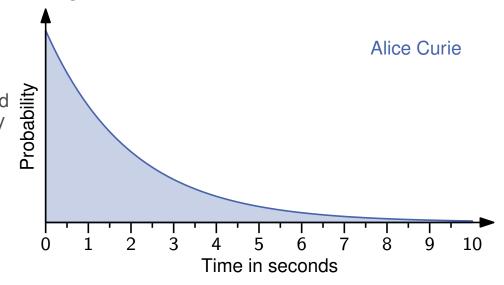




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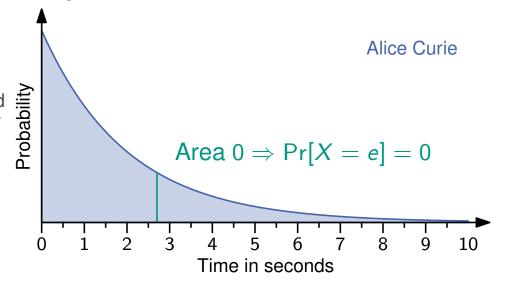




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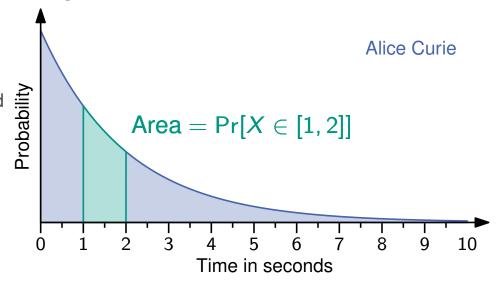




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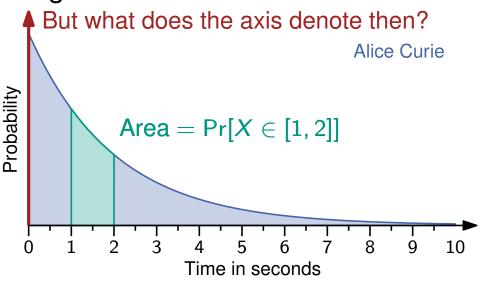


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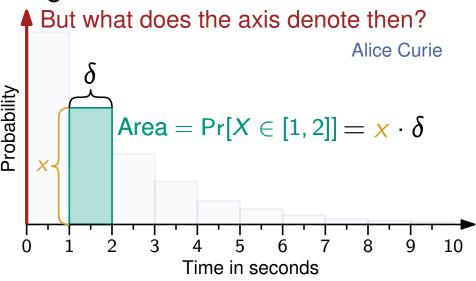




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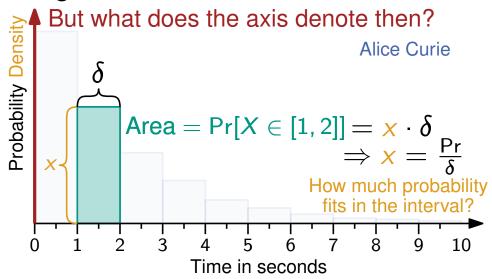




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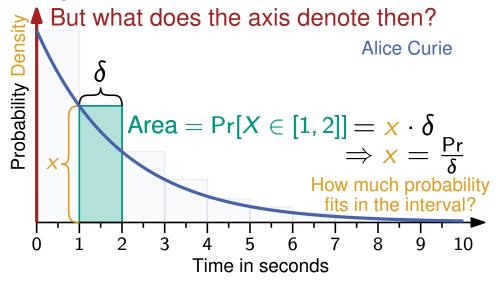
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- As bars get thinner, areas (probabilities) decrease
- We describe distributions using probability density functions



youtube.com/watch?v=ZA4JkHKZM50



Discrete Random Variable X

Continuous Random Variable *X*



Discrete Random Variable *X*

Cumulative distribution function

$$F_X(x) = \Pr[X \leq x]$$

Continuous Random Variable *X*



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Continuous Random Variable *X*

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Discrete Random Variable X

Cumulative distribution function

$$F_X(x) = \Pr[X \leq x]$$

Probability mass function

$$f_X(x) = \Pr[X = x] \ge 0$$

Continuous Random Variable *X*

$$F_X(x) = \Pr[X \leq x]$$



Discrete Random Variable X

Cumulative distribution function

$$F_X(x) = \Pr[X \leq x]$$

Probability mass function

$$f_X(x) = \Pr[X = x] \ge 0$$

$$\sum_{x} \Pr[X = x] = 1$$

Continuous Random Variable *X*

$$F_X(x) = \Pr[X \leq x]$$



Discrete Random Variable X

Cumulative distribution function

$$F_X(x) = \Pr[X \leq x]$$

Probability mass function

$$f_X(x) = \Pr[X = x] \ge 0$$

$$\sum_{x} \Pr[X=x]=1$$
 $f_X(x) \geq 0$

Continuous Random Variable X

Cumulative distribution function

$$F_X(x) = \Pr[X \leq x]$$

Probability density function

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Probability mass function
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Expectation

$$\mathbb{E}[X] = \sum_{x} x \cdot \Pr[X = x]$$

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Example: Uniform Distribution

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Example: Uniform Distribution

- You build a fence that is at least 2m tall at each point
- In the hardware store they have 5m planks



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- You build a fence that is at least 2m tall at each point
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→ Over [0, 5]



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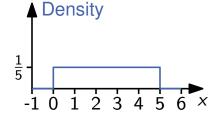
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 $\sum_{x} \Pr[X = x] = 1$

Expectation

$$\mathbb{E}[X] = \sum_{x} x \cdot \Pr[X = x]$$

Continuous Random Variable *X*

Cumulative distribution function

$$F_X(x) = \Pr[X \le x] = \int_{-\infty}^x f_X(y) dy$$

Probability density function •

$$f_X(x) \ge 0$$

$$\int_{-\infty}^{\infty} f_X(x) dx = 1$$

$$\int_{-\infty}^{\infty} f_X(x) \mathrm{d}x = 1$$

Expectation

$$\mathbb{E}[X] = \int x \cdot f_X(x) \mathrm{d}x$$

Example: Uniform Distribution -

- You build a fence that is at least 2m tall at each point
- In the hardware store they have 5m planks
- The staff member cutting your planks wears hearing protection and cuts uniformly at random
- \blacksquare What is the probability that you get two $\geq 2m$ boards out of one 5m plank?

$$f_X(x) = \begin{cases} \frac{1}{5}, & \text{if } x \in [0, 5] \\ 0, & \text{o.w.} \end{cases}$$

$$\int_{-\infty}^{\infty} f_X(x) dx = \int_0^5 \frac{1}{5} dx = \left[\frac{x}{5}\right]_0^5 = 1 \checkmark$$



Discrete Random Variable X

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Over
$$[0, 5]$$

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$$\int_{-\infty}^{\infty} f_X(x) dx = \int_{0}^{5} \frac{1}{5} dx = \left[\frac{x}{5}\right]_{0}^{5} = 1 \checkmark$$

$$\int_{a}^{b} f_{X}(x) dx = \left[\frac{x}{5}\right]_{a}^{b} = \frac{1}{5}(b-a) \checkmark$$
for $a \le b \in [0, 5]$



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Over
$$[0, 5]$$

$$f_X(x) = \begin{cases} \frac{1}{5}, & \text{if } x \in [0, 5] \\ 0, & \text{o.w.} \end{cases}$$
For all $x \in [0, 5]$

$$\int_{-\infty}^{\infty} f_X(x) dx = \int_0^5 \frac{1}{5} dx = \left[\frac{x}{5}\right]_0^5 = 1 \checkmark$$

$$\int_a^b f_X(x) dx = \left[\frac{x}{5}\right]_a^b = \frac{1}{5}(b-a) \checkmark$$
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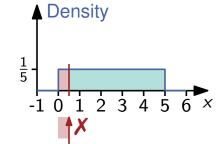
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Example: Uniform Distribution

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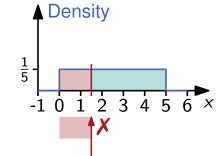
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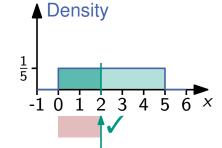
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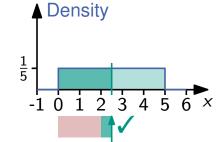
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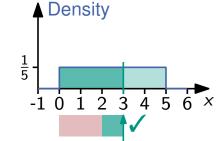
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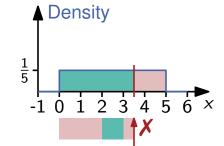
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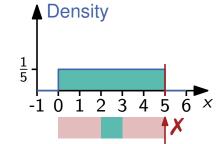
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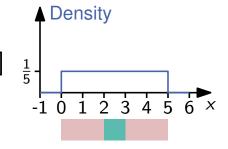
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$$ightharpoonup \Pr[X \in [2, 3]]$$





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Over
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Density
$$\frac{1}{5}$$

→
$$\Pr[X \in [2, 3]] = \Pr[X \le 3] - \Pr[X < 2]$$



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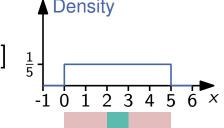
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$$= \left[\frac{x}{5}\right]_0^3 - \left[\frac{x}{5}\right]_0^2 = \frac{3}{5} - \frac{2}{5} = \frac{1}{5} \checkmark$$



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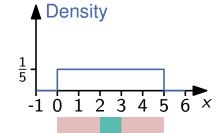
$$\mathbb{E}[X] = \int x \cdot f_X(x) \mathrm{d}x$$

Example: Uniform Distribution

- You build a fence that is at least 2m tall at each point
- In the hardware store they have 5m planks
- The staff member cutting your planks wears hearing protection and cuts uniformly at random
- What is the probability that you get two ≥ 2m boards out of one 5m plank?

→ Over [0, 5]

$$f_X(x) = \begin{cases} \frac{1}{5}, & \text{if } x \in [0, 5] \\ 0, & \text{o.w.} \end{cases}$$



■ In general: $X \sim \mathcal{U}([a, b])$

$$\Pr[X \in [c, d] \subseteq [a, b]] = \frac{d-c}{b-a}$$



Exponential Distribution $X \sim Exp(\lambda)$

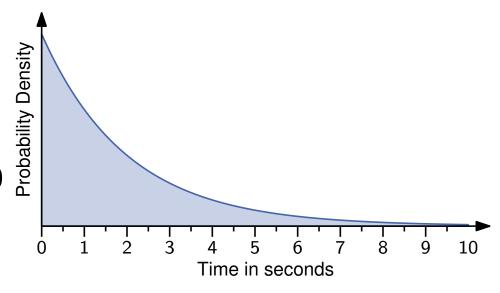
• "Rate" parameter $\lambda > 0$



- "Rate" parameter $\lambda > 0$
- Continuous equivalent to geometric distribution
- "Time until first success"



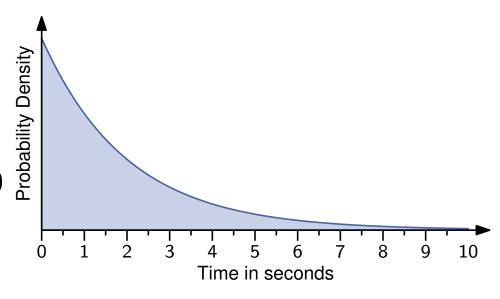
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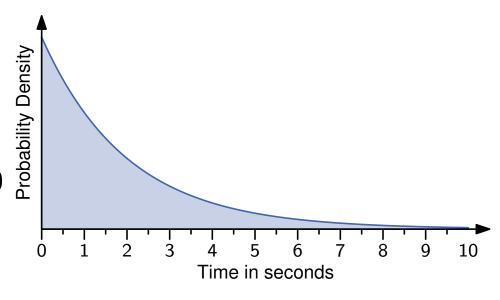
$$F_X(x) = \int_{-\infty}^x f_X(y) \mathrm{d}y$$





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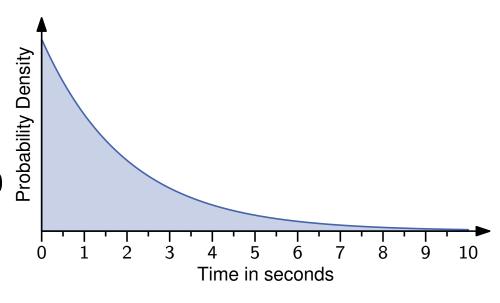
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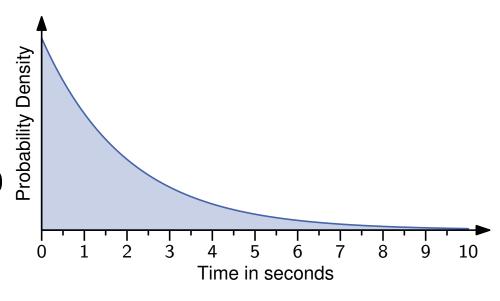
$$F_X(x) = \int_{-\infty}^x f_X(y) dy = \lambda \int_0^x e^{-\lambda y} dy$$
$$= \frac{\lambda}{-\lambda} \left[e^{-\lambda y} \right]_0^x$$





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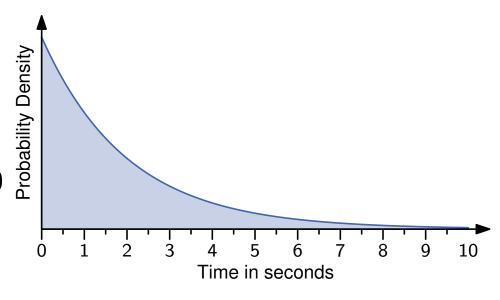
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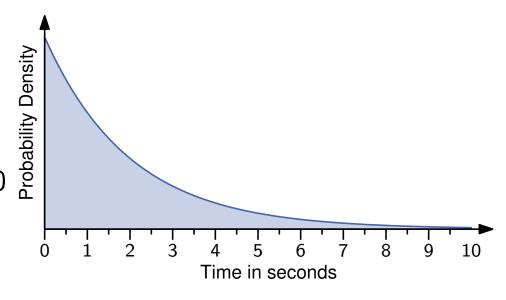
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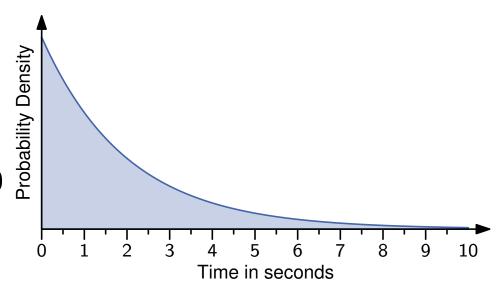




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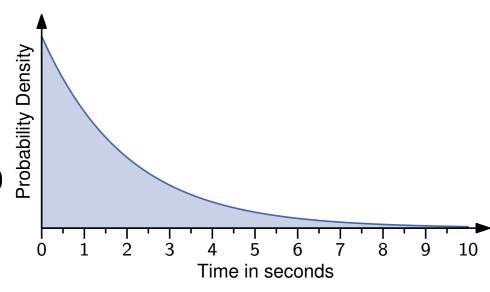


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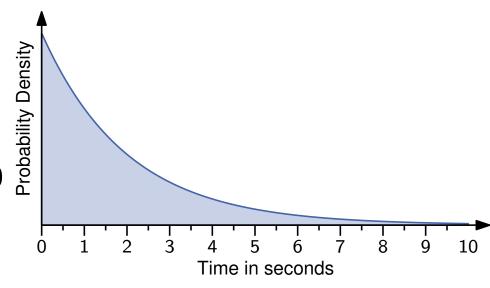


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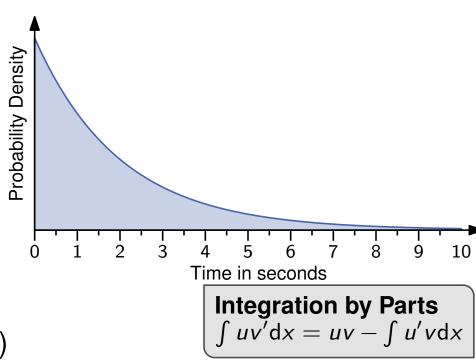


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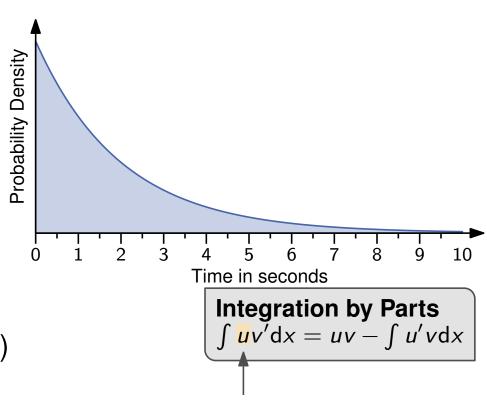


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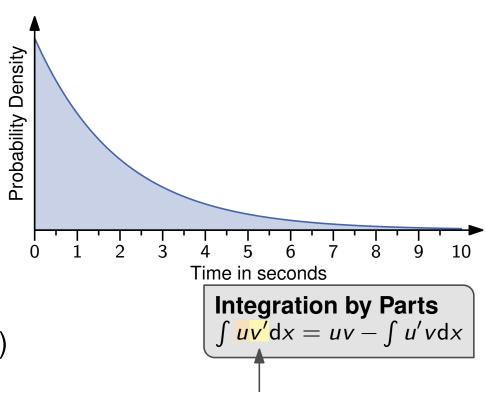
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$$u = x$$

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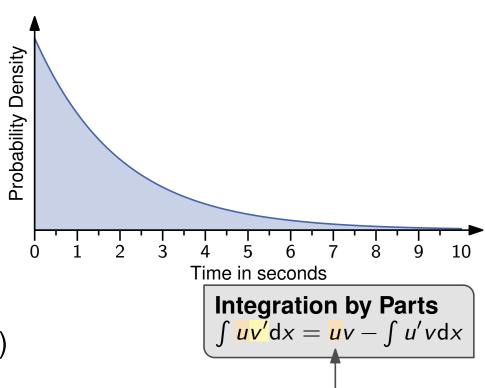
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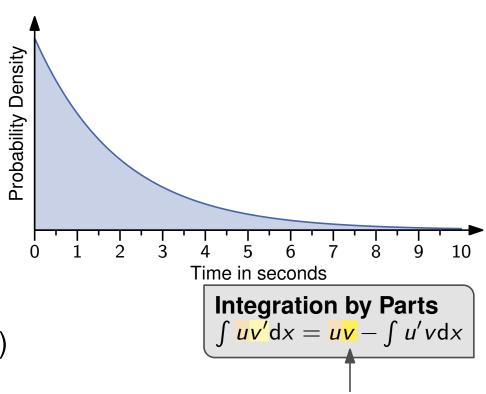


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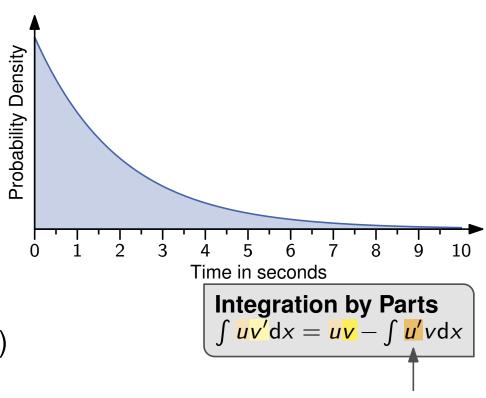


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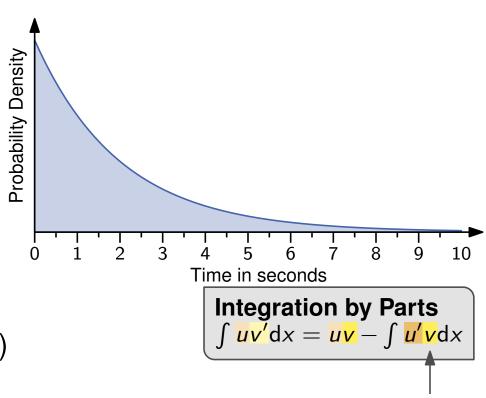


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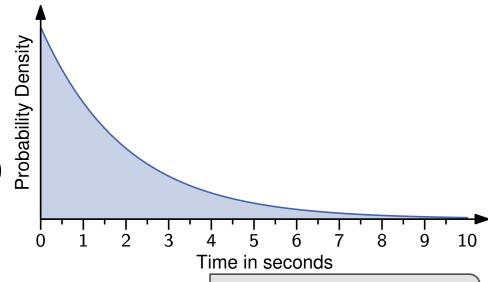
Maximilian Katzmann, Stefan Walzer - Probability & Computing

Characterization via Moments (*n*-th moment: $\mathbb{E}[X^n]$)

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$$u = x \quad v = \frac{1}{-\lambda} e^{-\lambda x}$$

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Integration by Parts $\int \frac{uv'}{dx} dx = \frac{uv}{uv'} - \int \frac{u'v}{u'} dx$



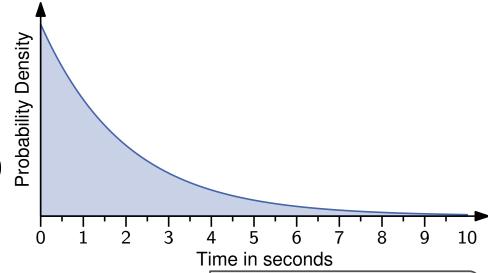
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Integration by Parts $\int uv'dx = uv - \int u'vdx$

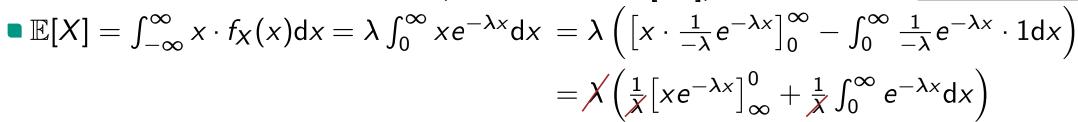


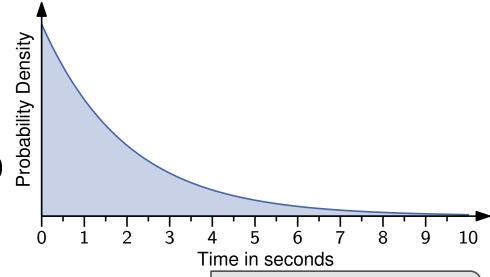
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Integration by Parts $\int uv' dx = uv - \int u'v dx$



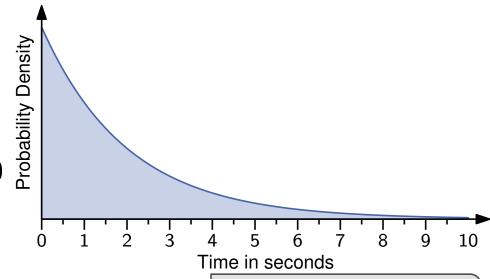
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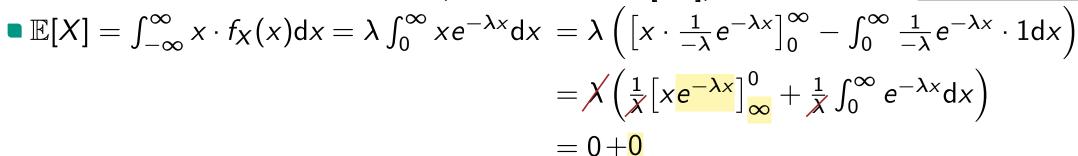


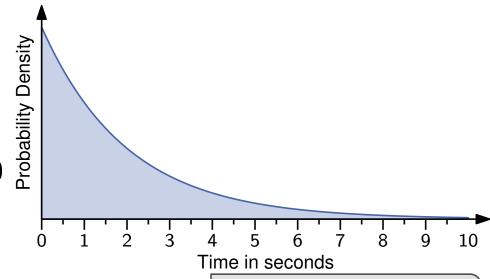
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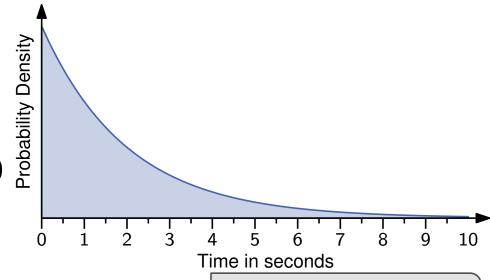
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$$= 0 + 0$$



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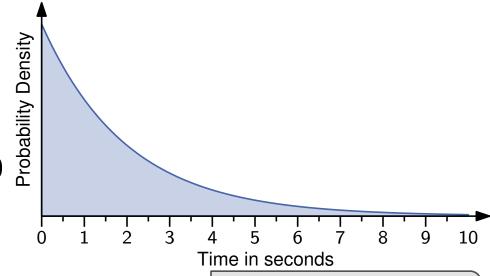
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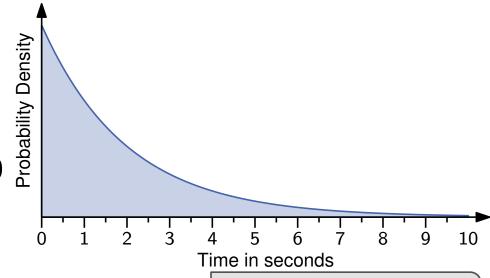
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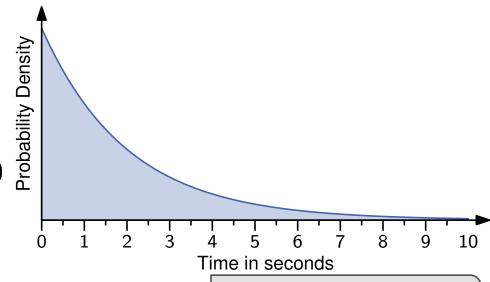
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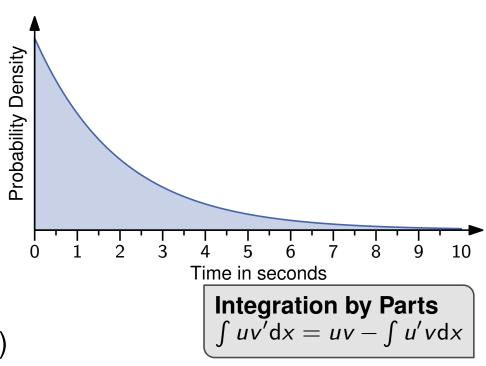


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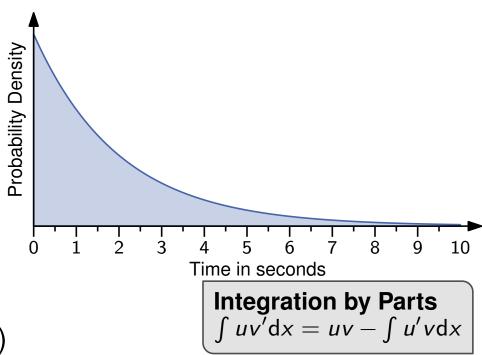
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$$\blacksquare \mathbb{E}[X^2] = \int_{-\infty}^{\infty} x^2 \cdot f_X(x) \mathrm{d}x$$





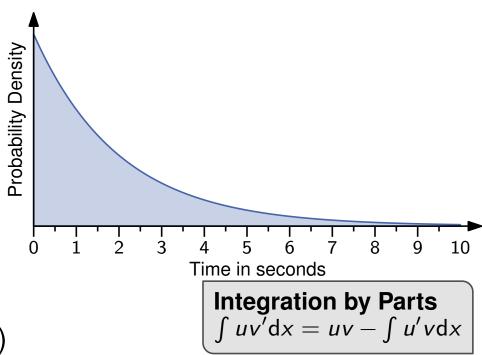
Exponential Distribution $X \sim Exp(\lambda)$

- "Rate" parameter $\lambda > 0$
- Continuous equivalent to geometric distribution
- "Time until first success"
- Probability density function $f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \geq 0 \\ 0, & \text{o.w.} \end{cases}$
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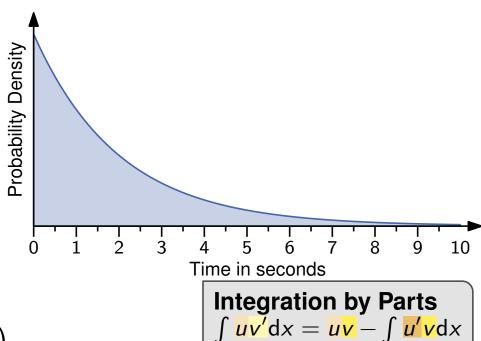
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$$u = x^2$$
 $v = \frac{1}{-\lambda}e^{-\lambda x}$
 $u' = 2x$ $v' = e^{-\lambda x}$



Exponential Distribution $X \sim Exp(\lambda)$

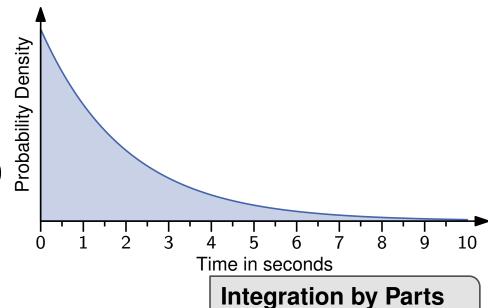
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$$= \lambda \left(\left[x^2 - \frac{1}{\lambda} e^{-\lambda x} \right]_0^{\infty} \right)$$



$$= \lambda \int_0^\infty x^2 e^{-\lambda x} dx$$

$$= \lambda \left(\left[x^2 \frac{1}{-\lambda} e^{-\lambda x} \right]_0^\infty - \frac{2}{-\lambda} \int_0^\infty x \cdot e^{-\lambda x} dx \right) \quad u = x^2 \quad v = \frac{1}{-\lambda} e^{-\lambda x}$$

$$= \lambda \left(\left[x^2 \frac{1}{-\lambda} e^{-\lambda x} \right]_0^\infty - \frac{2}{-\lambda} \int_0^\infty x \cdot e^{-\lambda x} dx \right) \quad u' = 2x \quad v' = e^{-\lambda x}$$

 $\int \frac{\mathbf{u}\mathbf{v}'}{\mathbf{d}x} = \frac{\mathbf{u}\mathbf{v}}{\mathbf{v}} - \int \frac{\mathbf{u}'\mathbf{v}}{\mathbf{d}x}$



Exponential Distribution $X \sim Exp(\lambda)$

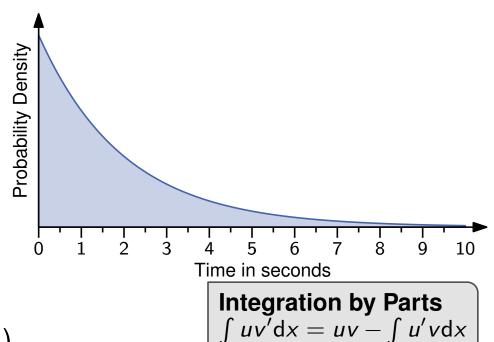
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Exponential Distribution $X \sim Exp(\lambda)$

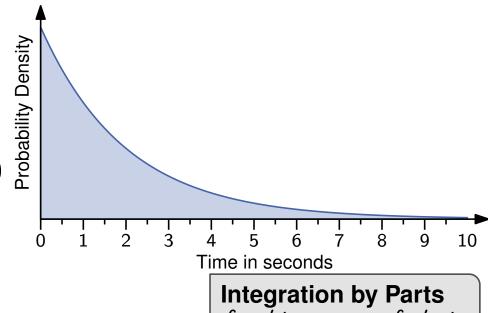
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Integration by Parts $\int uv' dx = uv - \int u'v dx$

$$= \lambda \int_0^\infty x^2 e^{-\lambda x} dx$$

$$= \lambda \left(\left[x^2 \frac{1}{-\lambda} e^{-\lambda x} \right]_0^\infty - \frac{2}{-\lambda} \int_0^\infty x \cdot e^{-\lambda x} dx \right) = \lambda ([0+0] + \frac{2}{\lambda^3})$$



Exponential Distribution $X \sim Exp(\lambda)$

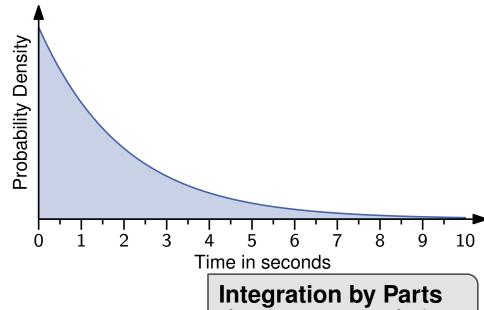
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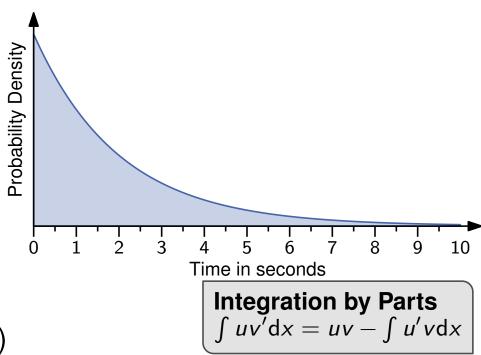
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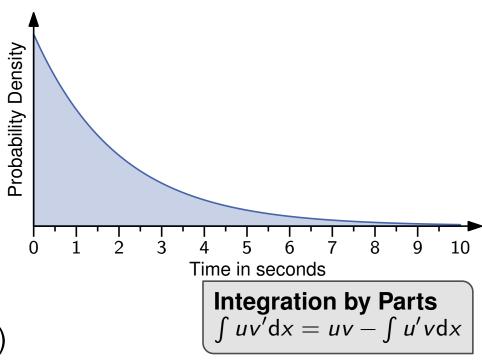
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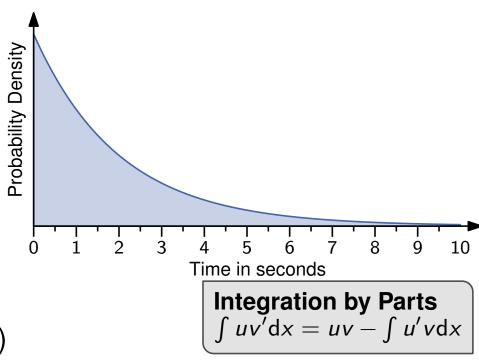
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$$Var[X] = \frac{\mathbb{E}[X^2]}{\mathbb{E}[X]^2} = \frac{2}{\lambda^2} - (\frac{1}{\lambda})^2$$





Exponential Distribution $X \sim Exp(\lambda)$

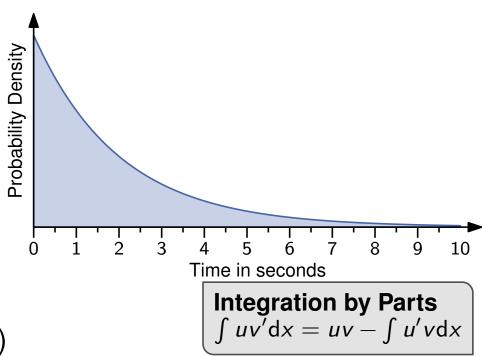
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•
$$Var[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{2}{\lambda^2} - (\frac{1}{\lambda})^2 = \frac{1}{\lambda^2}$$







$$f_X(x) = \lambda e^{-\lambda x}$$
 $f_X(x) = 1 - e^{-\lambda x}$





$$\Pr[X > s + t \mid X > t]$$

$$f_X(x) = \lambda e^{-\lambda x}$$
 $f_X(x) = 1 - e^{-\lambda x}$





$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]}$$

$$X \sim \text{Exp}(\lambda)$$

 $f_X(x) = \lambda e^{-\lambda x}$
 $F_X(x) = 1 - e^{-\lambda x}$

$$\Pr[A \mid B] = \frac{\Pr[A \land B]}{\Pr[B]}$$





$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \quad \begin{array}{c} X > s + t \Rightarrow X > t \\ \end{array}$$

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$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \quad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline \\ \Pr[X > t] \end{array}$$

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 $f_X(x) = 1 - e^{-\lambda x}$

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$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \xrightarrow{X > s + t \Rightarrow X > t}$$

$$= \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]}$$

$$f_X(x) = \lambda e^{-\lambda x}$$
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$$= \frac{e^{-\lambda(s+t)}}{e^{-\lambda(s+t)}}$$

$$f_X(x) = \lambda e^{-\lambda x}$$
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$$\Pr[A \mid B] = \frac{\Pr[A \land B]}{\Pr[B]}$$





$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \quad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline Pr[X > t] \end{array}$$

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■ What is the probability of having to wait longer than an additional time s > 0 after already having waited time t > 0?

$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \quad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline Pr[X > t] \end{array}$$

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$$\Pr[A \mid B] = \frac{\Pr[A \land B]}{\Pr[B]}$$

No matter how long we already waited, waiting time is distributed as if we just started

Exponential Distribution: Memorylessness



Motivation

• What is the probability of having to wait longer than an additional time s>0 after already having waited time t > 0?

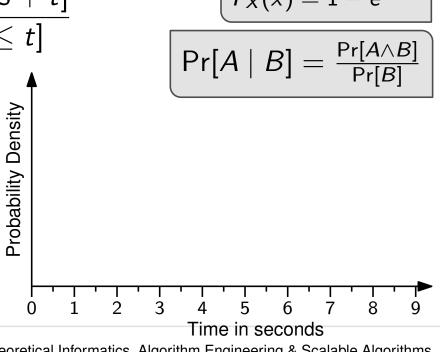
$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \qquad \begin{array}{l} X > s + t \Rightarrow X > t \\ \\ = \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]} \\ \\ = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \\ \\ = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \end{array}$$
No matter how long we already waited, waiting time

No matter how long we already waited, waiting time is distributed as if we just started

Observing Multiple Particles

Maximilian Katzmann, Stefan Walzer - Probability & Computing

How long do we have to wait for the second particle after having just seen the first?





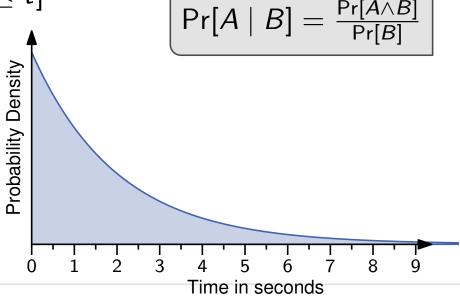
Motivation

• What is the probability of having to wait longer than an additional time s>0 after already having waited time t > 0?

$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \qquad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline Pr[X > t] \\ \hline = \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]} \\ \hline = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \\ \hline \text{No matter how long we already waited, waiting time is distributed as if we just started} \\ \hline \textbf{bserving Multiple Particles} \\ \hline \text{How long do we have to wait for the second} \end{array}$$

No matter how long we already waited, waiting time is distributed as if we just started

Observing Multiple Particles





Motivation

• What is the probability of having to wait longer than an additional time s>0 after already having waited time t > 0?

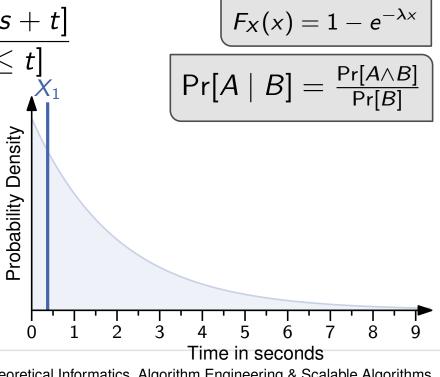
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$$= \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]}$$

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No matter how long we already waited, waiting time
$$= \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s]$$
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No matter how long we already waited, waiting time is distributed as if we just started

Observing Multiple Particles





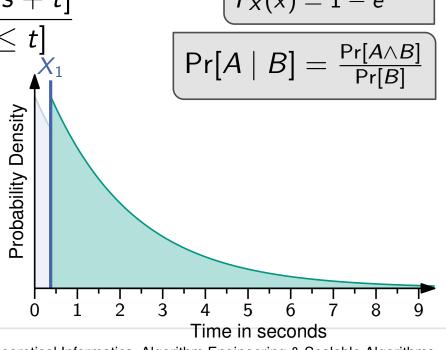
Motivation

• What is the probability of having to wait longer than an additional time s>0 after already having waited time t > 0?

$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \qquad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline Pr[X > s + t \mid X > t] \\ \hline = \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]} \\ \hline = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \\ \hline \text{No matter how long we already waited, waiting time} \end{array}$$

No matter how long we already waited, waiting time is distributed as if we just started

Observing Multiple Particles





Motivation

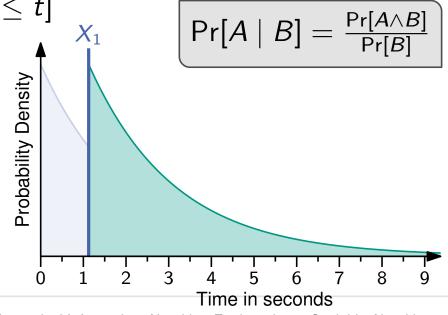
• What is the probability of having to wait longer than an additional time s>0 after already having waited time t > 0?

$$\Pr[X > s + t \mid X > t] = \frac{\Pr[X > s + t \land X > t]}{\Pr[X > t]} \qquad \begin{array}{l} X > s + t \Rightarrow X > t \\ \hline Pr[X > t] \\ \hline = \frac{\Pr[X > s + t]}{\Pr[X > t]} = \frac{1 - \Pr[X \le s + t]}{1 - \Pr[X \le t]} \\ \hline = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \\ \hline = \frac{e^{-\lambda(s + t)}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s] \end{array}$$
No matter how long we already waited, waiting time is distributed as if we just started
$$\frac{e^{-\lambda t}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s]$$

$$\frac{e^{-\lambda t}}{e^{-\lambda t}} = e^{-\lambda s} = \Pr[X > s]$$
How long do we have to wait for the second

No matter how long we already waited, waiting time is distributed as if we just started

Observing Multiple Particles





 $X \sim \operatorname{Exp}(\lambda)$

Motivation

■ What is the probability of having to wait longer than an additional time s > 0 after already having waited time t > 0?

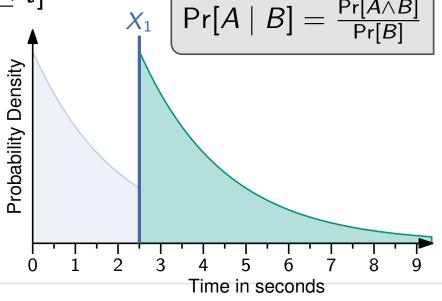
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No matter how long we already waited, waiting time is distributed as if we just started

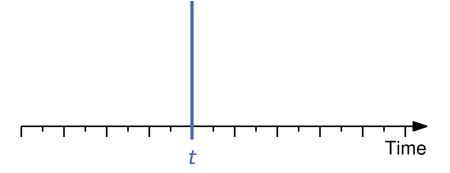
Observing Multiple Particles





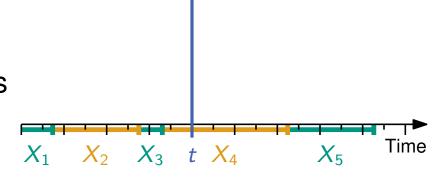
Motivation

Count number of particles emitted within a given time t



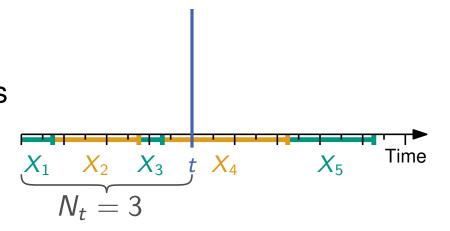


- Count number of particles emitted within a given time t
- Let $X_1, X_2, X_3, ... \sim \text{Exp}(\lambda)$ be independent waiting times





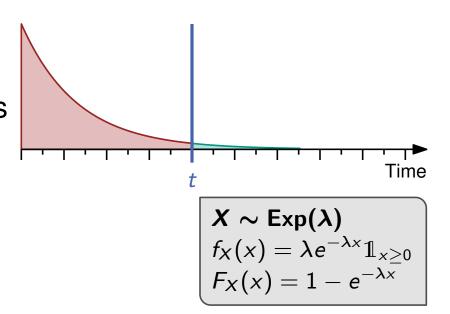
- Count number of particles emitted within a given time t
- Let $X_1, X_2, X_3, ... \sim \text{Exp}(\lambda)$ be independent waiting times
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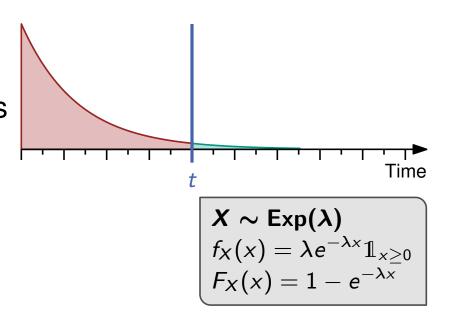
$$Pr[N_t = 0]$$





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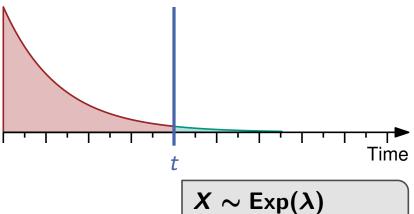
$$Pr[N_t = 0] = Pr[X_1 > t]$$





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$$\Pr[N_t = 0] = \Pr[X_1 > t] = 1 - \Pr[X_1 \le t] = 1 - F_{X_1}(t)$$



$$f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x \ge 0}$$

 $F_X(x) = 1 - e^{-\lambda x}$

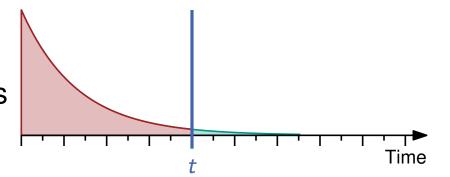


Motivation

- Count number of particles emitted within a given time t
- Let $X_1, X_2, X_3, ... \sim \text{Exp}(\lambda)$ be independent waiting times
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Specific Values

$$\Pr[N_t = 0] = \Pr[X_1 > t] = 1 - \Pr[X_1 \le t] = 1 - F_{X_1}(t) = e^{-\lambda t}$$



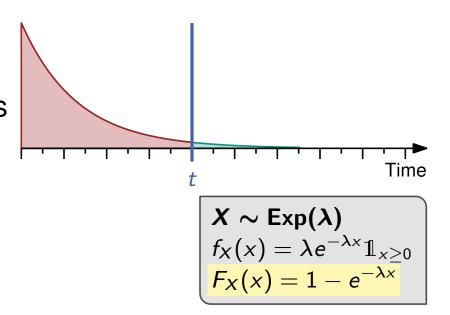
$$X \sim \text{Exp}(\lambda)$$

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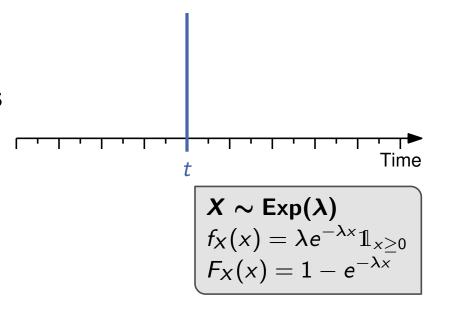




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$$\Pr[N_t = 0] = e^{-\lambda t}$$

 $\Pr[N_t = 1]$

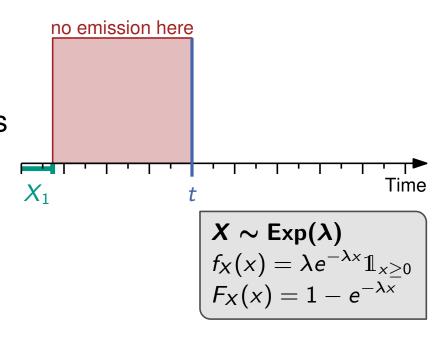




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 $\Pr[N_t = 1] = \Pr[X_1 \le t \land N(X_1, t) = 0]$

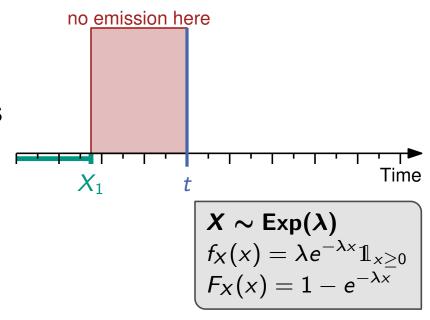




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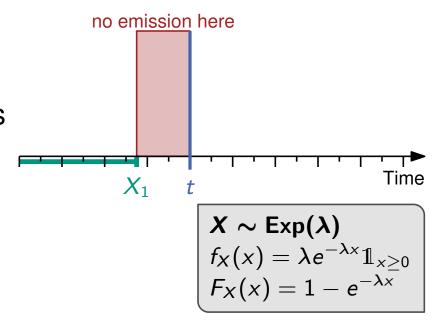




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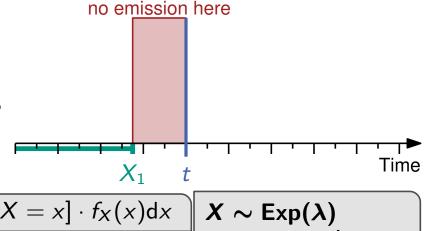


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$$\Pr[A] = \int_{-\infty}^{\infty} \Pr[A \mid X = x] \cdot f_X(x) dx$$
 $X \sim \operatorname{Exp}(\lambda)$

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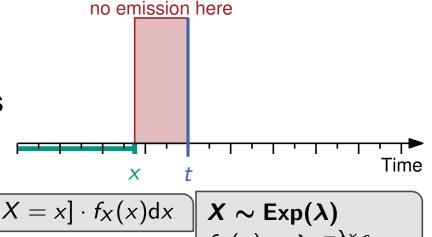


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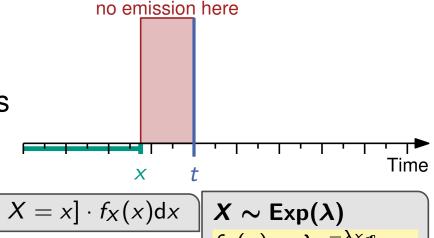


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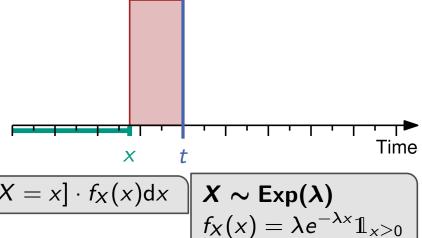


Motivation

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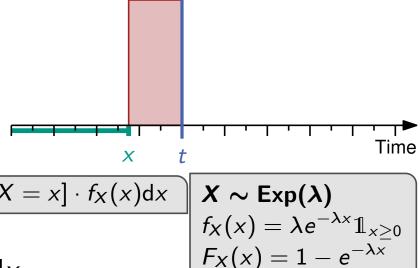


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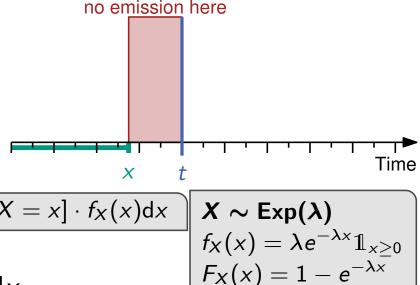
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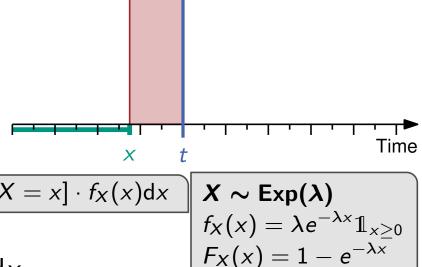


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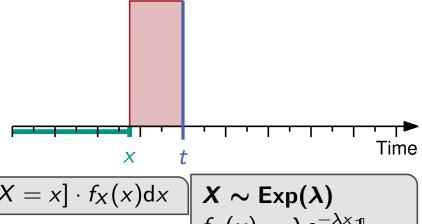


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$$f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x \ge 0}$$

 $F_X(x) = 1 - e^{-\lambda x}$

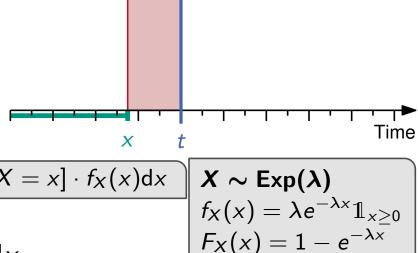


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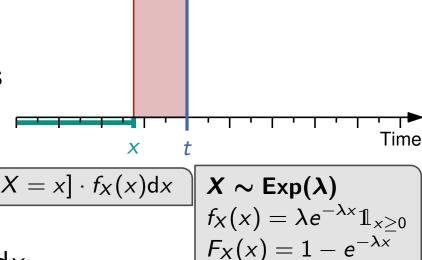


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independent





Motivation

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independent

Time
$$X = x \cdot f_X(x) dx$$

$$f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x \ge 0}$$

$$F_X(x) = 1 - e^{-\lambda x}$$



Motivation

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$$= x] \cdot f_X(x) dx$$

$$= x \cdot f_X(x) dx$$

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Due to memorylessness
$$\Pr[N(\square) = k] = \Pr[N(\square) = k]$$
 $X = x \cdot f_X(x) dx$
 $X \sim \text{Exp}(\lambda)$

$$f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x \geq 0}$$

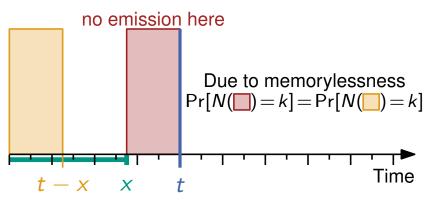
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 $=\lambda e^{-\lambda t} \int_0^t 1 dx = \lambda t e^{-\lambda t}$

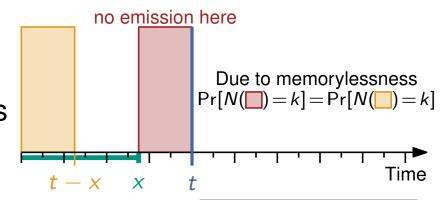


Motivation

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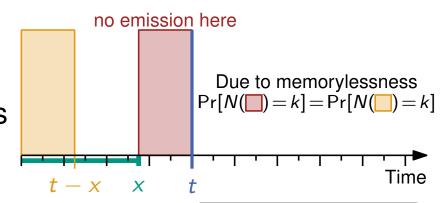
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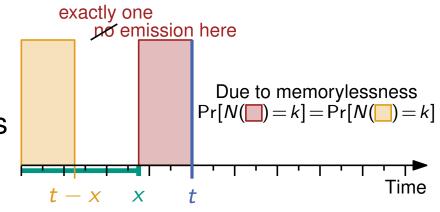
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 $\Pr[N_t = 2] = \Pr[X_1 \le t \land N(X_1, t) = 1]$



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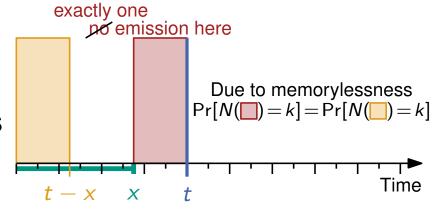
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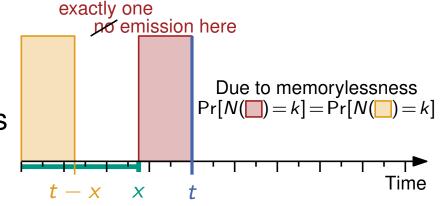
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Example 2 Preserve Law of Total Probability:
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$$\Pr[N_t = 2] = \int_{0}^{\infty} \Pr[X_t = t] \cdot N(x, t) = 1 \mid X_t = x \mid f_{X_t}(x) dx$$
independent



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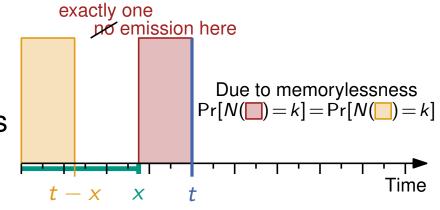


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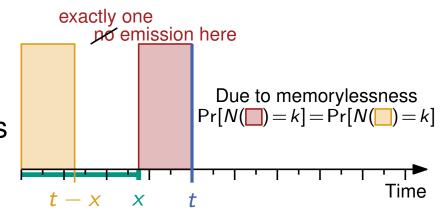


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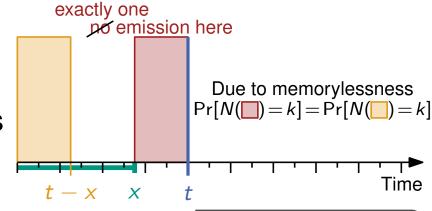


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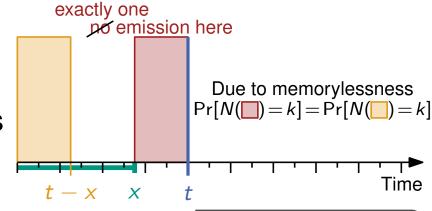


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exactly one potential emission here

Due to memorylessness
$$Pr[N(\square) = k] = Pr[N(\square) = k]$$
 $t - x \times t$

Time

$$X \sim \text{Exp}(\lambda)$$

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$$\Pr[A] = \int_{-\infty}^{\infty} \Pr[A \mid X = x] \cdot f_X(x) dx$$
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$$\begin{aligned} \Pr[N_t = 0] &= e^{-\lambda t} \quad \Pr[N_t = 1] = \lambda t e^{-\lambda t} \\ \Pr[N_t = 2] &= \int_0^t \Pr[N(x, t) = 1] \lambda e^{-\lambda x} \mathrm{d}x \\ &= \int_0^t \Pr[N_{t-x} = 1] \lambda e^{-\lambda x} \mathrm{d}x = \int_0^t \lambda (t - x) e^{-\lambda (t - x)} \cdot \lambda e^{-\lambda x} \mathrm{d}x \\ &= \lambda^2 e^{-\lambda t} \int_0^t t - x \mathrm{d}x \\ &= \lambda^2 e^{-\lambda t} \left(t \cdot \int_0^t 1 \mathrm{d}x - \int_0^t x \mathrm{d}x \right) \end{aligned}$$

$$f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x \ge 0}$$

 $F_X(x) = 1 - e^{-\lambda x}$

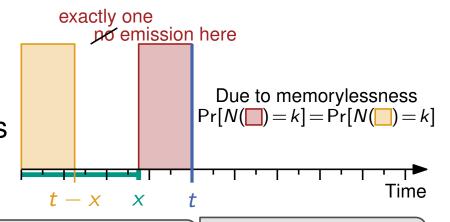
 $= \lambda^2 e^{-\lambda t} \left(t^2 - \left[\frac{1}{2} x^2 \right]_0^t \right) = \lambda^2 e^{-\lambda t} \cdot \frac{1}{2} t^2$



- Count number of particles emitted within a given time t
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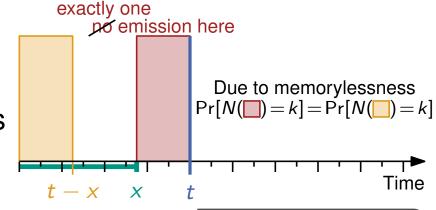


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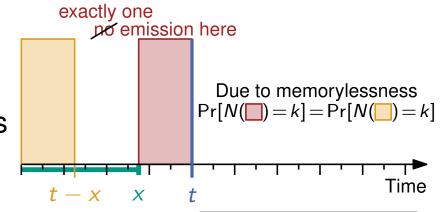


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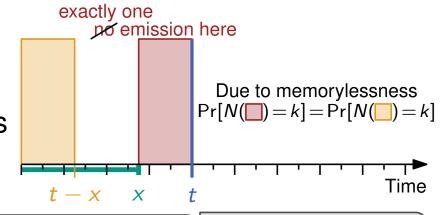


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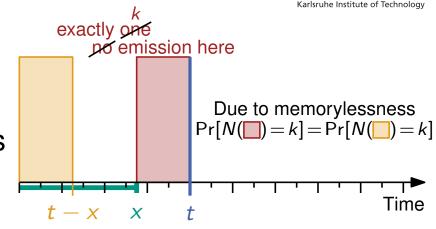
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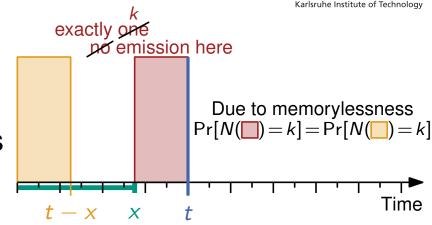
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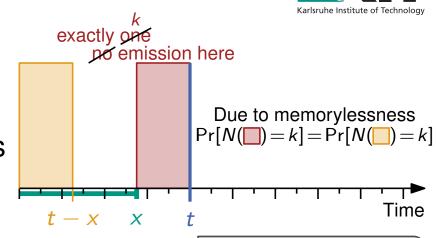
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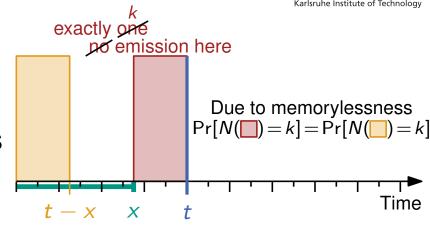
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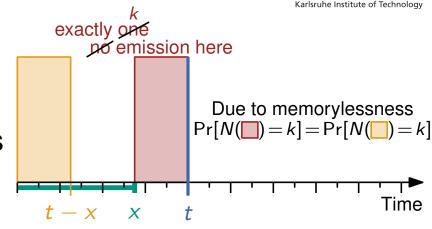
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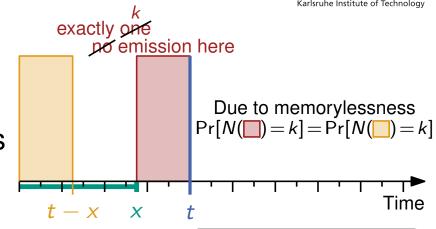
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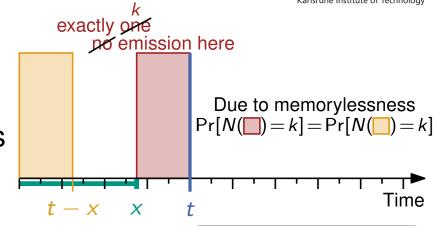
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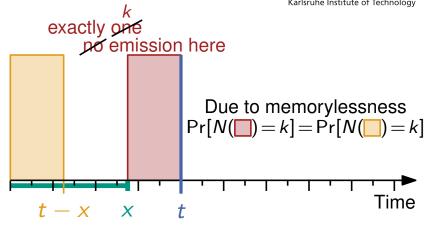
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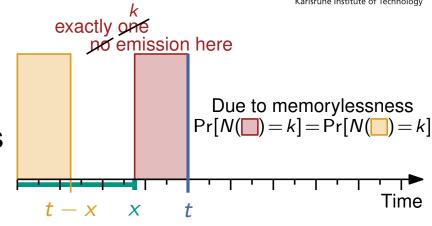
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$$= \frac{\lambda^{(k+1)} e^{-\lambda t}}{k!} \left[-\frac{1}{k+1} u^{(k+1)} \right]_{t}^{0} = \frac{\lambda^{(k+1)} e^{-\lambda t}}{(k+1)!} \left[u^{(k+1)} \right]_{0}^{t} = \frac{(\lambda t)^{(k+1)} e^{-\lambda t}}{(k+1)!} \checkmark$$

Integration by Substitution
$$u = g(x)$$

$$\int_{a}^{b} f(g(x)) dx = \int_{g(a)}^{g(b)} \frac{f(u)}{\left(\frac{dg(x)}{dx}\right)} du$$

exactly one por emission here

Due to memorylessness
$$Pr[N(\square) = k] = Pr[N(\square) = k]$$

Time

Integration by Substitution
$$u = g(x)$$

$$\int_{0}^{b} f(g(x)) dx = \int_{0}^{g(b)} \frac{f(u)}{f(u)} du$$

$$rac{\mathsf{d} g(x)}{\mathsf{d} x} = -1$$

 $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}_{x>0}$

 $F_X(x) = 1 - e^{-\lambda x}$



Definition: A **Poisson process** with *intensity* λ is a collection of random variables

 $X_1, X_2, ... \in \mathbb{R}$ such that, if $N(a, b) = |\{i \mid X_i \in [a, b]\}|$, then

 $ightharpoonup N(a,b) \sim \mathsf{Pois}(\lambda(b-a))$

(homogeneity)

lacksquare a < b < c < d: N(a, b) and N(c, d) are independent

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$$\Pr[N(a,b)=k]=\frac{(\lambda(b-a))^k e^{-\lambda(b-a)}}{k!}$$





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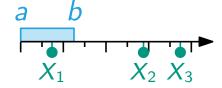
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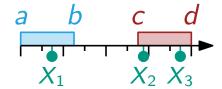
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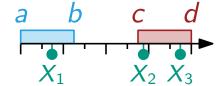
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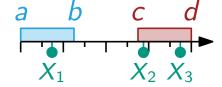
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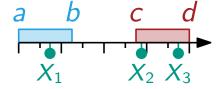
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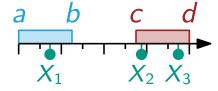
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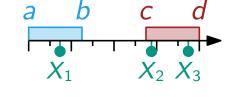
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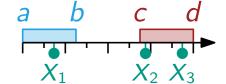
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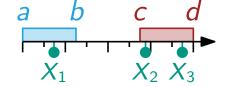
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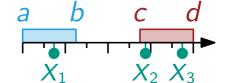
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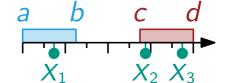
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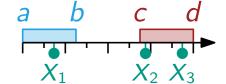
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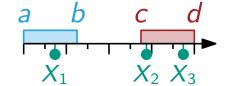
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■ In general: the positions of the points are distributed uniformly in an interval



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The **joint density function** $f_{X,Y}(a,b)$ satisfies $F_{X,Y}(a,b) = \int_{-\infty}^{a} \int_{-\infty}^{b} f_{X,Y}(x,y) dy dx$.



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$$f_{X|A}(x) = \begin{cases} f_X(x) / \Pr[A], & \text{if } x \in A, \\ 0, & \text{otherwhise.} \end{cases}$$



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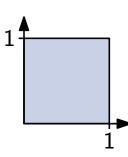
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Definition: Random variables X, Y are **independent** if $F_{X,Y}(x,y) = F_X(x) \cdot F_Y(y)$.

Karlsruhe Institute of Technology

Uniform Distribution on the Unit Square

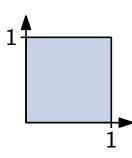
• We want to draw a point P uniformly at random from $[0, 1]^2$





Uniform Distribution on the Unit Square

- We want to draw a point P uniformly at random from $[0, 1]^2$
- Let X, Y be the x- and y-coordinates of P, respectively
- $f_P(x, y) = f_{X,Y}(x, y) = 1$ for $(x, y) \in [0, 1]^2$ and $f_P(x, y) = 0$, otherwise

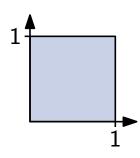




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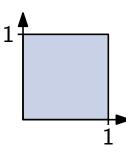


Marginal Density $f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$



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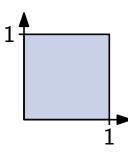
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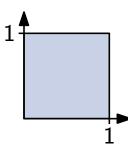
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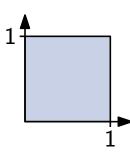
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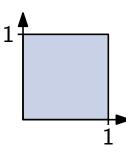
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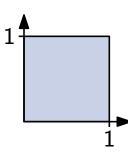
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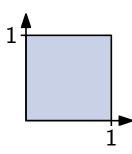
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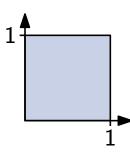
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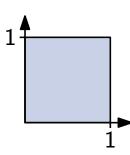
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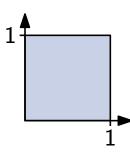
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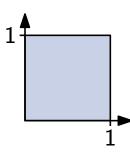
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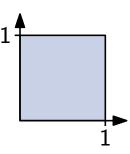
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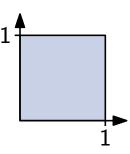
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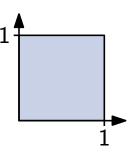
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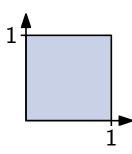
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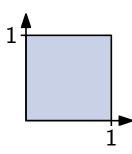
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$$E_{1}(b)$$
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■ Sample $P = (X, Y) \sim \mathcal{U}([0, 1]^2)$ by independently sampling $X, Y \sim \mathcal{U}([0, 1])!$

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Application: Random Geometric Graphs



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- Average-case analysis: analyze models that represent the real world
- So far: Erdős-Rényi random graphs (connect two vertices independently with equal prob)

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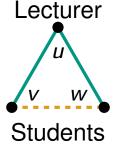
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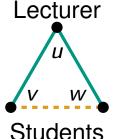
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Idea

- Vertices are likelier to connect if their distance is already small
 - ⇒ Define vertex distances in advance by introducing geometry

Students



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- Problem: In real networks, edges do *not* form independently
 - Two vertices are more likely to be adjacent if they have a common neighbor
 - ► This property is called *locality* or *clustering*
 - ER-graph: $\Pr[\{v, w\} \in E \mid \{u, v\} \in E \land \{u, w\} \in E] = \Pr[\{v, w\} \in E]$ 🗶

Idea

- Vertices are likelier to connect if their distance is already small
 - ⇒ Define vertex distances in advance by introducing geometry

Definition: A random geometric graph is obtained by distributing vertices in a metric space and connecting any two with a probability that depends on their distance.

How many? Which space? Which metric? Which distribution? Which probability?



Motivation

- Average-case analysis: analyze models that represent the real world
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v w Students

Lecturer

Idea

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Definition: A random geometric graph is obtained by distributing vertices in a metric space and connecting any two with a probability that depends on their distance.

How many? Which space? Which metric? Which distribution? Which probability? Simple & Realistic!



Random Geometric Graph



Number: n vertices

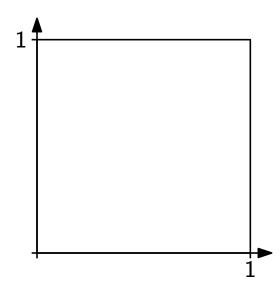
Random Geometric Graph



Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

Random Geometric Graph

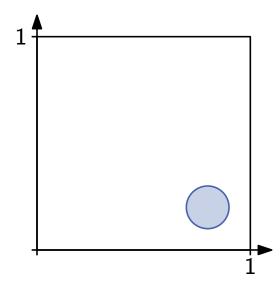




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Random Geometric Graph

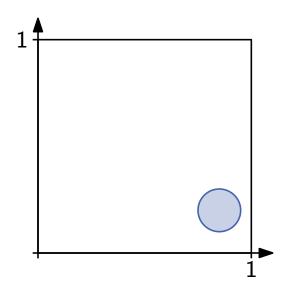




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Random Geometric Graph

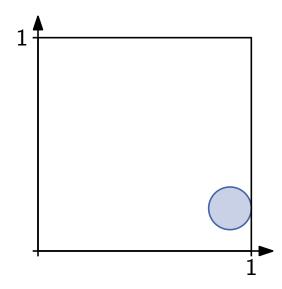




Number: n vertices

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Random Geometric Graph

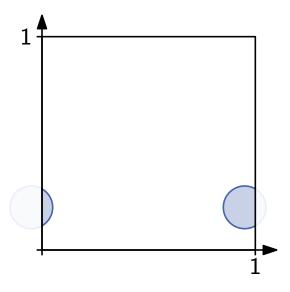




Number: n vertices

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Random Geometric Graph

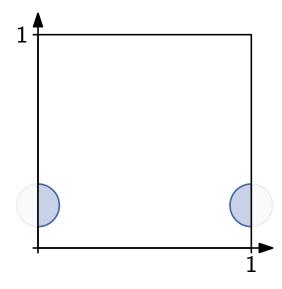




Number: n vertices

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Random Geometric Graph

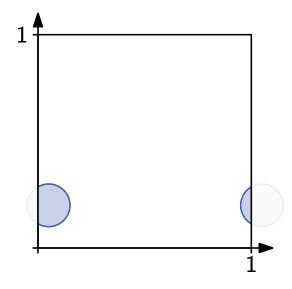




Number: n vertices

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Random Geometric Graph

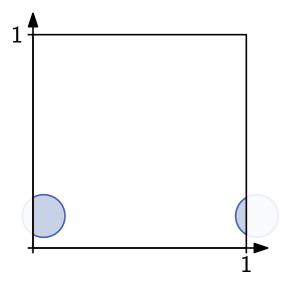




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Random Geometric Graph

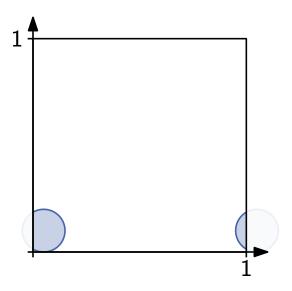




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Random Geometric Graph

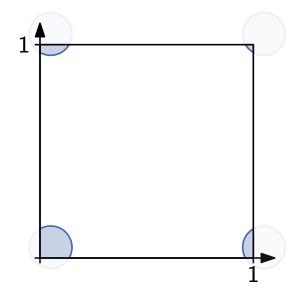




Number: n vertices

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Random Geometric Graph

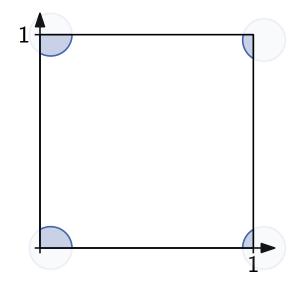




Number: n vertices

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Random Geometric Graph

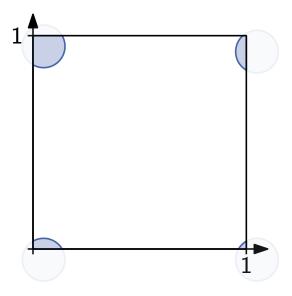




Number: n vertices

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Random Geometric Graph

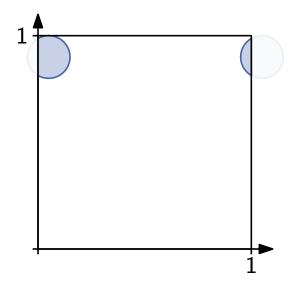




Number: n vertices

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Random Geometric Graph





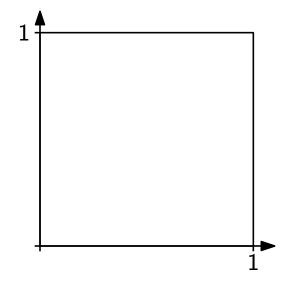
Number: n vertices

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■ Metric: for $p = (p_1, p_2)$, $q = (q_1, q_2)$: $d_i = |p_i - q_i|$ L_{∞} norm: $d(p, q) = \max_{i \in \{1, 2\}} \min\{d_i, 1 - d_i\}$

"Chebychev distance"

Random Geometric Graph



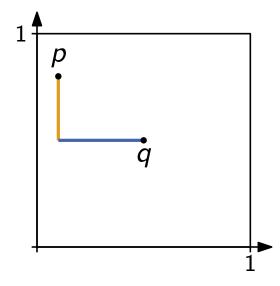


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Random Geometric Graph



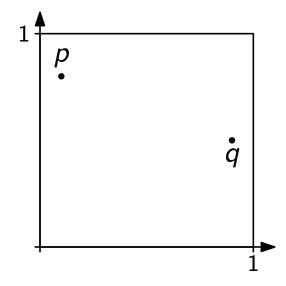


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Random Geometric Graph



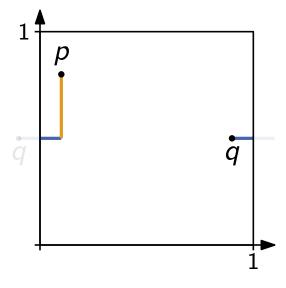


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Random Geometric Graph



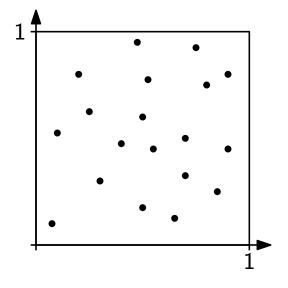


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Random Geometric Graph



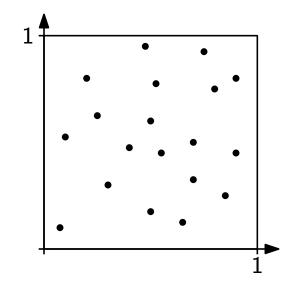


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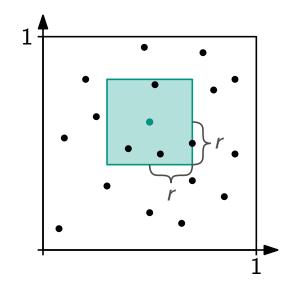
Random Geometric Graph





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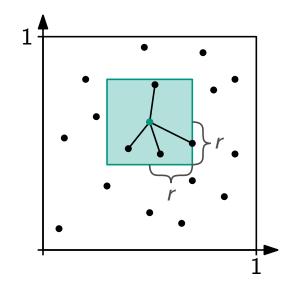
Random Geometric Graph





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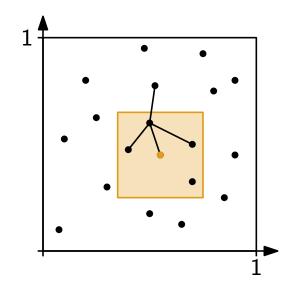
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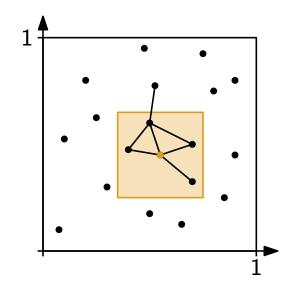


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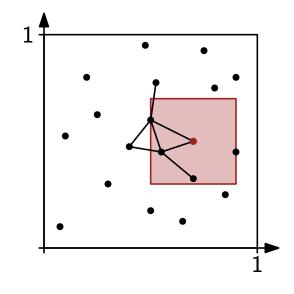


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Random Geometric Graph



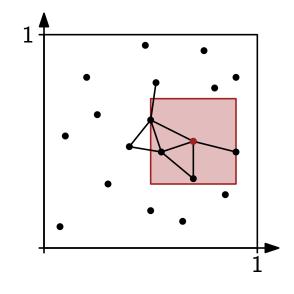


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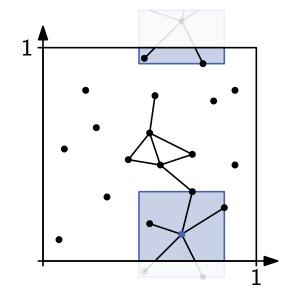
Random Geometric Graph





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Random Geometric Graph



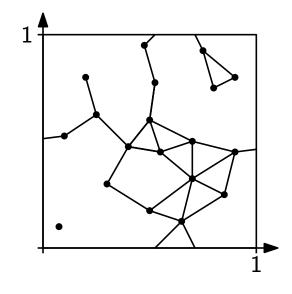


Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

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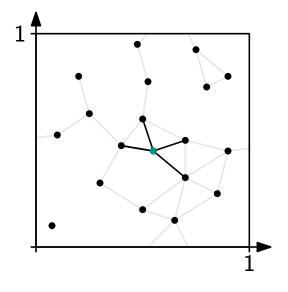
Random Geometric Graph





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Random Geometric Graph





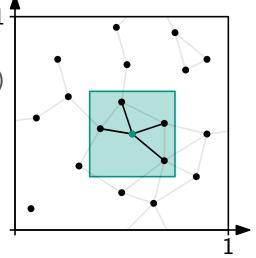
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- Distribution: For each ν independently: $P_{\nu} \sim \mathcal{U}([0,1]^2)$
- Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases} \text{ threshold parameter} \\ \text{Expected Degree of } v \end{cases}$

■ Neighbors of v are in N(v) (here N(v) denotes the *region* in the ground space)

Random Geometric Graph





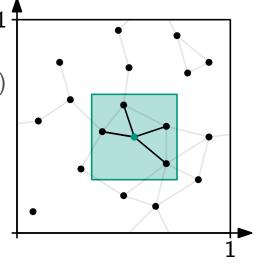
Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

- Metric: for $p = (p_1, p_2), q = (q_1, q_2)$: $d_i = |p_i q_i|$ L_{∞} norm: $d(p,q) = \max_{i \in \{1,2\}} \min\{d_i, 1-d_i\}$
- Distribution: For each ν independently: $P_{\nu} \sim \mathcal{U}([0,1]^2)$
- Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases}$ threshold parameter 0, otherwise

- Neighbors of v are in N(v) (here N(v) denotes the *region* in the ground space)
- $\blacksquare \mathbb{E}[\deg(v)] = \mathbb{E}[\sum_{u \in V \setminus \{v\}} \mathbb{1}_{\{P_u \in N(v)\}}] = \sum_{u \in V \setminus \{v\}} \Pr[d(P_u, P_v) \le r]$

Random Geometric Graph





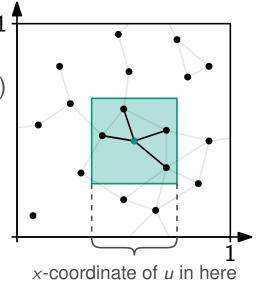
Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

- Metric: for $p = (p_1, p_2), q = (q_1, q_2)$: $d_i = |p_i q_i|$ L_{∞} norm: $d(p,q) = \max_{i \in \{1,2\}} \min\{d_i, 1-d_i\}$
- Distribution: For each ν independently: $P_{\nu} \sim \mathcal{U}([0,1]^2)$
- Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases}$ threshold parameter 0, otherwise

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Random Geometric Graph





Number: n vertices

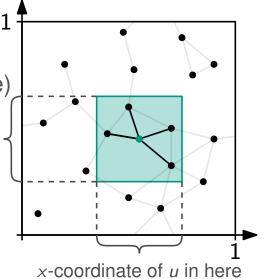
■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

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$$\mathbb{E}[\deg(v)] = \mathbb{E}[\sum_{u \in V \setminus \{v\}} \mathbb{1}_{\{P_u \in N(v)\}}] = \sum_{u \in V \setminus \{v\}} \Pr[d(P_u, P_v) \leq r]$$
 and y-coordinate of u in here

Random Geometric Graph





Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

• Metric: for $p = (p_1, p_2), q = (q_1, q_2)$: $d_i = |p_i - q_i|$ L_{∞} norm: $d(p,q) = \max_{i \in \{1,2\}} \min\{d_i, 1-d_i\}$

■ Distribution: For each ν independently: $P_{\nu} \sim \mathcal{U}([0,1]^2)$

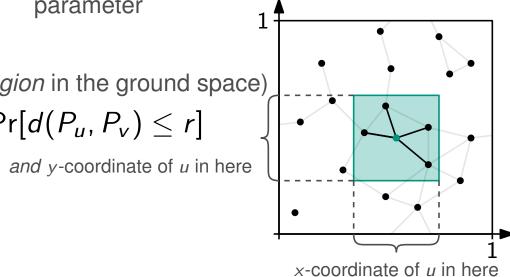
Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases}$ threshold parameter **Expected Degree of** v

■ Neighbors of v are in N(v) (here N(v) denotes the *region* in the ground space)

$$\blacksquare \mathbb{E}[\deg(v)] = \mathbb{E}[\sum_{u \in V \setminus \{v\}} \mathbb{1}_{\{P_u \in N(v)\}}] = \sum_{u \in V \setminus \{v\}} \Pr[d(P_u, P_v) \le r]$$

■ Draw $P_u = (X, Y)$ as independent $X, Y \sim \mathcal{U}([0, 1])$

Random Geometric Graph





Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

■ Metric: for $p = (p_1, p_2), q = (q_1, q_2)$: $d_i = |p_i - q_i|$ L_{∞} norm: $d(p,q) = \max_{i \in \{1,2\}} \min\{d_i, 1-d_i\}$

■ Distribution: For each ν independently: $P_{\nu} \sim \mathcal{U}([0,1]^2)$

Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases}$ threshold parameter 0, otherwise

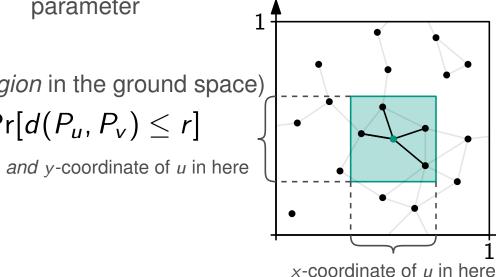
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■ Draw $P_u = (X, Y)$ as independent $X, Y \sim \mathcal{U}([0, 1])$

$$X \sim \mathcal{U}([a,b])$$
 : $\mathsf{Pr}[X \in [c,d] \subseteq [a,b]] = rac{d-c}{b-a}$

Random Geometric Graph





Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

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Probability $\Pr[\{u,v\} \in E] = \begin{cases} 1, & \text{if } d(P_u,P_v) \leq r \end{cases}$ threshold parameter 0, otherwise

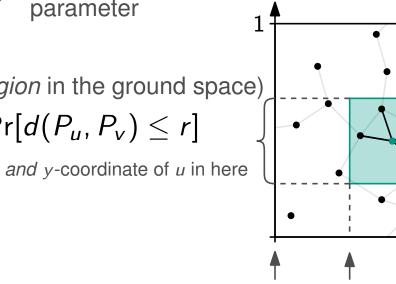
■ Neighbors of v are in N(v) (here N(v) denotes the *region* in the ground space)

$$\blacksquare \mathbb{E}[\deg(v)] = \mathbb{E}[\sum_{u \in V \setminus \{v\}} \mathbb{1}_{\{P_u \in N(v)\}}] = \sum_{u \in V \setminus \{v\}} \Pr[d(P_u, P_v) \le r]$$

■ Draw $P_u = (X, Y)$ as independent $X, Y \sim \mathcal{U}([0, 1])$

$$X \sim \mathcal{U}([a,b])$$
 : $\Pr[X \in [c,d] \subseteq [a,b]] = rac{d-c}{b-a}$

Random Geometric Graph





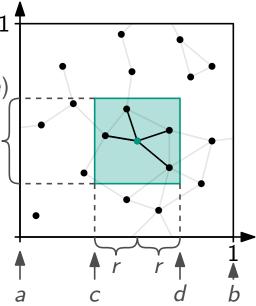
- Number: n vertices
- Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)
- Metric: for $p = (p_1, p_2), q = (q_1, q_2)$: $d_i = |p_i q_i|$ L_{∞} norm: $d(p,q) = \max_{i \in \{1,2\}} \min\{d_i, 1-d_i\}$
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Random Geometric Graph

Nodes distributed in metric space Connection probability depends on distance



and v-coordinate of u in here



Number: n vertices

■ Space: 2-dimensional torus \mathbb{T}^2 (unit square with opposite sides identified)

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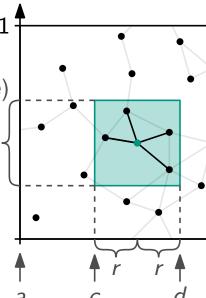
$$= \sum_{u \in V \setminus \{v\}} \Pr[d(P_u, P_v) \leq r]$$

$$= \sum_{u \in V \setminus \{v\}} \frac{2r}{1 - 0} \cdot \frac{2r}{1 - 0}$$

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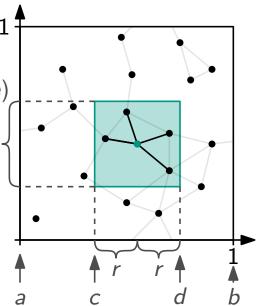
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and y-coordinate of u in here

$$= \sum_{u \in V \setminus \{v\}} \frac{2r}{1-0} \cdot \frac{2r}{1-0}$$

$$- \sum_{u \in V \setminus \{v\}} 1 - 1 = (n-1) \cdot 4r^2$$

Random Geometric Graph





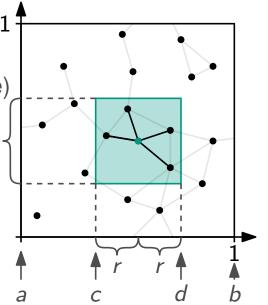
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$$X \sim \mathcal{U}([a,b]) : \Pr[X \in [c,d] \subseteq [a,b]] = \frac{d-c}{b-a}$$
 = $(n-1) \cdot 4r^2$ (area of the region $N(v)$)

Random Geometric Graph

Nodes distributed in metric space Connection probability depends on distance



and y-coordinate of u in here

 $\stackrel{\blacktriangledown}{=} \sum_{u \in V \setminus \{v\}} \frac{2r}{1-0} \cdot \frac{2r}{1-0}$



Locality



Locality Realistic assumption: $r = \Theta(n^{-1/2})$ such that $\mathbb{E}[\deg(v)] = \Theta(1)$



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Locality

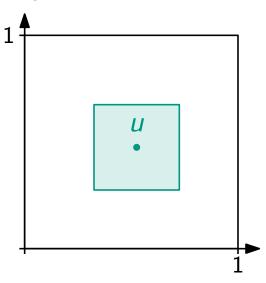
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Locality

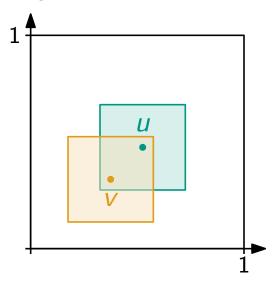
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Locality

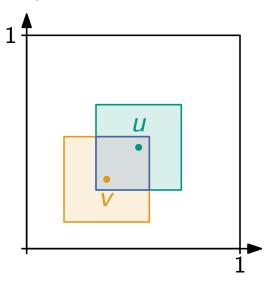
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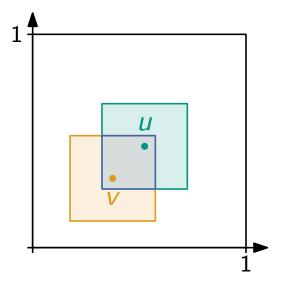
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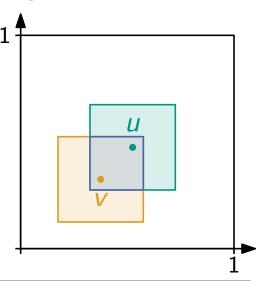
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Law of Total Probability
$$Pr[A] = \int_{-\infty}^{\infty} Pr[A \mid X = x] f_X(x) dx$$



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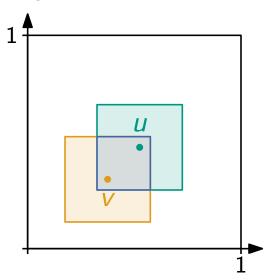
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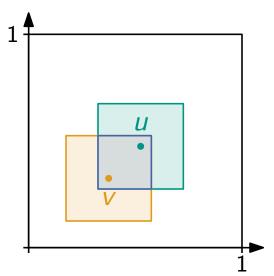
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$$[0, 1]^{2}$$

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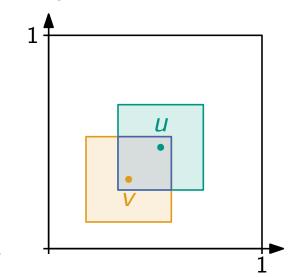
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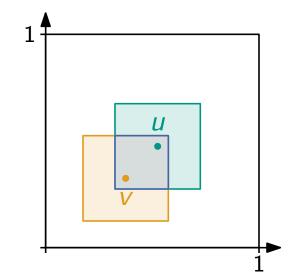
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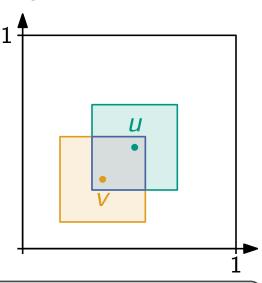
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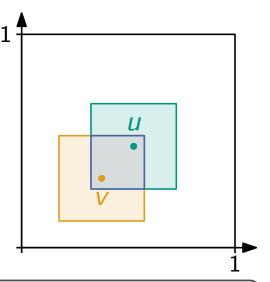
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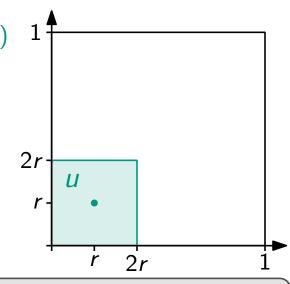
Two vertices v and w are likelier to connect if they have a common neighbor u

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 w.l.o.g assume $u = (r, r)$ 1 $\Pr[\{v, w\} \in E \mid \{u, v\} \in E \land \{u, w\} \in E]$ $= \Pr[w \in N(v) \mid v \in N(u) \land w \in N(u)] = \frac{\Pr[w \in N(v) \land v \in N(u) \land w \in N(u)]}{\Pr[v \in N(u) \land w \in N(u)]}$

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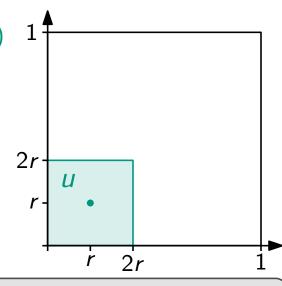
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Law of Total Probability
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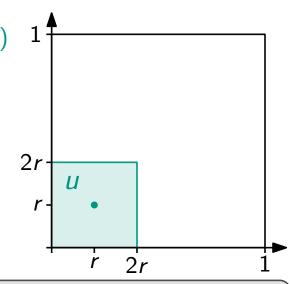
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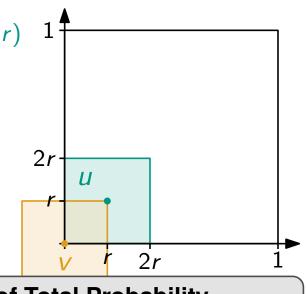
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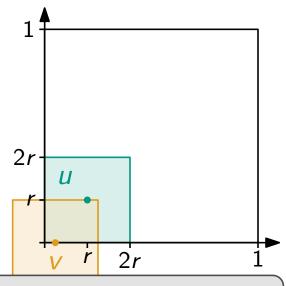
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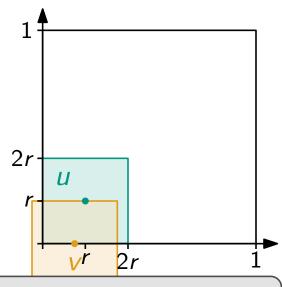
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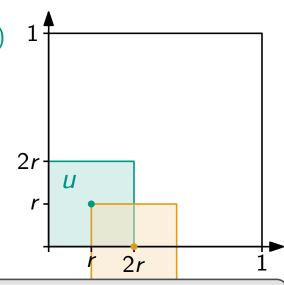
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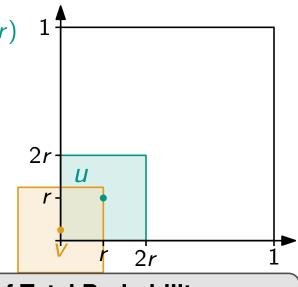
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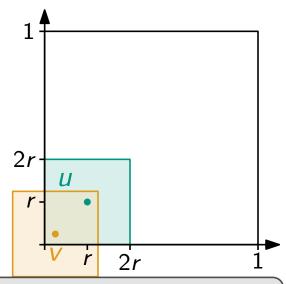
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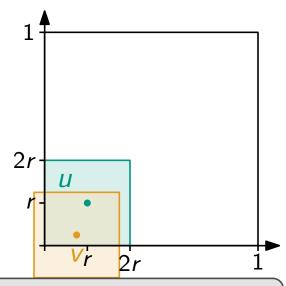
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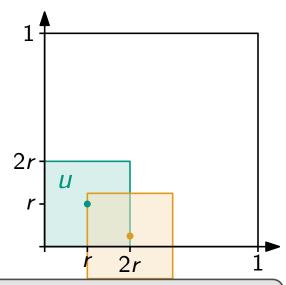
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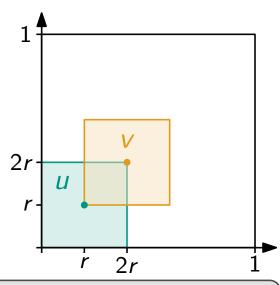
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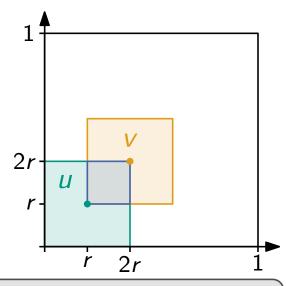
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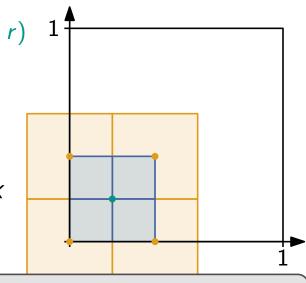
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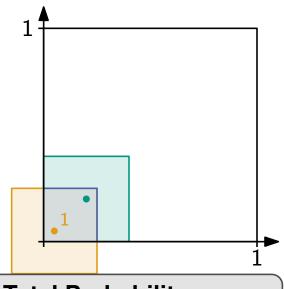
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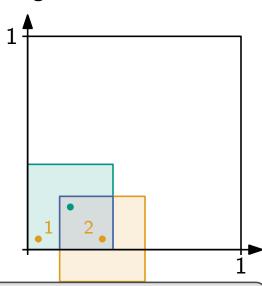
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Convention: $v = P_v$

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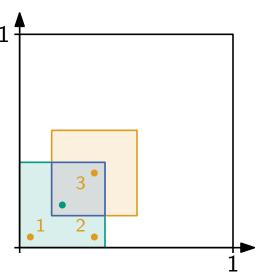
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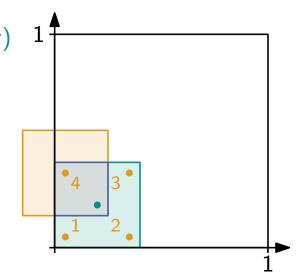
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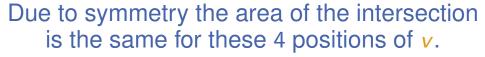
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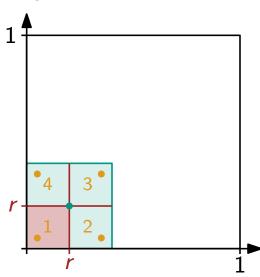
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⇒ Integrate only one quarter and multiply by 4



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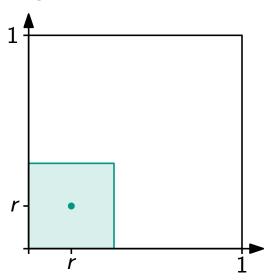
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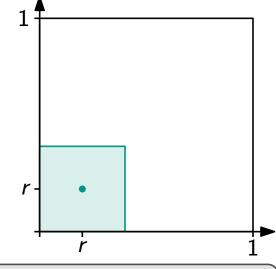
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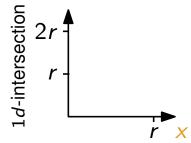
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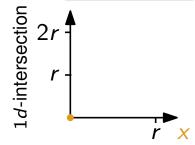
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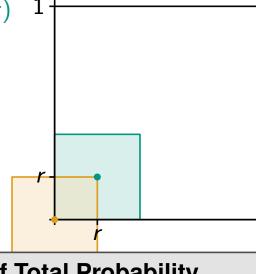
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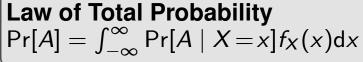
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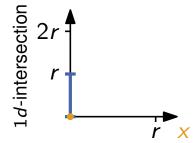
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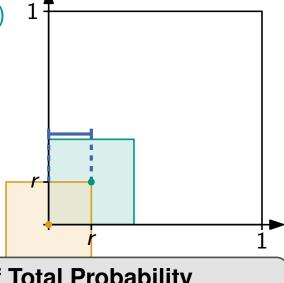
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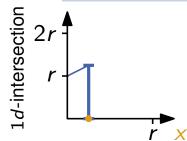
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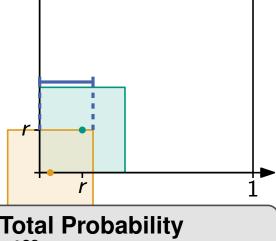
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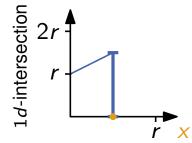
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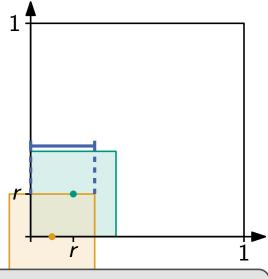
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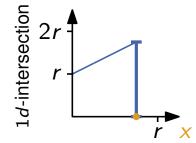
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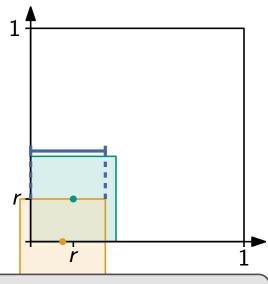
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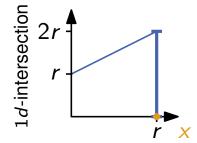
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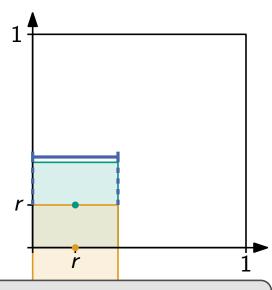
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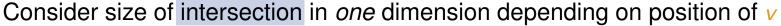
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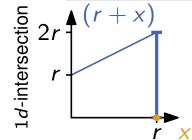
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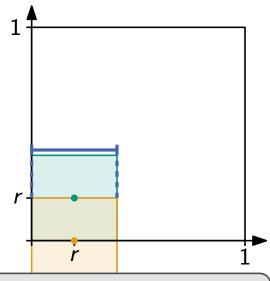
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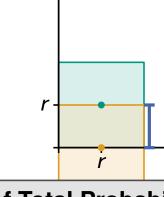
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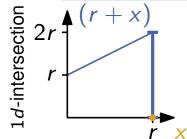
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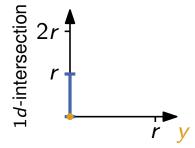
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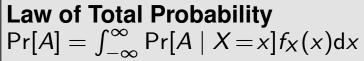
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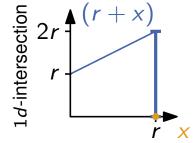
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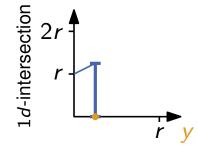
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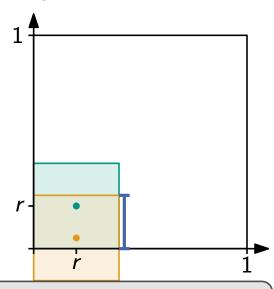
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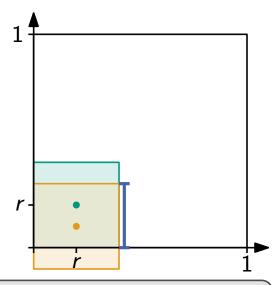
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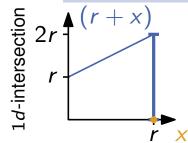
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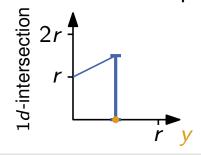
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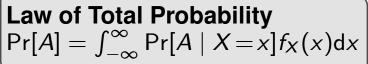
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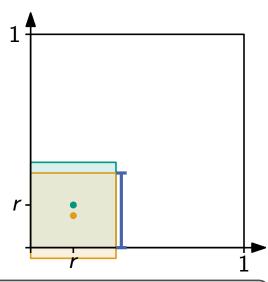
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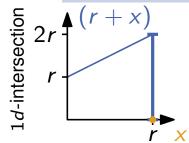
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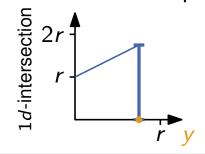
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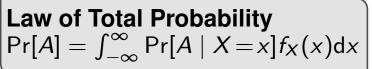
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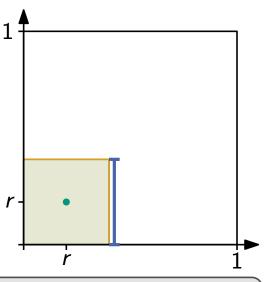
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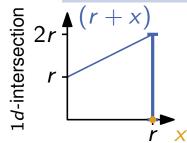
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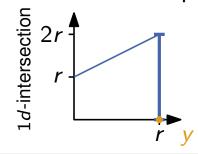
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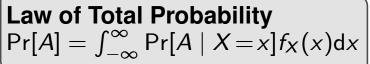
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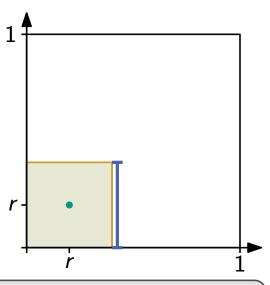
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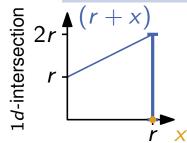
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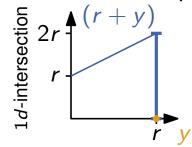
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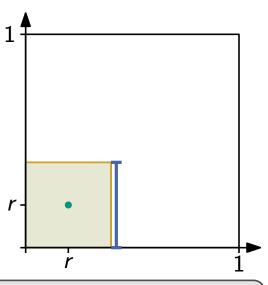
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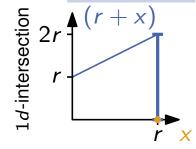
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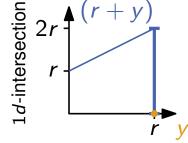
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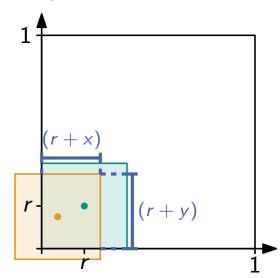
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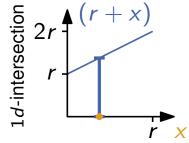
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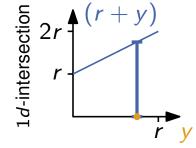
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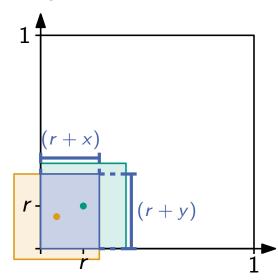
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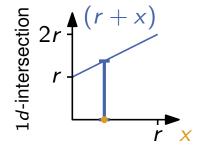
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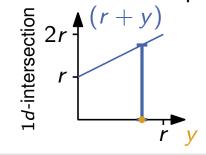
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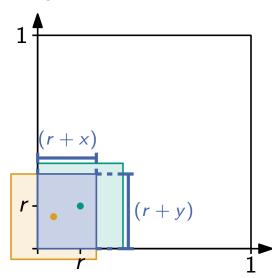
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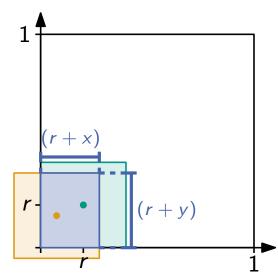
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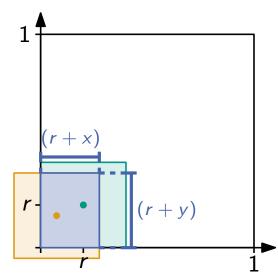
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Locality

Realistic assumption: $r = \Theta(n^{-1/2})$ such that $\mathbb{E}[\deg(v)] = \Theta(1)$

Convention: $v = P_v$

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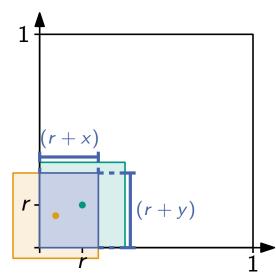
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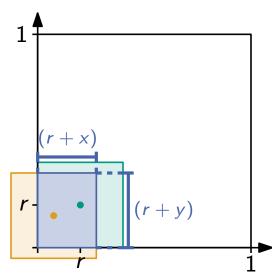
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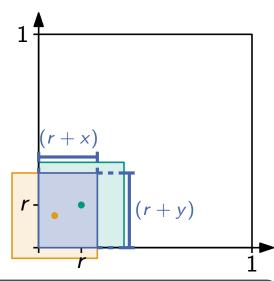
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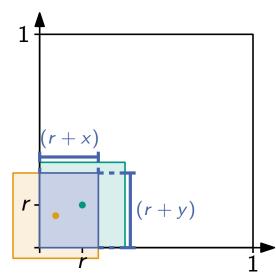
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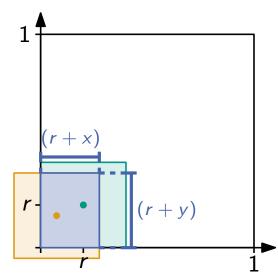
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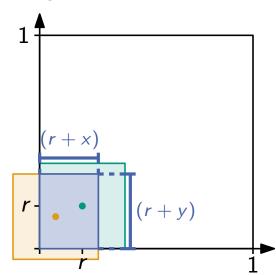
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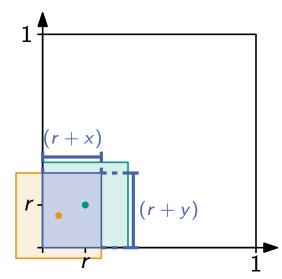
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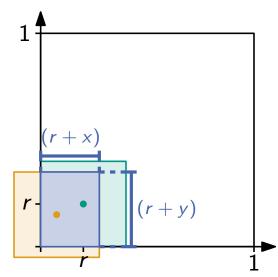
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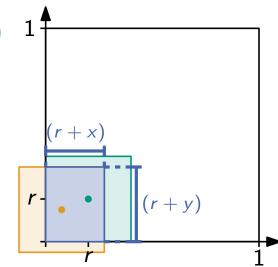
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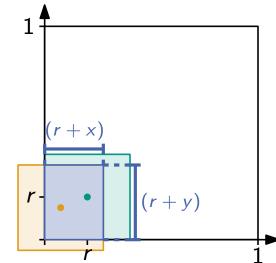
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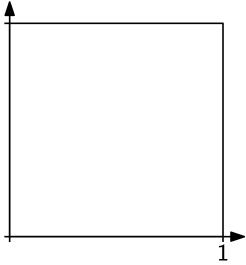
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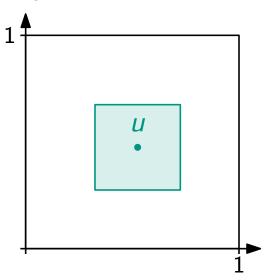
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Denominator

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Locality

Realistic assumption: $r = \Theta(n^{-1/2})$ such that $\mathbb{E}[\deg(v)] = \Theta(1)$

Convention: $v = P_v$

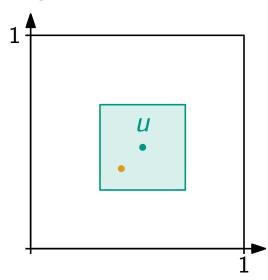
Two vertices v and w are likelier to connect if they have a common neighbor u

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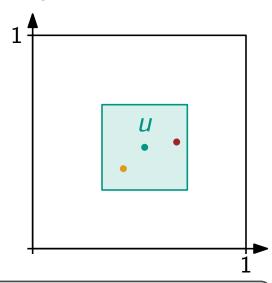
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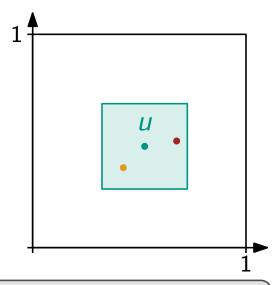
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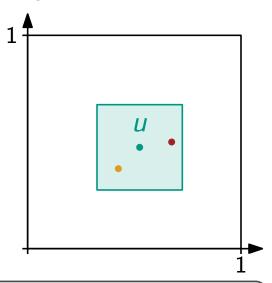
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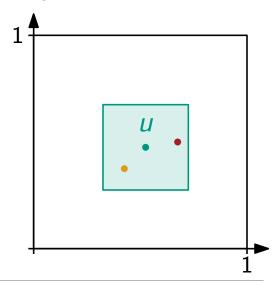
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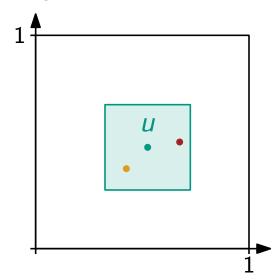
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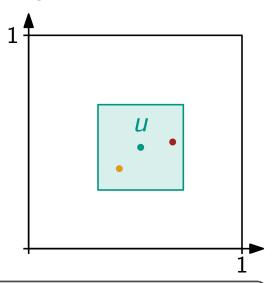
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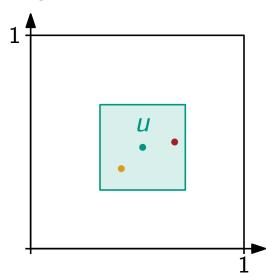
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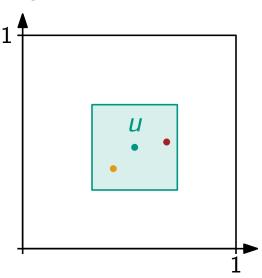
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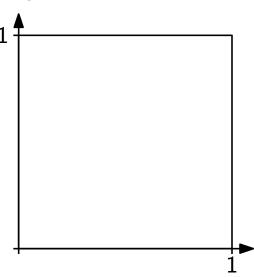
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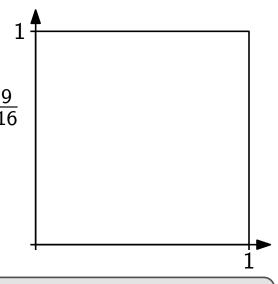
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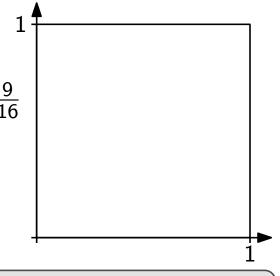
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 \Rightarrow $\Pr[\{v, w\} \in E \mid \{u, v\} \in E \land \{u, w\} \in E] = \Theta(1)\}$ \Rightarrow $\Pr[w \in N(v) \mid v \in N(u) \land w \in N(u)] = \frac{\Pr[w \in N(v) \land v \in N(u) \land w \in N(u)]}{\Pr[v \in N(u) \land w \in N(u)]} = \frac{9}{16}$ **Numerator** $\Pr[w \in N(v) \land v \in N(u) \land w \in N(u)] = 9r^4$ **Denominator**

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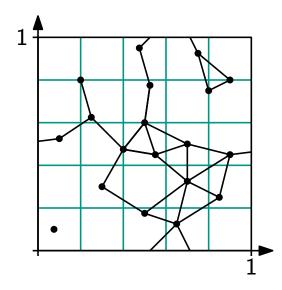


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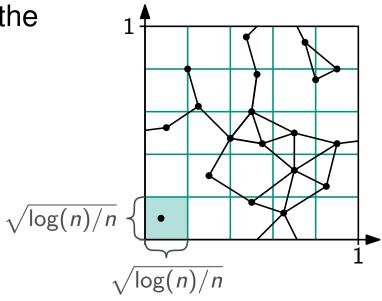
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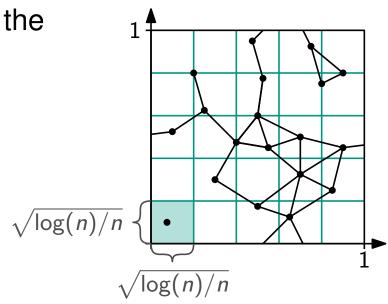
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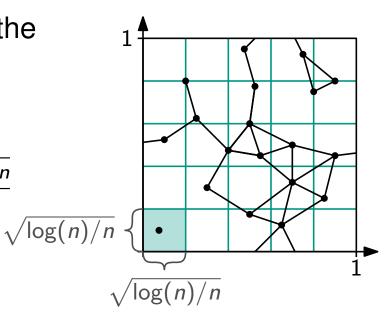
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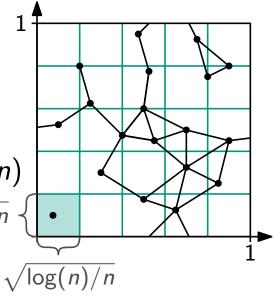




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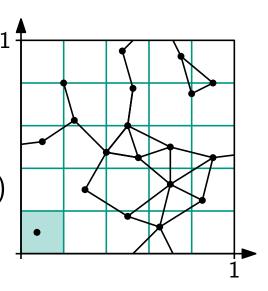
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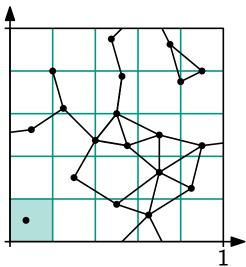




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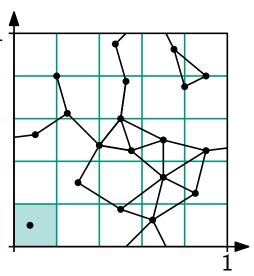
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■ Same distribution for all X_i : $\Pr[\forall i : X_i = \log(n)] = \prod_i \Pr[X_i = \log(n)]$





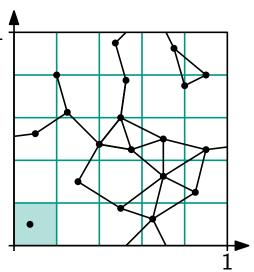
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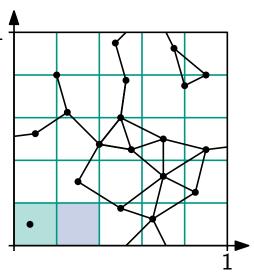


- Discretize the space into equally sized grid cells, such that the expected number of vertices in each cell is log(n)
- Each cell C_i has width and height $\sqrt{\log(n)/n}$
- \blacksquare Let X_i denote the number of vertices in C_i

$$\mathbb{E}[X_i] = \mathbb{E}[\sum_{v \in V} \mathbb{1}_{\{v \in C_i\}}] = n \cdot \Pr[v \in C_i] = n \frac{\sqrt{\log(n)/n}}{1-0} \frac{\sqrt{\log(n)/n}}{1-0} = \log(n)$$

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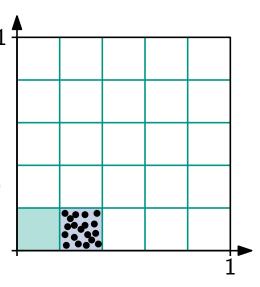


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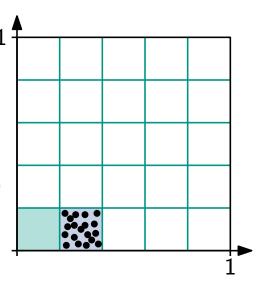
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https://i.imgflip.com/1pln6k.jpg?a471949



Idea

Avoid dependencies by replacing uniform point sampling with a Poisson point process



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Definition: A **Poisson** *Point* **process** with *intensity* λ is a collection of random variables $X_1, X_2, ... \in \mathbb{R}^2$ such that, if |A| is the area of A and $N(A) = |\{i \mid X_i \in A\}|$, then

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(homogeneity)

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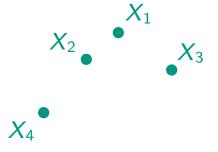
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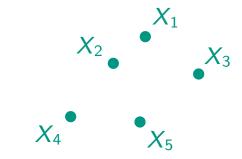
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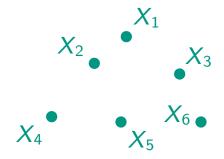
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Idea

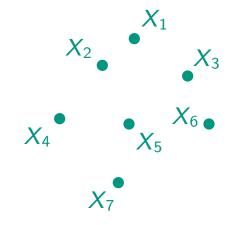
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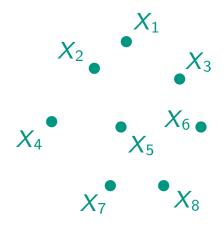
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ldea

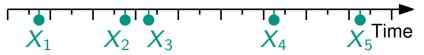
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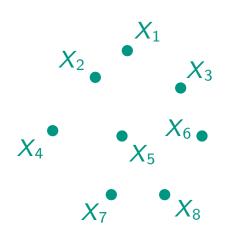
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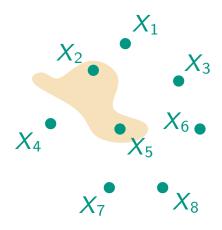
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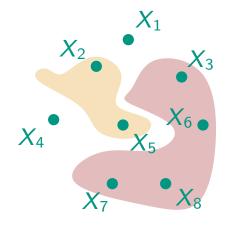
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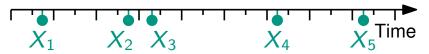
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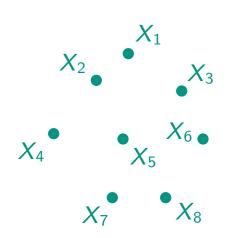
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(homogeneity)



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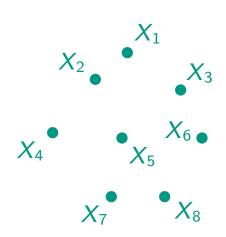
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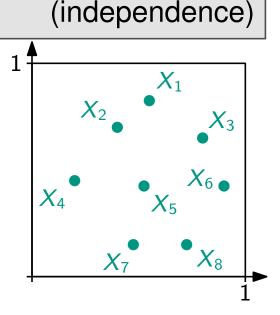
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Idea

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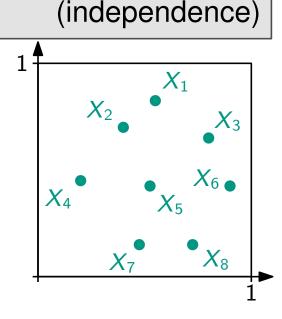
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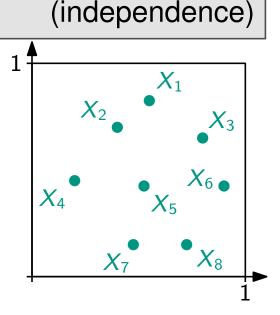
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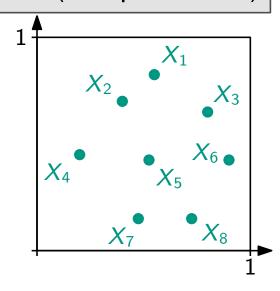
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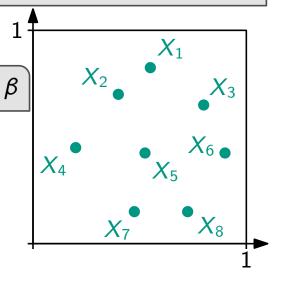
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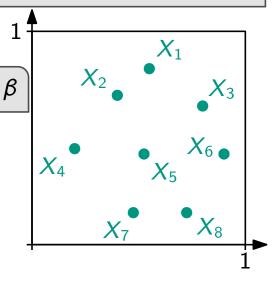
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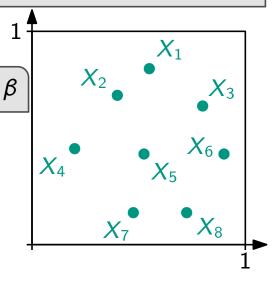
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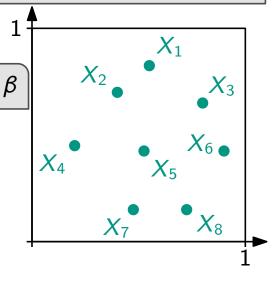
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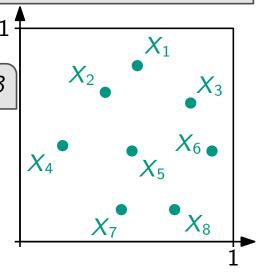
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(Generalizes to arbitrary dimension, 1d is the Poisson process seen earlier)

- Note: We do not know how many points we get!
- How do we choose λ ?
 - We should at least expect n points in our ground space $[0, 1]^2$

$$n = \mathbb{E}[|\{i \mid X_i \in [0, 1]^2\}|] = \mathbb{E}[N([0, 1]^2)] = \lambda|[0, 1]^2| = \lambda$$

Recall: conditioned on their number, points are distributed uniformly





Idea

Avoid dependencies by replacing uniform point sampling with a Poisson point process

Definition: A **Poisson** *Point* **process** with *intensity* λ is a collection of random variables $X_1, X_2, ... \in \mathbb{R}^2$ such that, if |A| is the area of A and $N(A) = |\{i \mid X_i \in A\}|$, then

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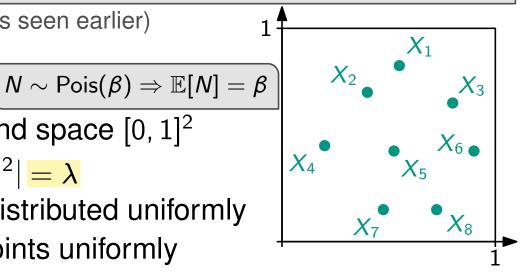
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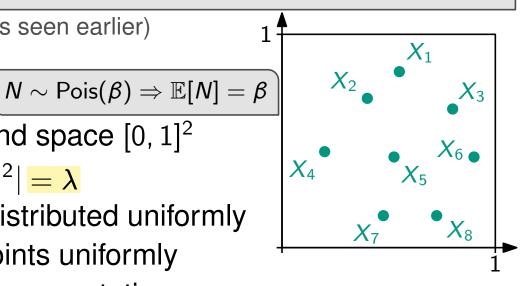
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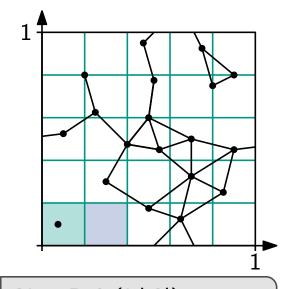
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- The resulting **Poissonized RGG** has *n* vertices in expectation





- Vertices of RGG distributed using Poisson point process with intensity $\lambda = n$
- Discretize the space into equally sized grid cells, such that the expected number of vertices in each cell is log(n)
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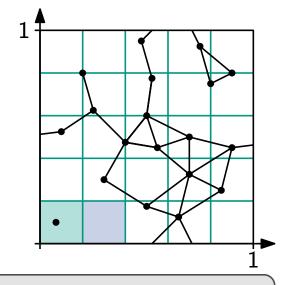


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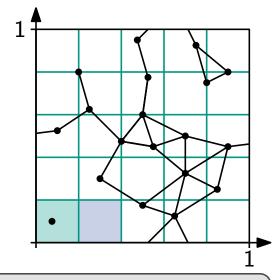
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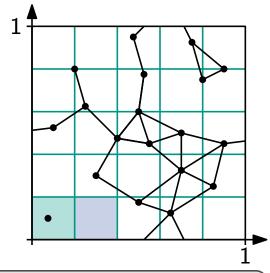
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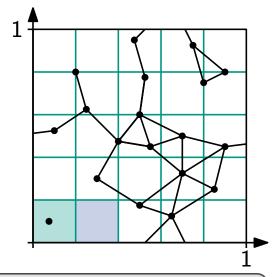
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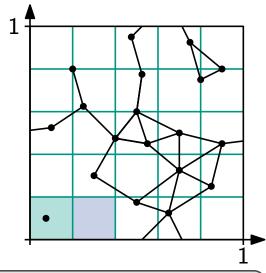
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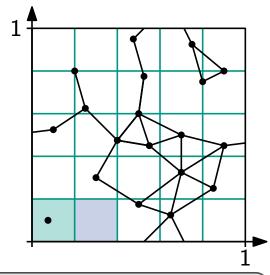
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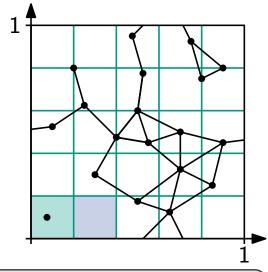
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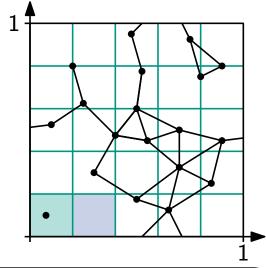
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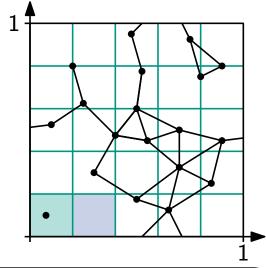


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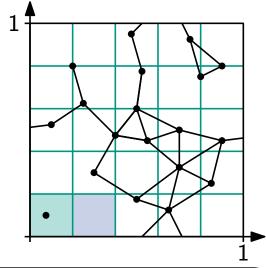
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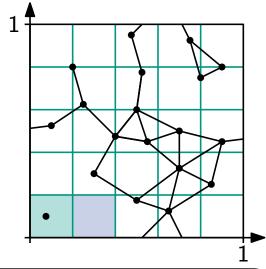
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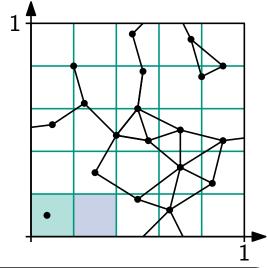
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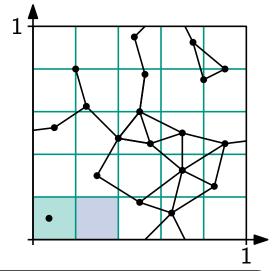
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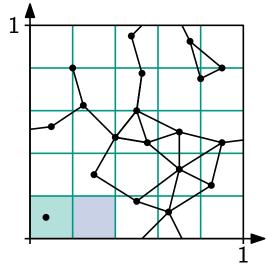
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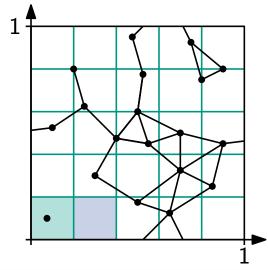
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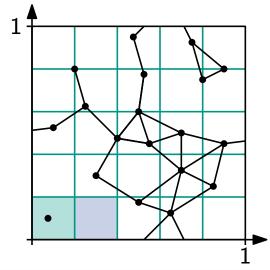
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- Each cell C_i has width and height $\sqrt{\log(n)/n} \Rightarrow |C_i| = \log(n)/n$
- Let X_i denote the number of vertices in $C_i \Rightarrow X_i \sim \text{Pois}(\lambda |C_i|)$

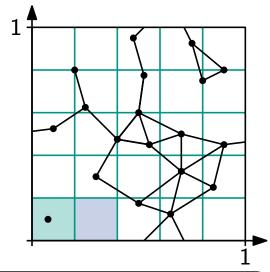
$$\mathbb{E}[X_i] = \lambda |C_i| = \log(n)$$

• What is the probability that each cell gets exactly log(n) vertices?

$$\Pr[X_i = \log(n)] = \frac{(\lambda |C_i|)^{\log(n)} e^{-\lambda |C_i|}}{\log(n)!} = \frac{(n \frac{\log(n)}{p'})^{\log(n)} e^{-p \frac{\log(n)}{p'}}}{\log(n)!} = \frac{\log(n)^{\log(n)} e^{-\log(n)}}{\log(n)!}$$

$$\leq \frac{\log(n)^{\log(n)} e^{-\log(n)}}{e(\frac{\log(n)}{p})^{\log(n)}} = \frac{1}{e} \quad \text{there are } n/\log(n) \text{ cells}$$

■ Same distribution for all X_i : $\Pr[\forall i : X_i = \log(n)] = \prod_i^{\bullet} \Pr[X_i = \log(n)]$ by definition, disjoint regions *are* independent $-\sqrt{e^{-n/\log(n)}}$



$$egin{aligned} \mathcal{N} & \sim \mathsf{Pois}(\pmb{\lambda}|\pmb{A}|) \ \mathbb{E}[\mathcal{N}] &= \lambda|A| \ \mathsf{Pr}[\mathcal{N} &= k] &= rac{(\lambda|A|)^k e^{-\lambda|A|}}{k!} \end{aligned}$$

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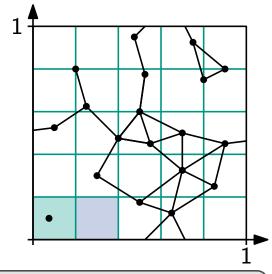
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$$N \sim \text{Pois}(\lambda|A|)$$
 $\mathbb{E}[N] = \lambda|A|$
 $\Pr[N = k] = \frac{(\lambda|A|)^k e^{-\lambda|A|}}{k!}$

$$k! \geq e(k/e)^k$$

but we cheated...



Situation

■ We started with a simple RGG $(n, \mathbb{T}^2, L_\infty$ -norm, $P_i \sim \mathcal{U}([0,1]^2)$, $\Pr[\{u,v\} \in E] = \mathbb{1}_{\{d(u,v) \leq r\}})$



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Recall

Conditioned on the number of points in area A, the points are distributed uniformly in A



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Seen so far

- Simple RGG
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Random Geometric Graph



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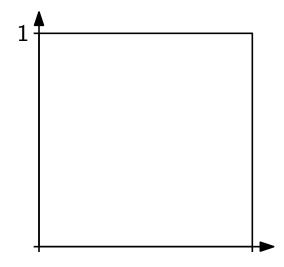
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More commonly used model

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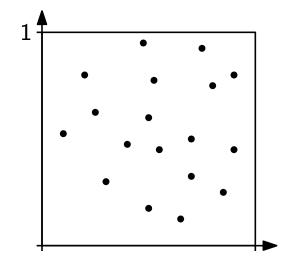
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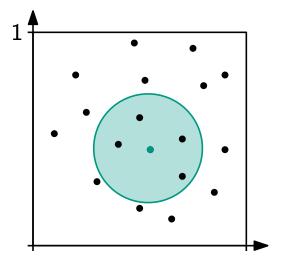
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Random Geometric Graph

Nodes distributed in metric space Connection probability depends on distance

N(v) is a disk





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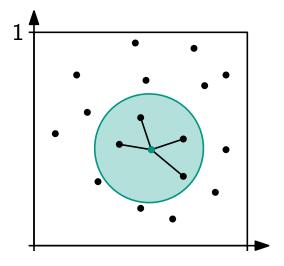
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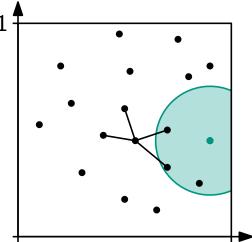


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- n, $[0,1]^2$, L_2 -norm, $P_i \sim \mathcal{U}([0,1]^2)$, $\Pr[\{u,v\} \in E] = \mathbb{1}_{d(u,v) \leq r}$

Random Geometric Graph

Nodes distributed in metric space Connection probability depends on distance



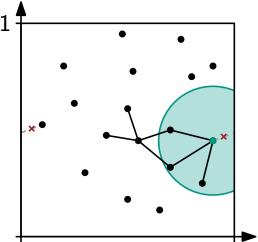


Seen so far

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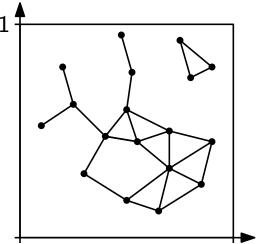


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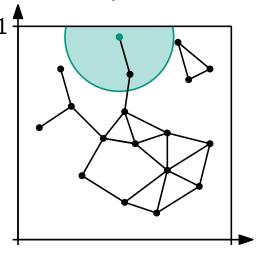


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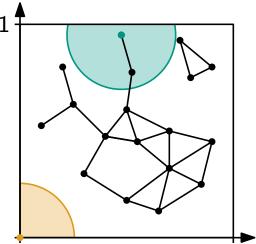


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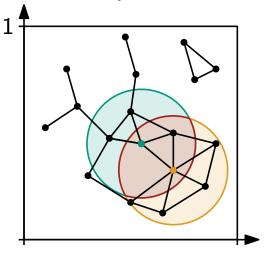


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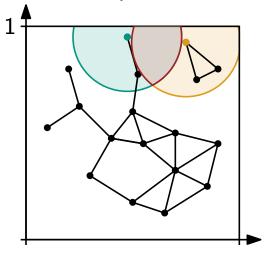


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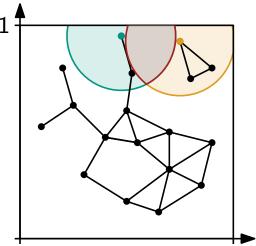


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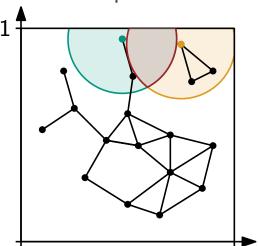
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Problem: Homogeneous degree distribution does not match many real-world graphs

Random Geometric Graph

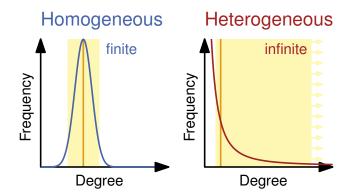
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Motivation

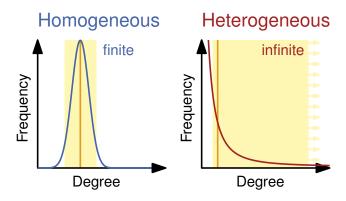
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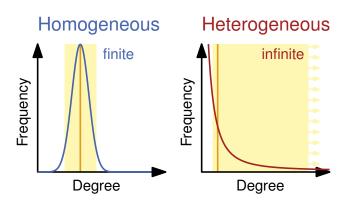
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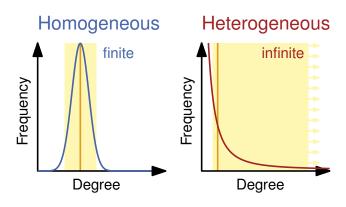


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 $lacksquare X \sim \mathsf{Par}(\alpha, x_{\mathsf{min}})$

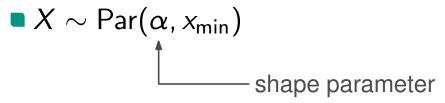


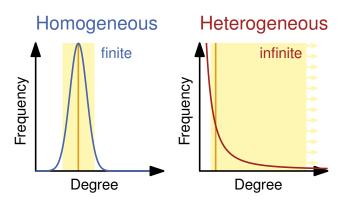


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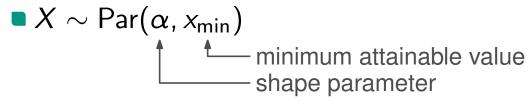


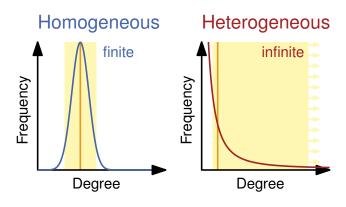


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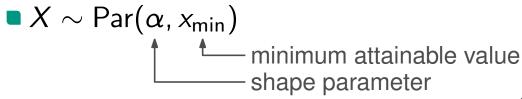


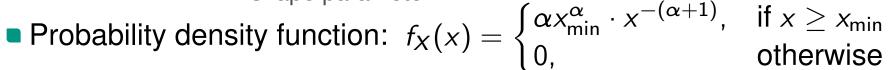


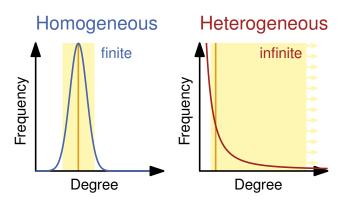
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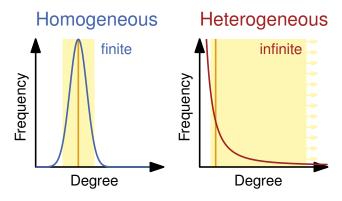


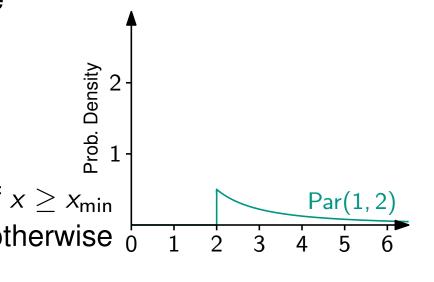
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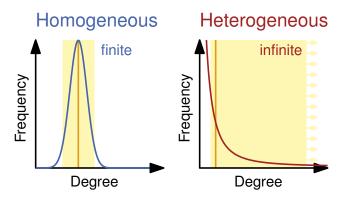


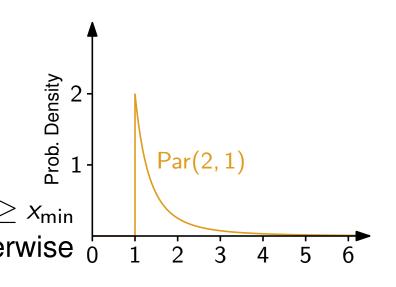
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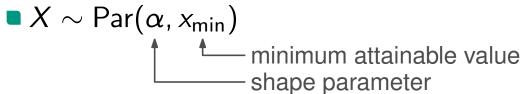


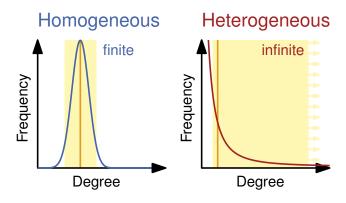


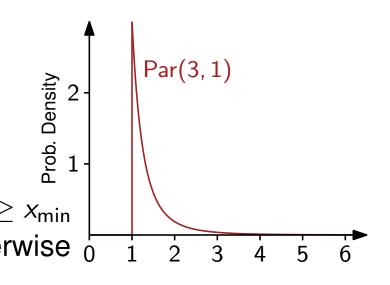
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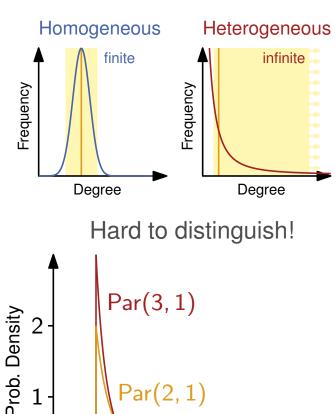


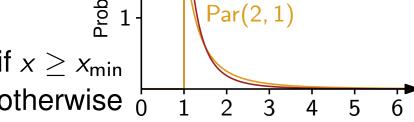
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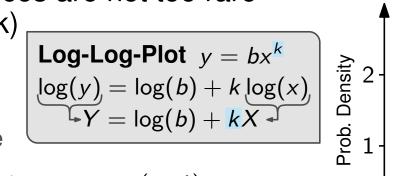
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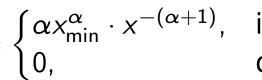
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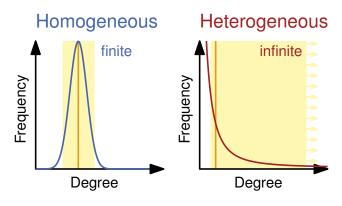
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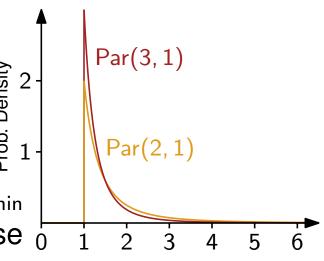
Probability density function: $f_X(x) = \begin{cases} \alpha x_{\min}^{\alpha} \cdot x^{-(\alpha+1)}, & \text{if } x \geq x_{\min} \\ 0. & \text{otherwise} \end{cases}$







Hard to distinguish!





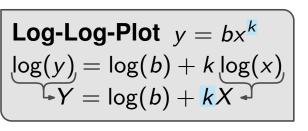
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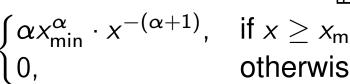
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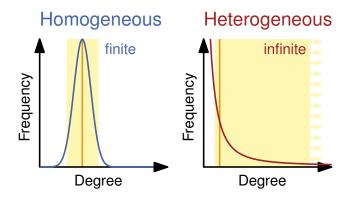
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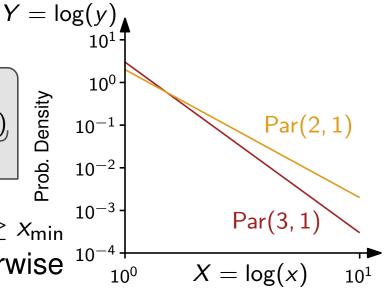
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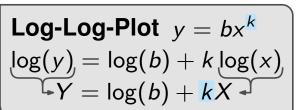
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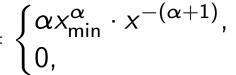
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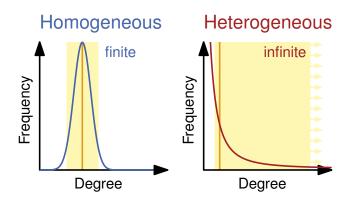
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konect.cc/plot/degree.a.youtube-links.full.png YouTube Frequency 01

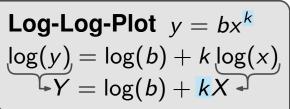


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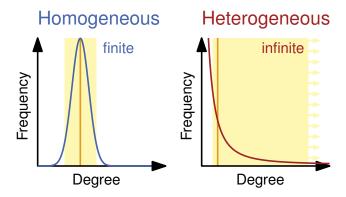
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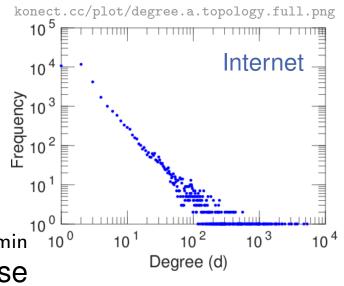
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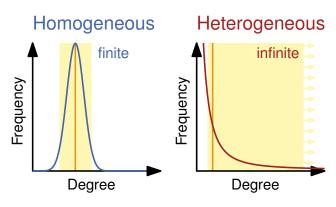
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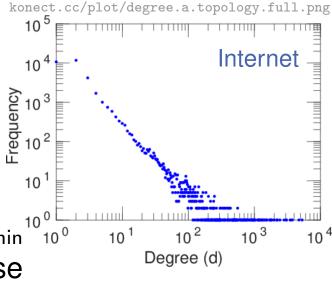
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Exercise: Determine for which values of α we have $\mathbb{E}[X] < \infty$ but $\text{Var}[X] = \infty$

Log-Log-Plot $y = bx^k$

 $\log(y) = \log(b) + k \log(x)$ $Y = \log(b) + kX$



Continuous Distributions

- For our purposes they are handled like discrete versions (replacing sums with integrals)
- Seen today: Uniform distribution, exponential distribution, Pareto distribution, joint distributions



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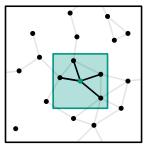
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- Vertices distributed at random in metric space
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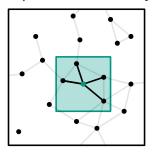
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(not discussed in lecture)

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Outlook: More realistic extension of RGGs featuring a heterogeneous degree distribution

