

Probability & Computing

Concentration



Expectation Management

What does it mean?

- “QuickSort has an *expected* running time of $O(n \log(n))$.”

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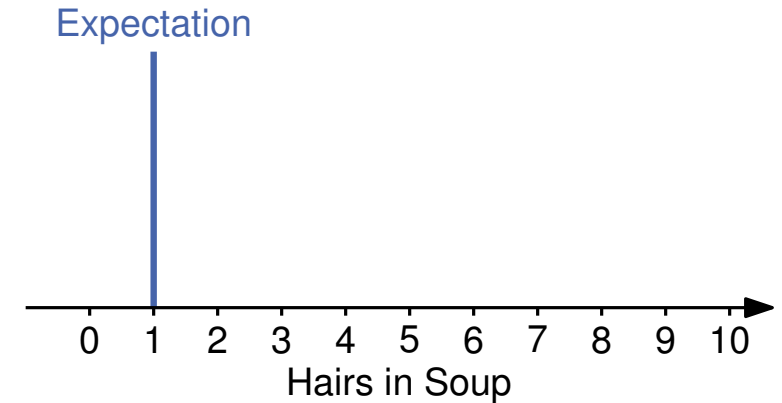
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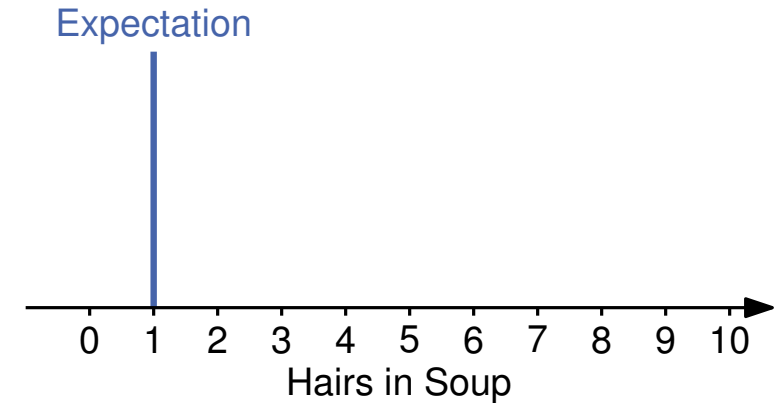
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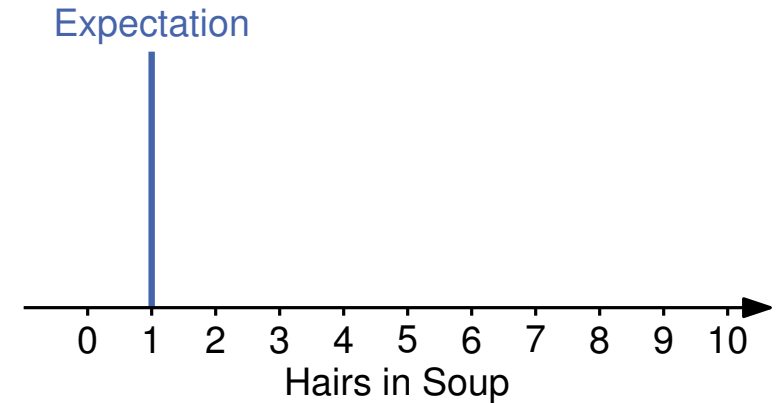
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- The average of infinitely many trials
- How useful is that information in practice?



I “expect” the sniper to hit the target...

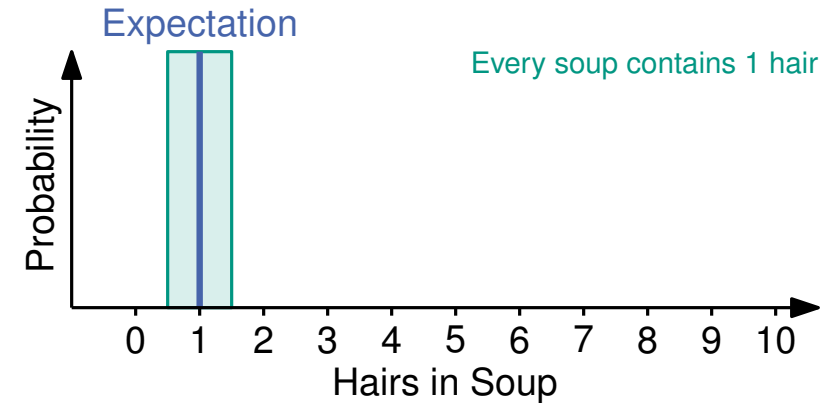
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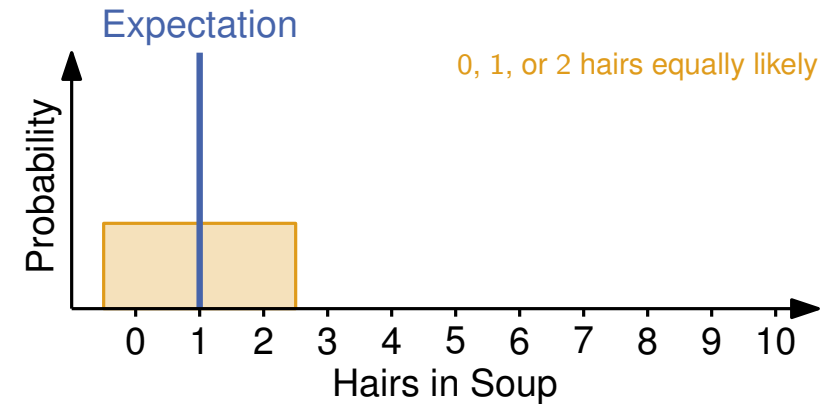
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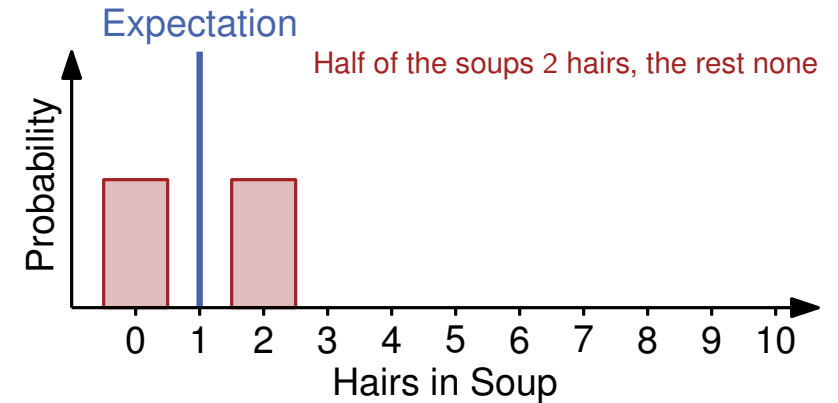
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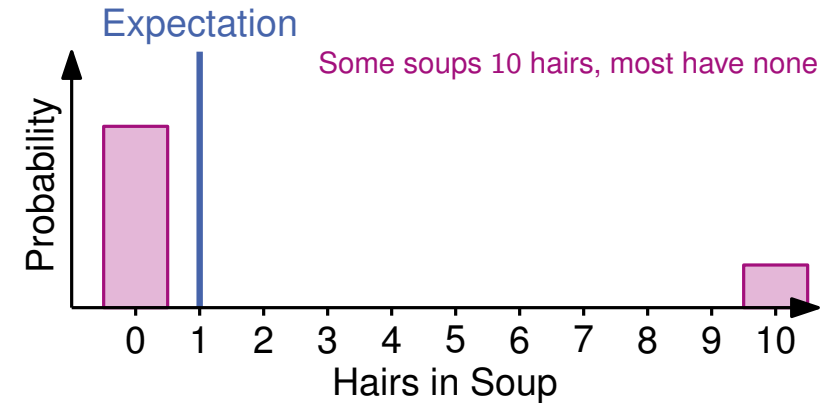
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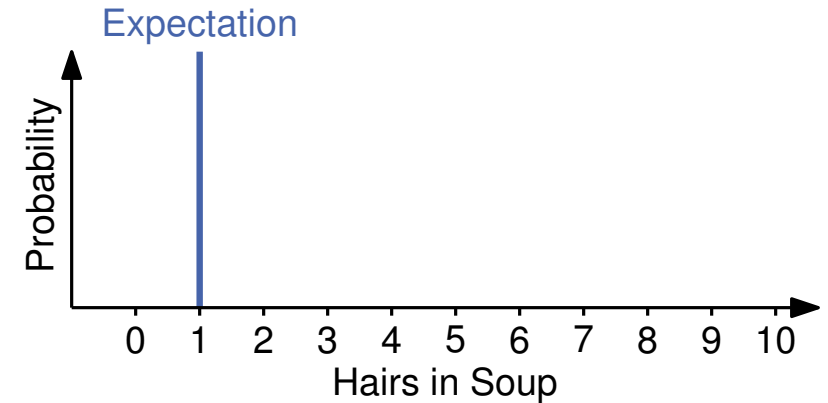
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Knowing that the expected value is 1 hair:
How likely is it that I get at least 10?

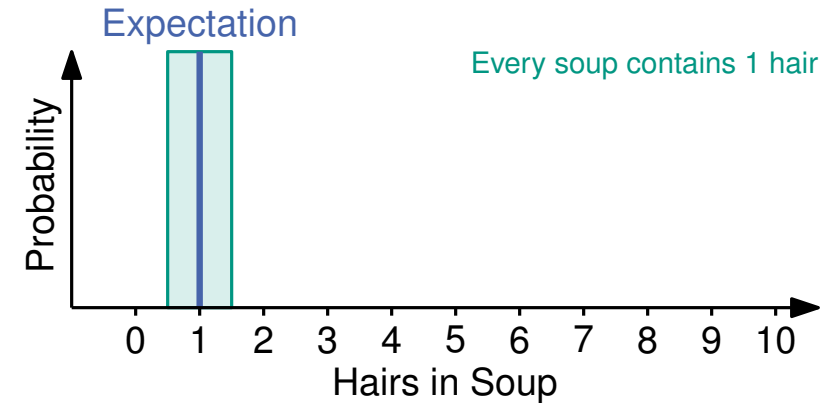
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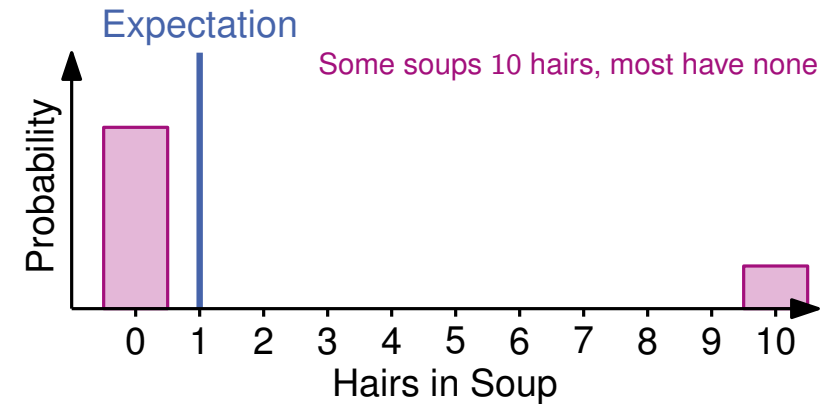
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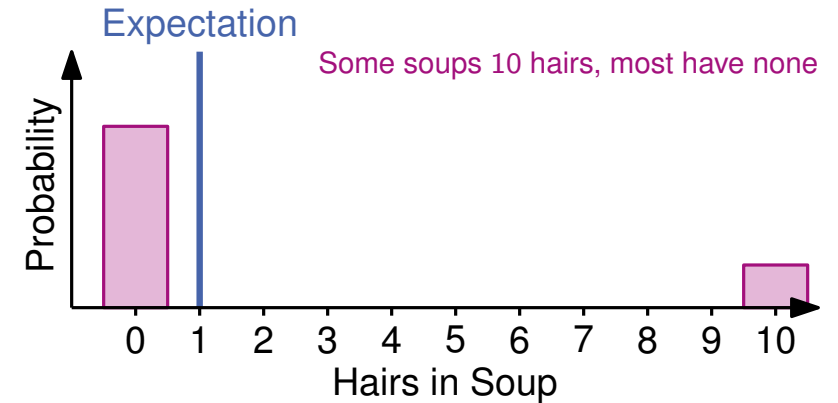
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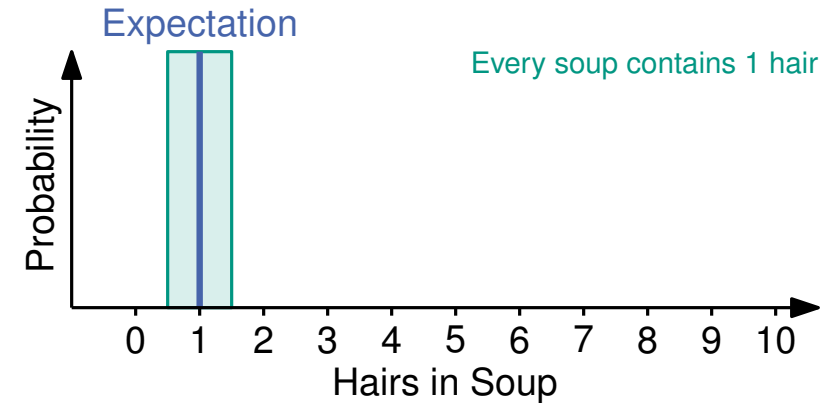
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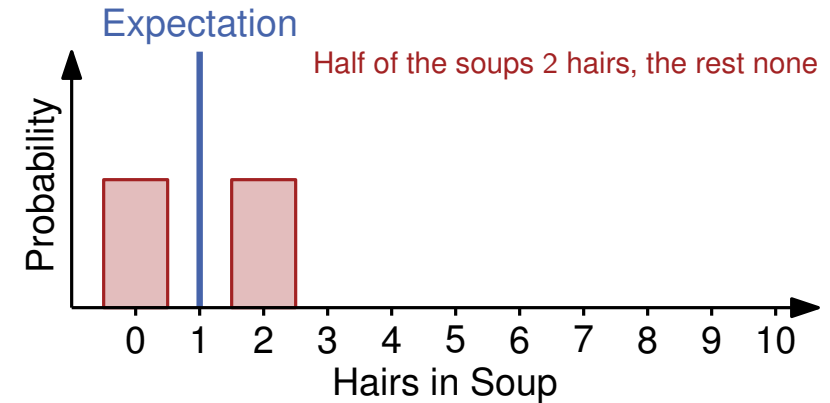
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- How useful is that information in practice?
- Does not tell us much about the shape of the distribution



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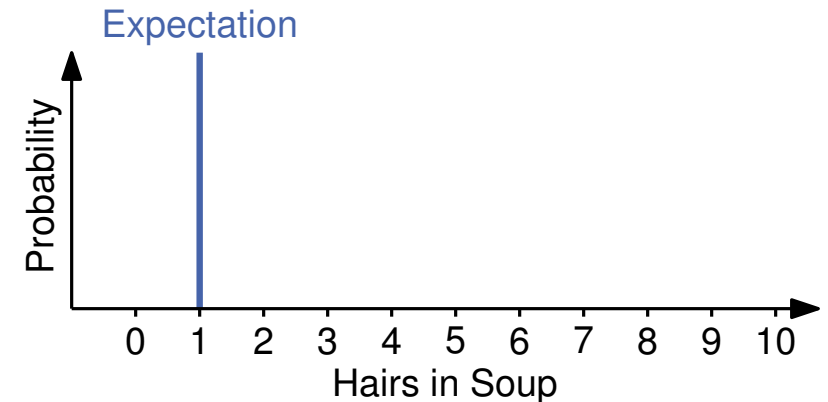
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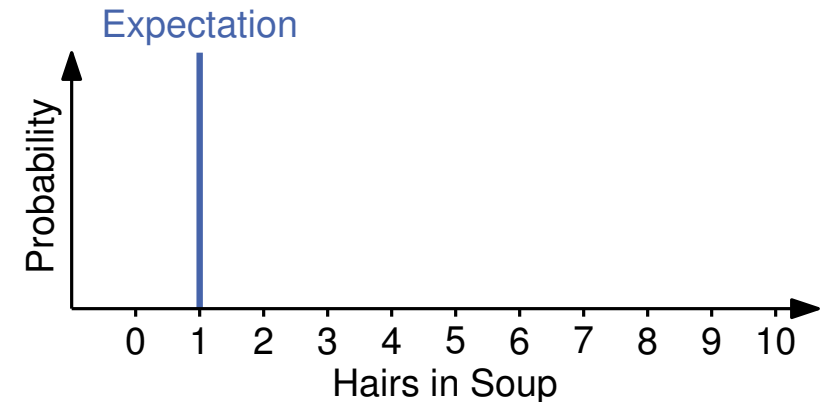
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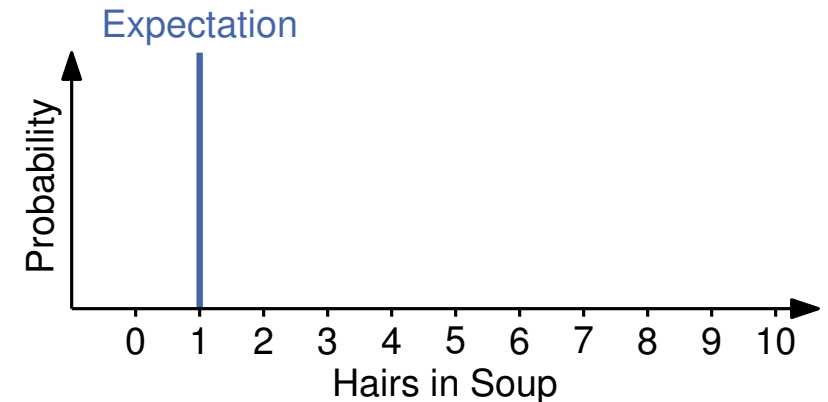
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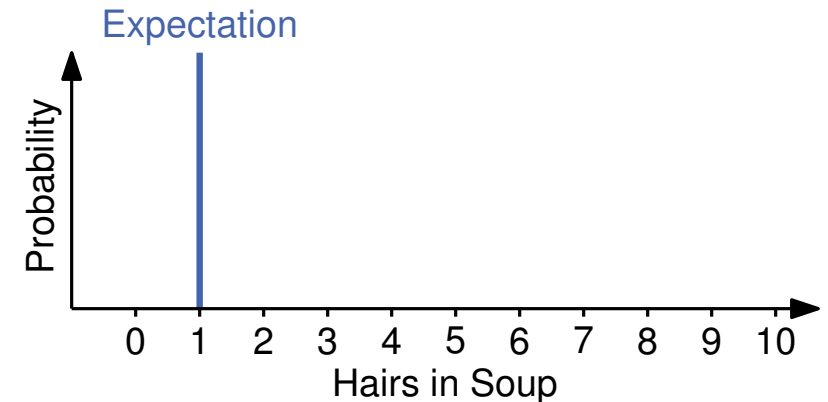
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Definition: A **concentration inequality** bounds the probability of a random variable to deviate from a given value (typically its expectation) by a certain amount.



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Markov's Inequality

About Markov

- Andrei “The Furious” Andreyevich Markov (Russian mathematician)
- Unhappy with the state of living at the time (1921)
- Informed St. Petersburg Academy of Sciences that he could not attend meetings due to not having shoes
- After getting shoes from the Communist Party he replied:

Finally, I received footwear. However, it is stupidly stitched together and does not accord with my measurements. Thus I cannot attend the meetings. I propose placing the footwear in a museum, as an example of the material culture of the current time.

“Shape, The hidden geometry of absolutely everything”, Jordan Ellenberg

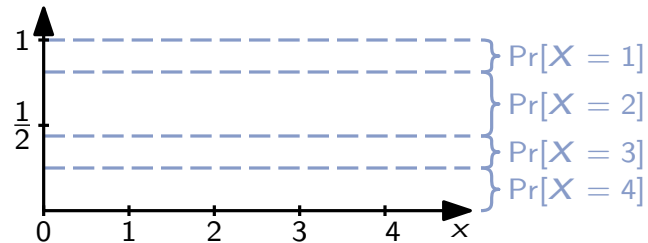
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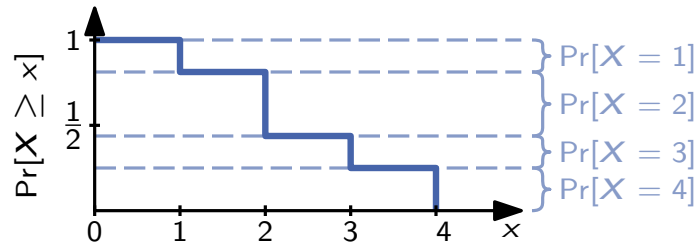
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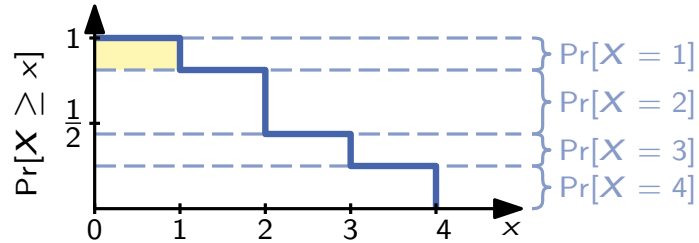
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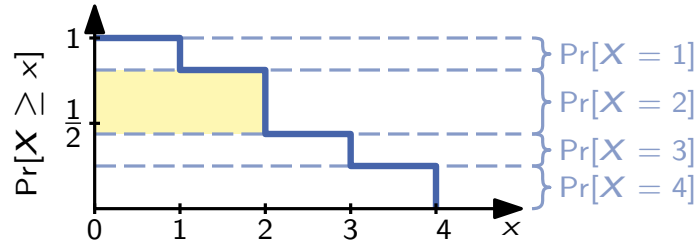


$$x \cdot \Pr[X = x]$$

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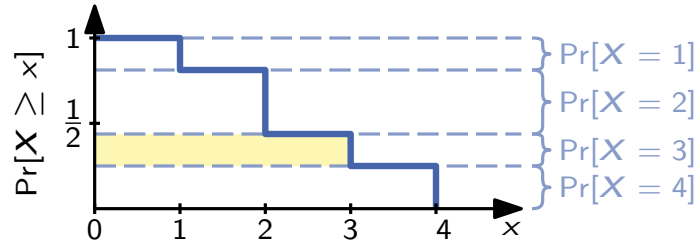


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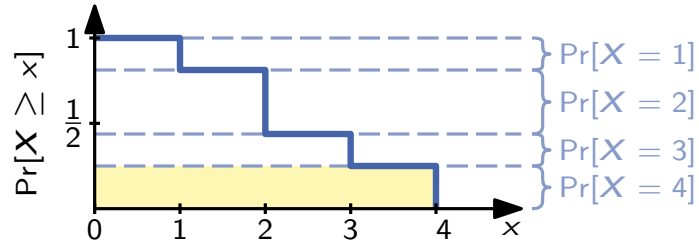


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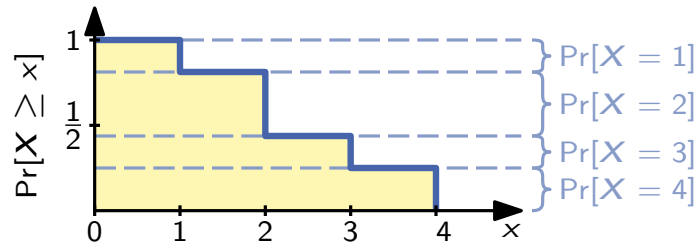


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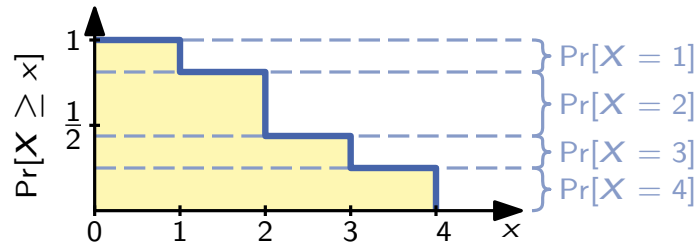


$$\sum_x x \cdot \Pr[X = x]$$

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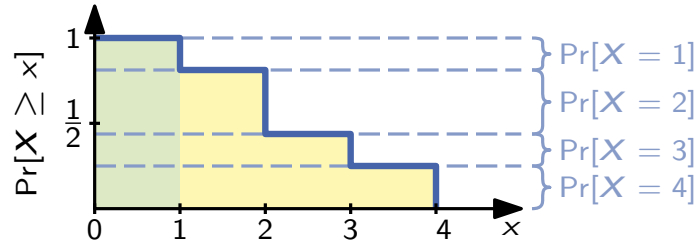


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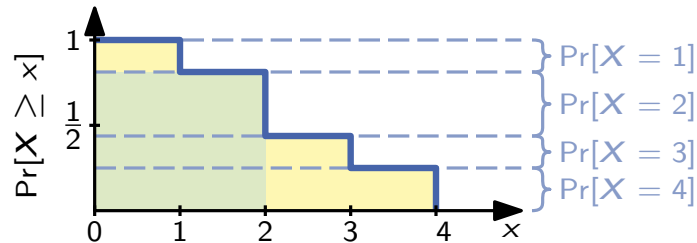
$$1 \cdot \Pr[X \geq 1]$$

fits into

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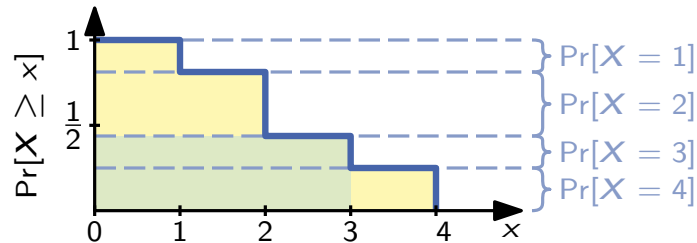
$$2 \cdot \Pr[X \geq 2]$$

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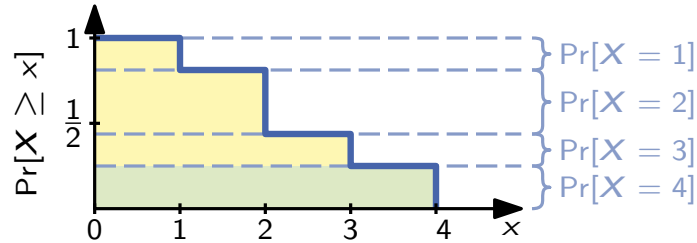
$$3 \cdot \Pr[X \geq 3]$$

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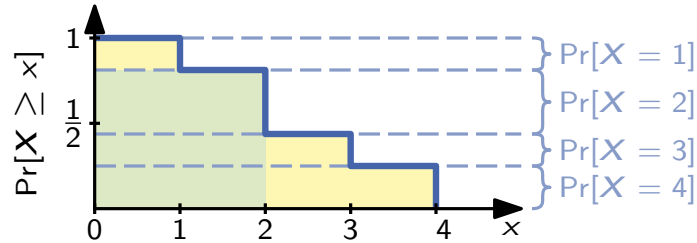
$$4 \cdot \Pr[X \geq 4]$$

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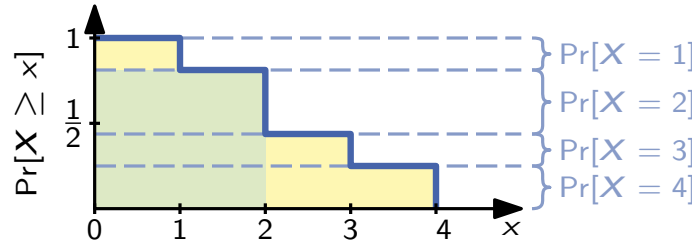
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← fits into →

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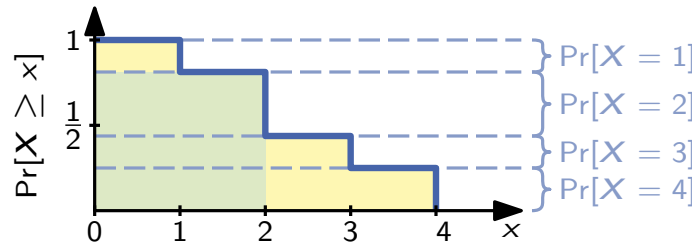
Proof
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Law of Total Expectation

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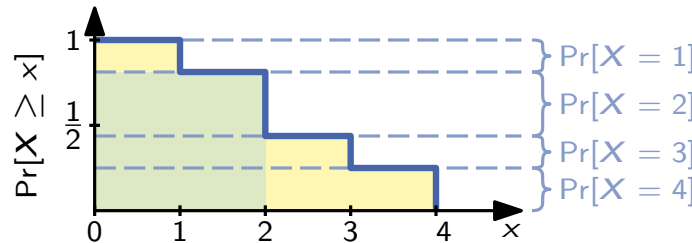
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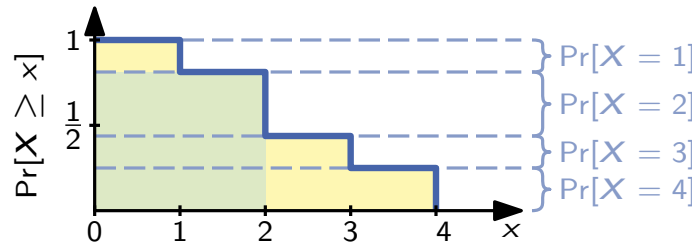
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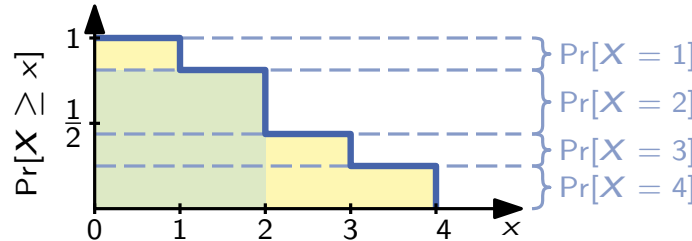
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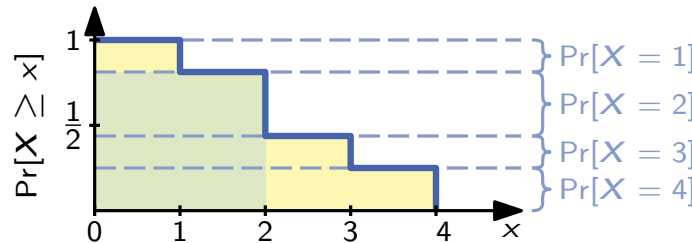
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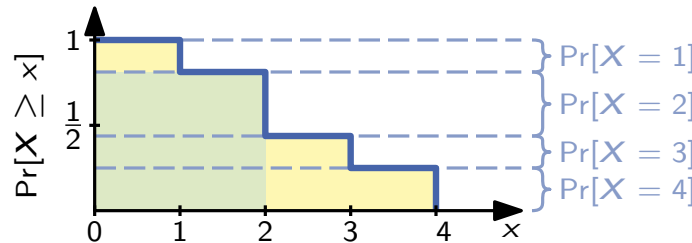
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Corollary: Let X be a non-negative rand. var. and $a > 0$. Then, $\Pr[X \geq a \cdot \mathbb{E}[X]] \leq 1/a$.

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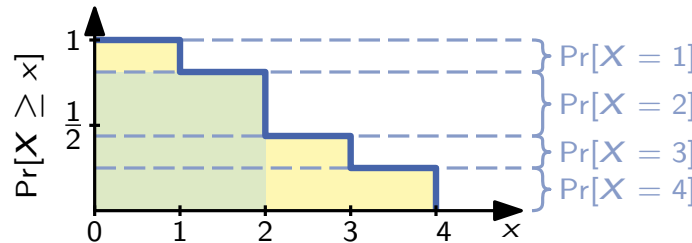
Corollary: Let X be a non-negative rand. var. and $a > 0$. Then, $\Pr[X \geq a \cdot \mathbb{E}[X]] \leq 1/a$.

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 - How likely is it that I get less than 2?

Markov's Inequality

Theorem (Markov's inequality): Let X be a non-negative random variable and let $a > 0$. Then, $\Pr[X \geq a] \leq \mathbb{E}[X]/a$.

Visual Proof



$$\mathbb{E}[X] = \sum_x x \cdot \Pr[X = x] \geq a \cdot \Pr[X \geq a]$$

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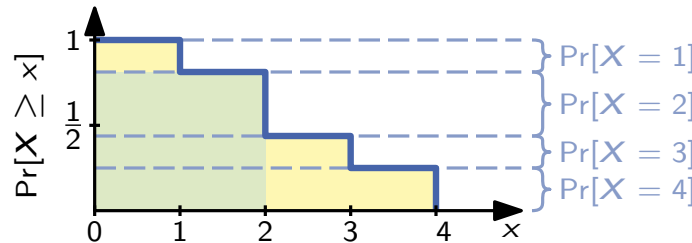
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$$\Pr[X < 2] = 1 - \Pr[X \geq 2] \geq 1 - 1/2 = 1/2$$

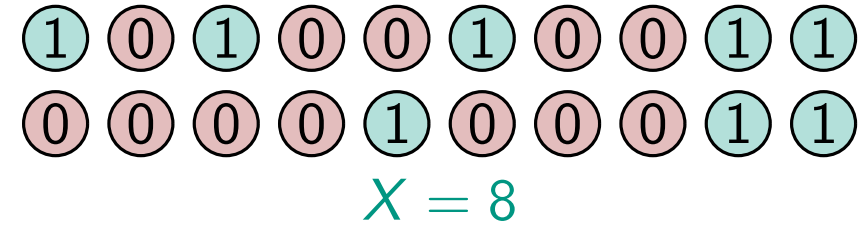
Oh no...

Application: Unfair Coins

- The sum of 20 unfair $\{0, 1\}$ -coin tosses: $X \sim \text{Bin}(20, \frac{1}{5})$

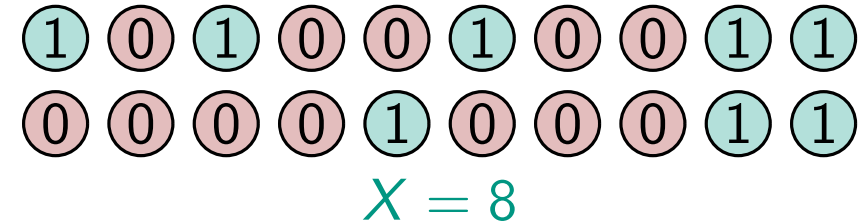
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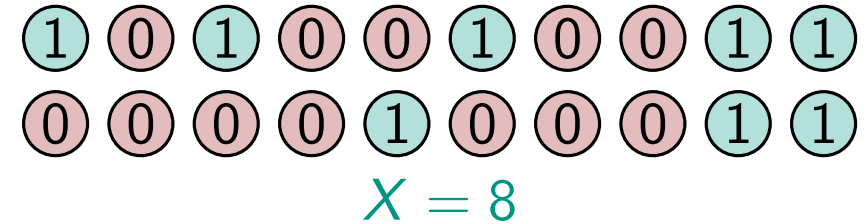
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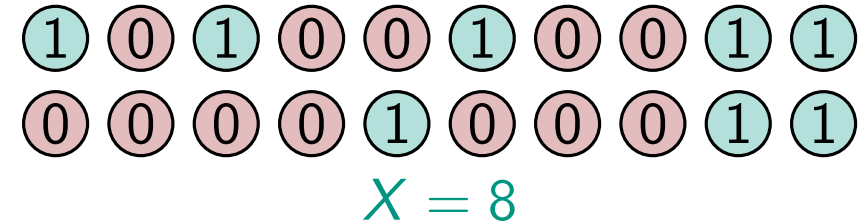


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$$\Pr[X \geq 16] \leq \mathbb{E}[X]/16$$

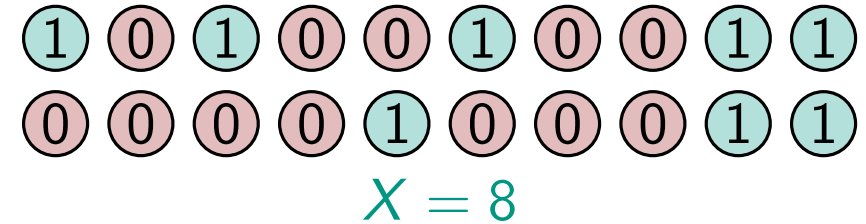


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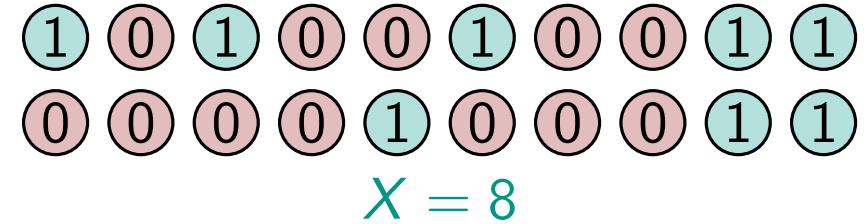


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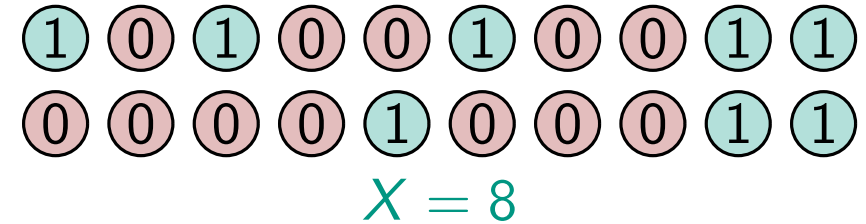
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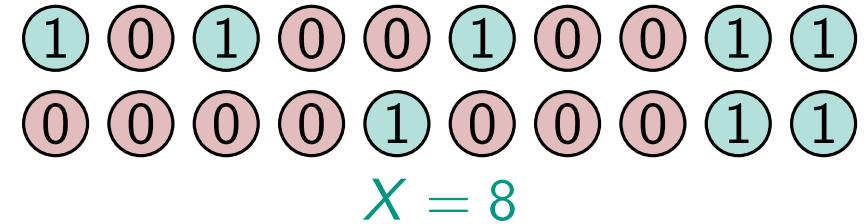
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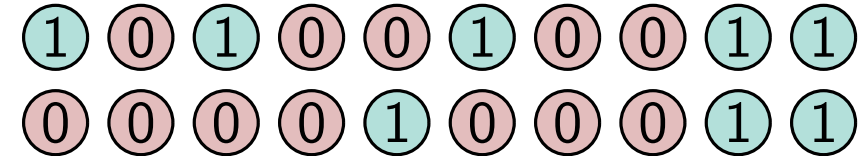
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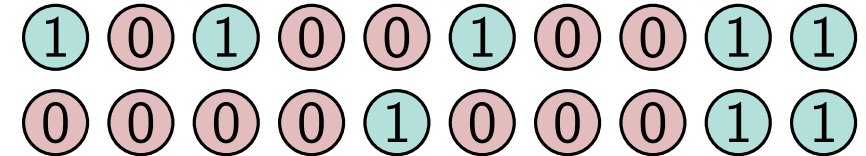
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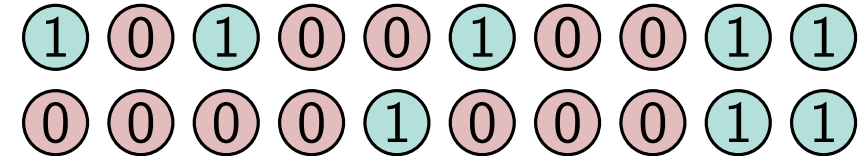
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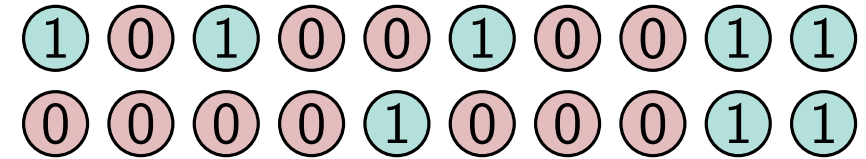
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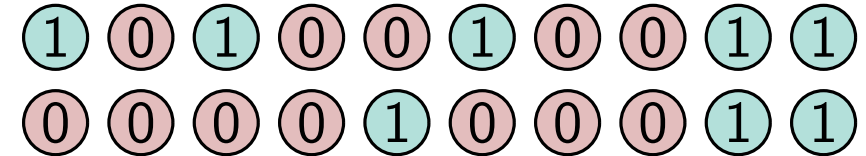
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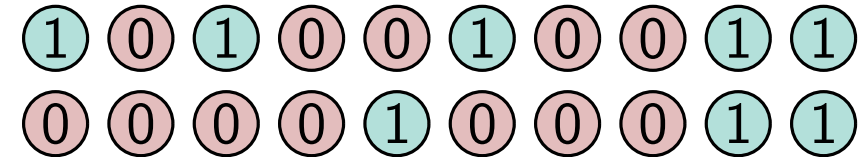
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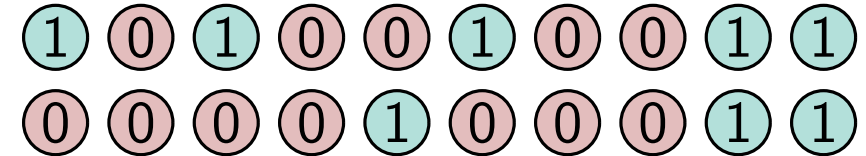
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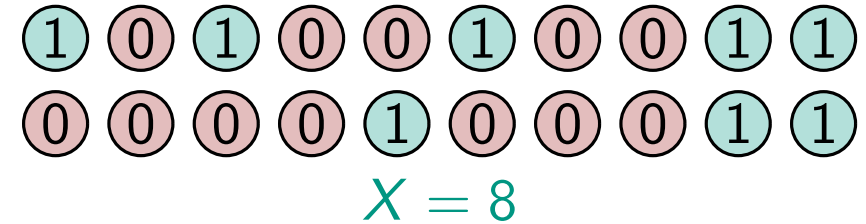
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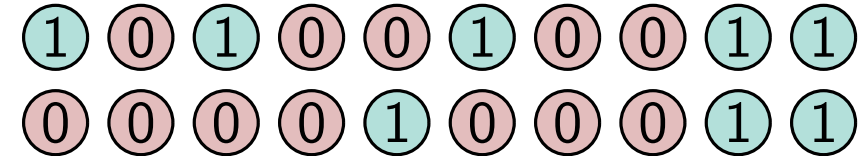
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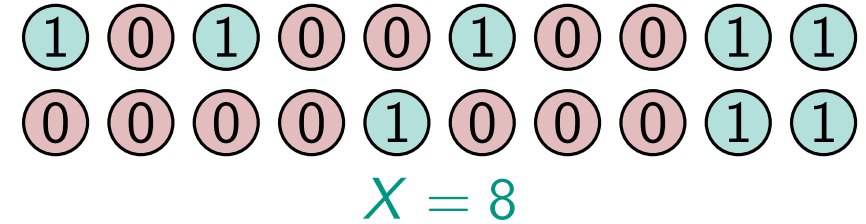
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\Rightarrow There is no better bound (that relies only on the expected value)

We need more information about the shape of the distribution!

Characterizing the Shape of a Distribution

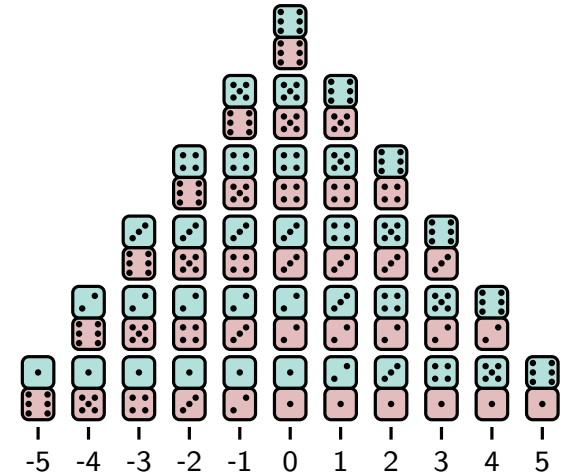
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Example

- X, Y independent fair die-rolls, $D = X - Y$

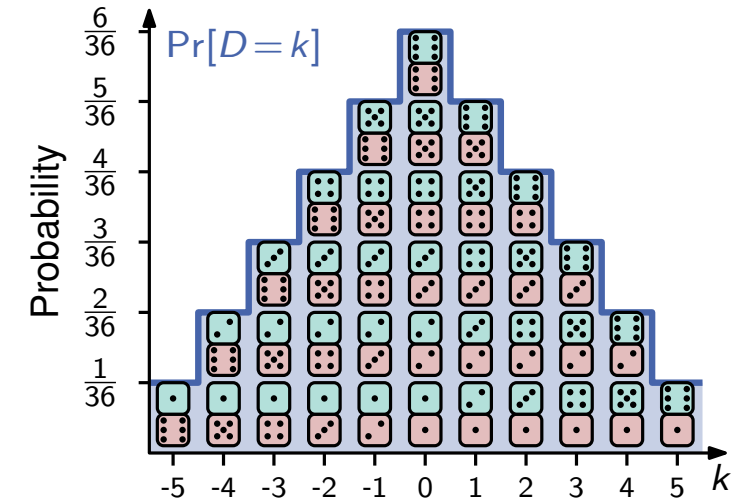


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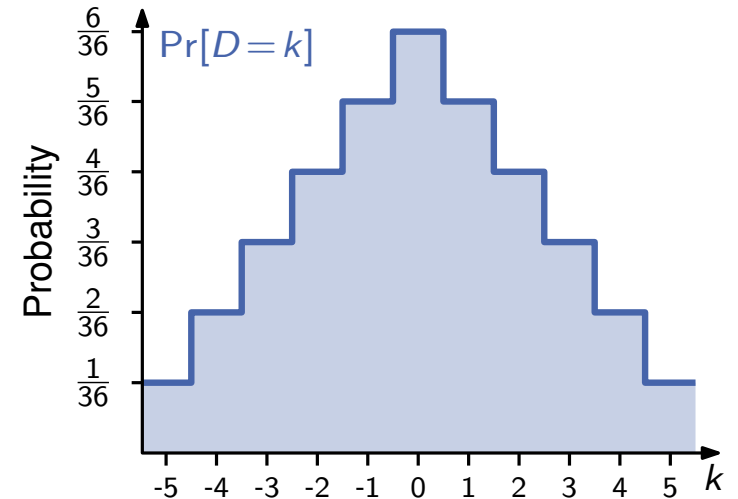


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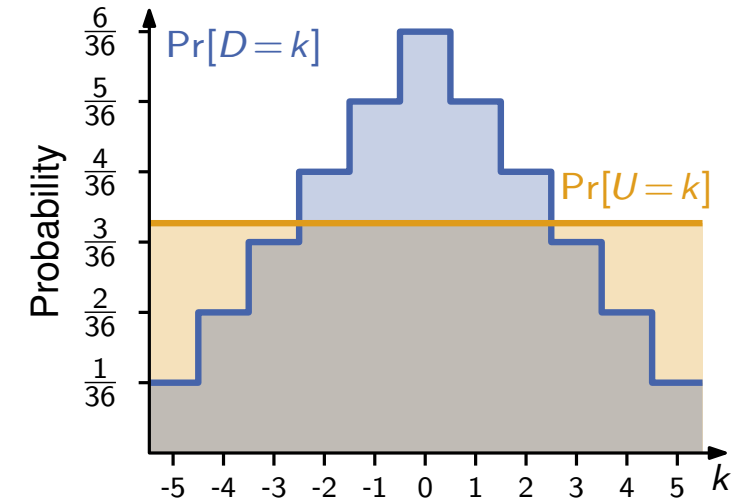


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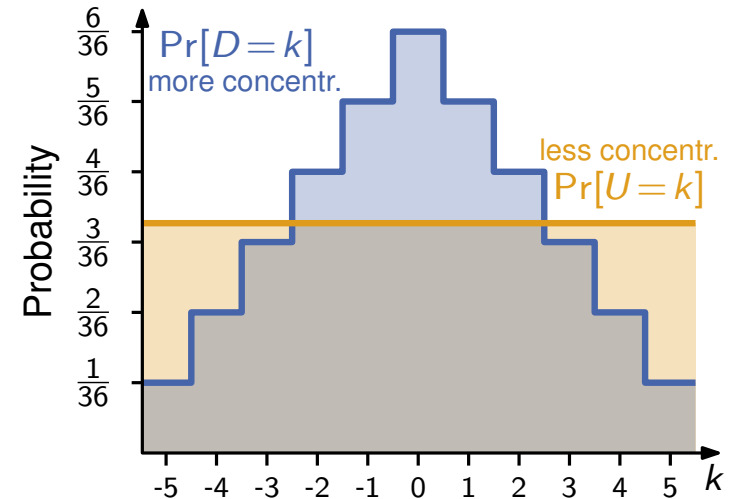


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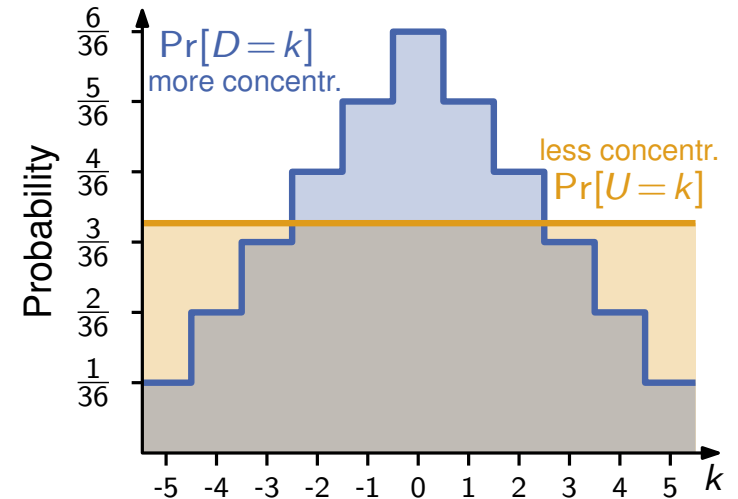


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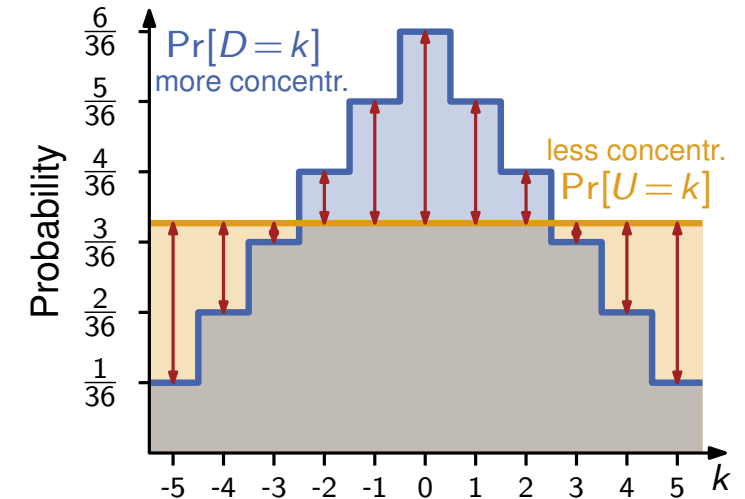


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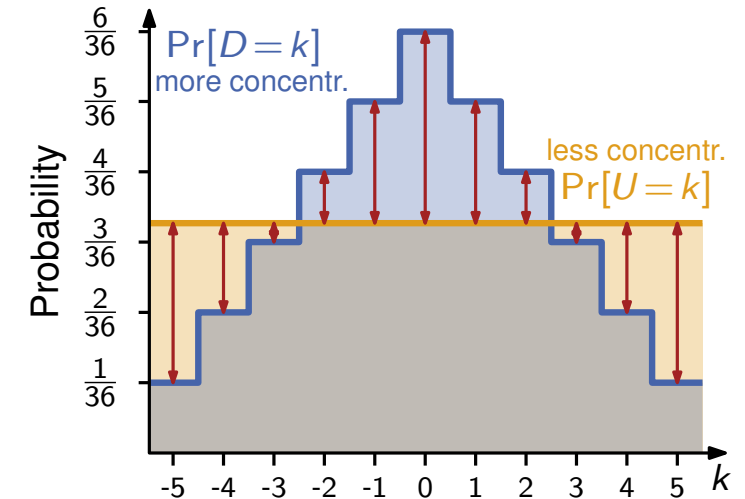


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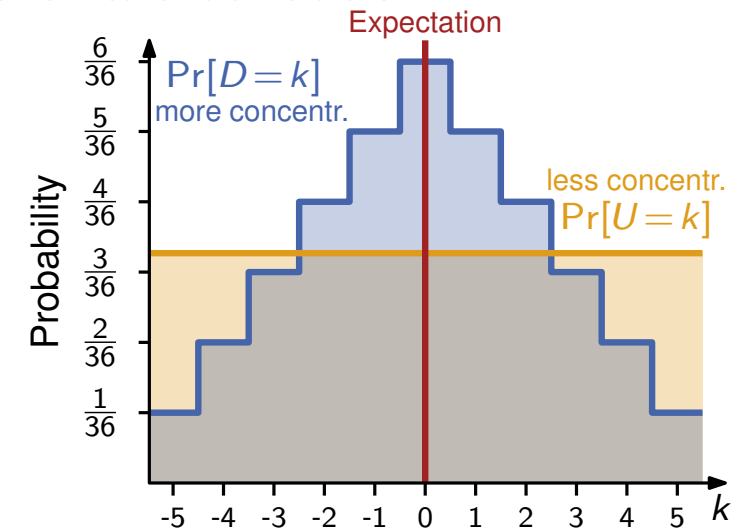
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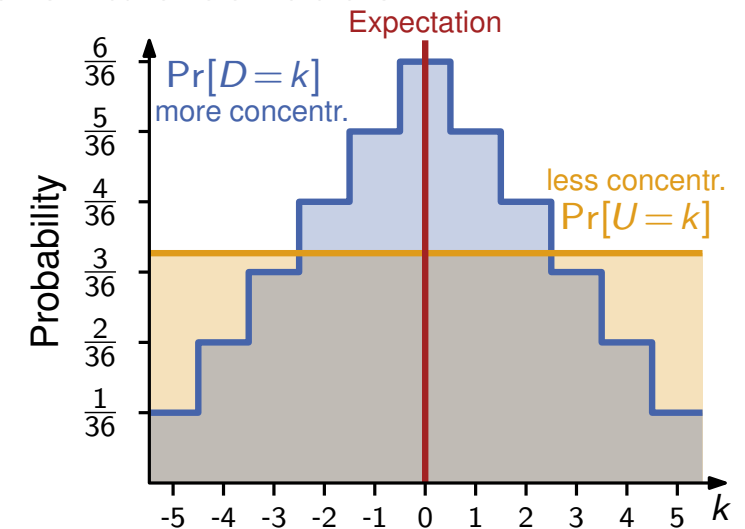
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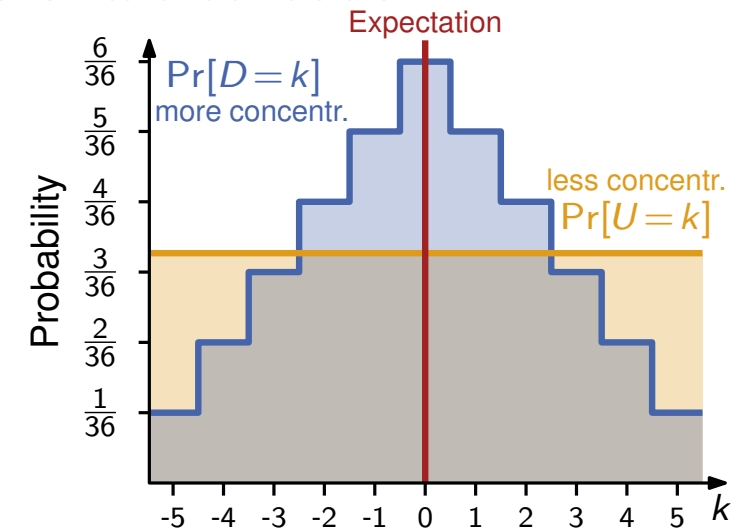
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Characterizing the Shape of a Distribution

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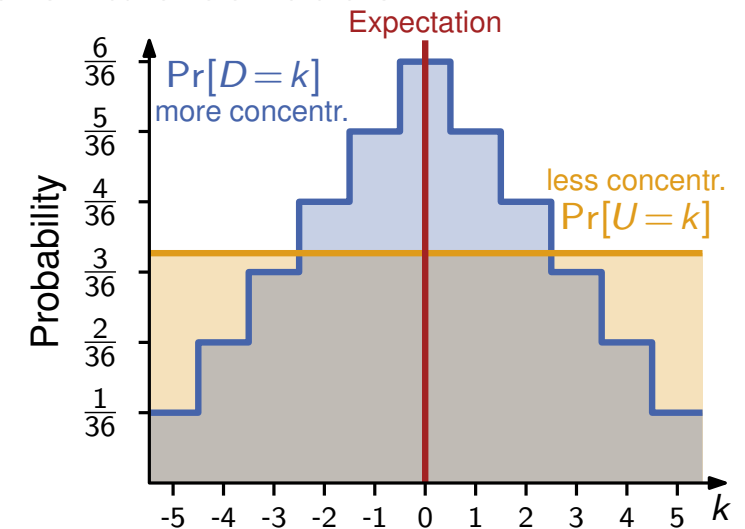
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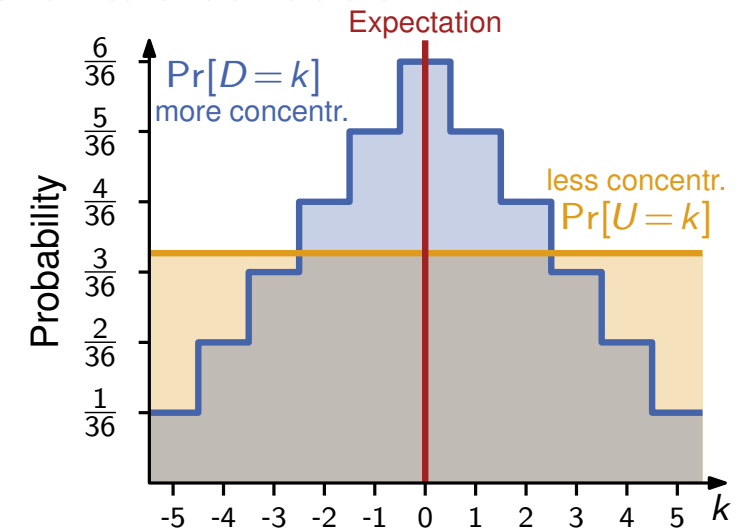
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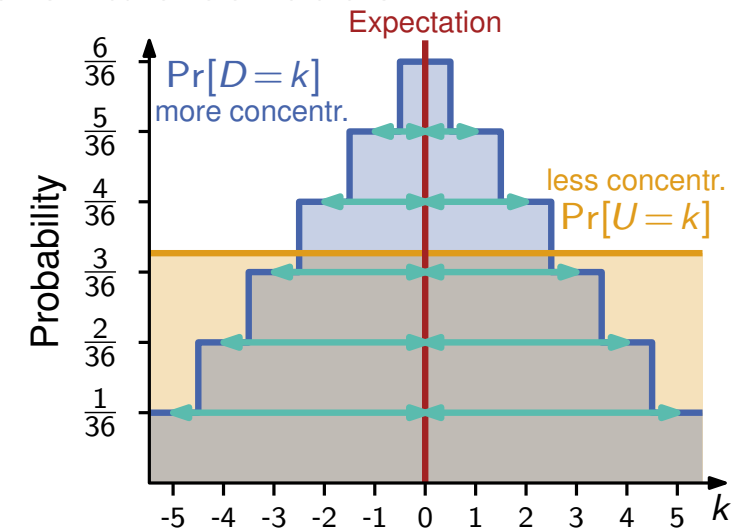
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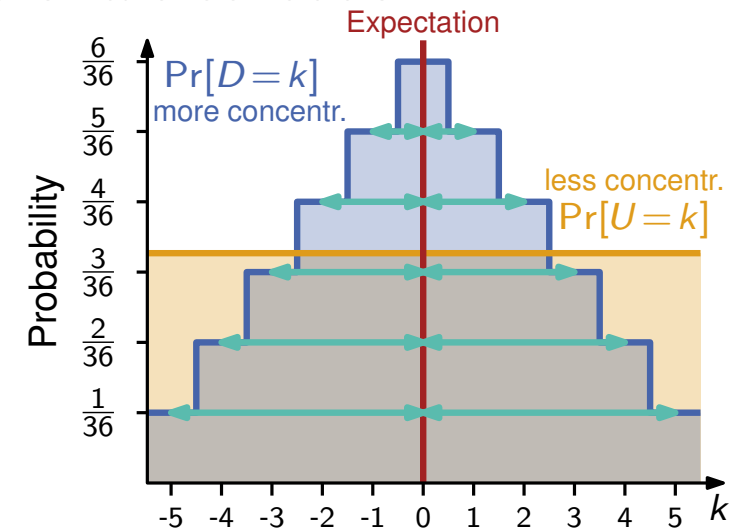
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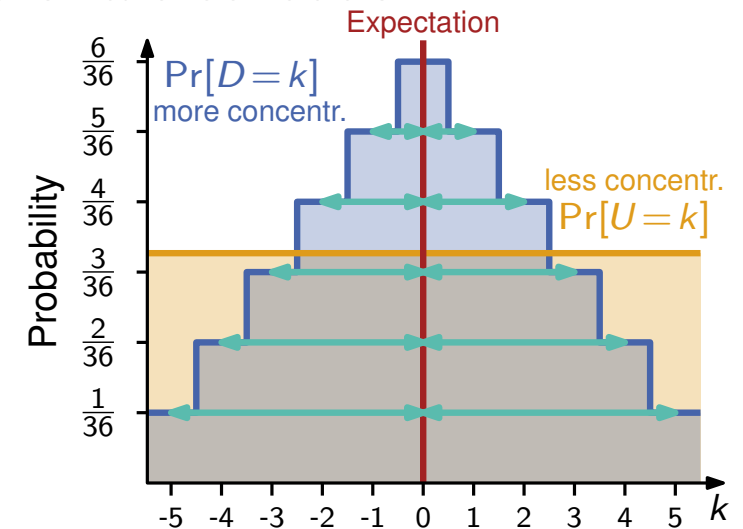
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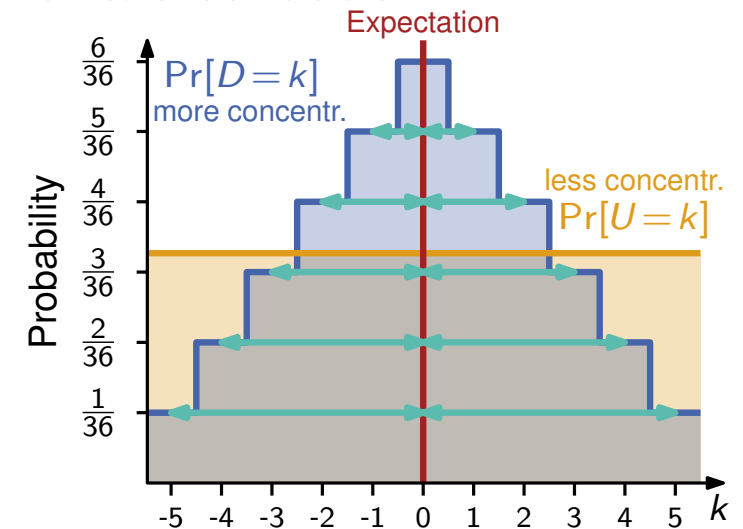
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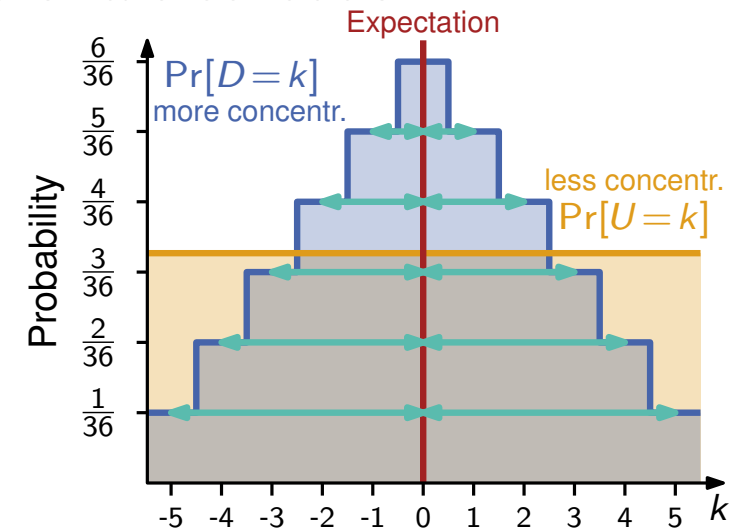
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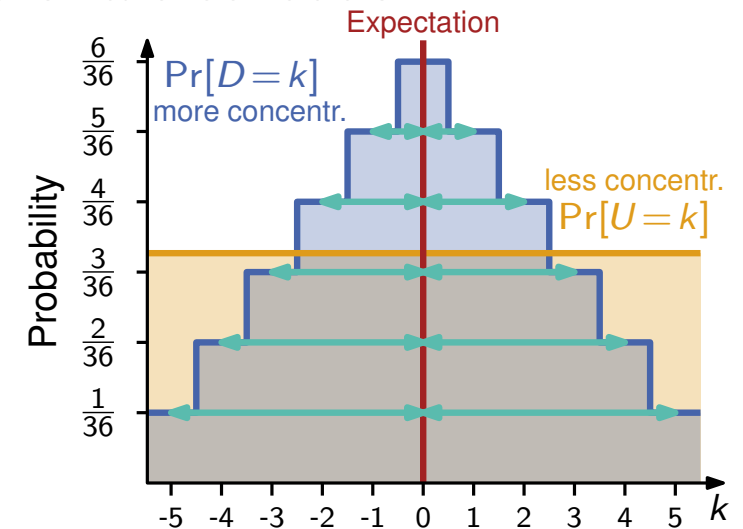
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Squared distance to \mathbb{E}



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$$f(k) = k$$

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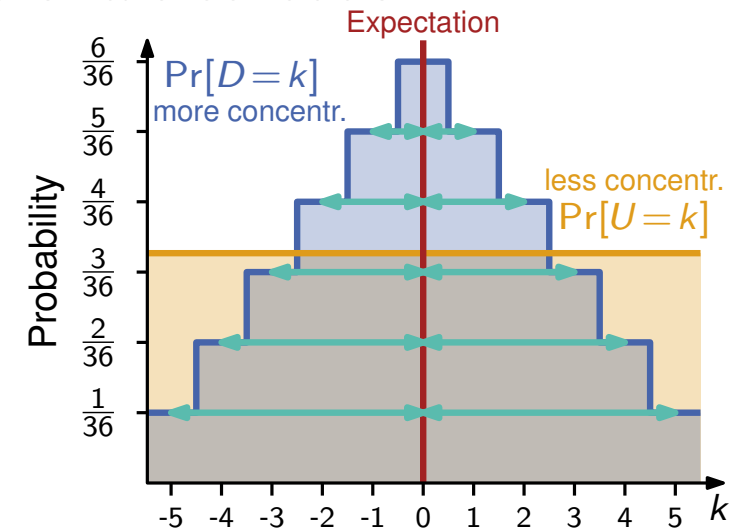
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These are just expectations of functions of random variables!

Do you have a Moment?

Expectation and Functions

- Random variable X taking values in a set S
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... and with Markov's help, we can turn that insight into a concentration inequality!

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Markov's teacher! (Markov's inequality actually appeared earlier in Chebychev's works)

Theorem (Chebychev's inequality): Let X be a random variable with finite variance and let $b > 0$. Then, $\Pr[|X - \mathbb{E}[X]| \geq b] \leq \text{Var}[X]/b^2$.

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Application: Unfair Coins



- $X \sim \text{Bin}(20, \frac{1}{5}), \Pr[X \geq 16]? \mid \mathbb{E}[X] = 20 \cdot \frac{1}{5} = 4$

$$\Pr[X \geq 16] = \sum_{k=16}^{20} \binom{20}{k} \left(\frac{1}{5}\right)^k \cdot \left(1 - \frac{1}{5}\right)^{20-k} \approx 0.0000000138$$

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Application: Unfair Coins



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Order of magnitude better than Markov!

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- Empirical distribution of the degrees of *all* vertices in a graph $G = (V, E)$

$$N_d = \sum_{v \in V} \mathbb{1}_{\{\deg(v)=d\}} \quad (\text{normalized: } \frac{1}{n} N_d, \text{ for } n = |V|)$$

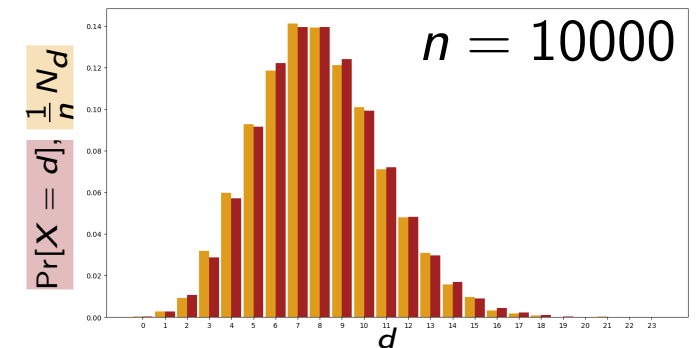
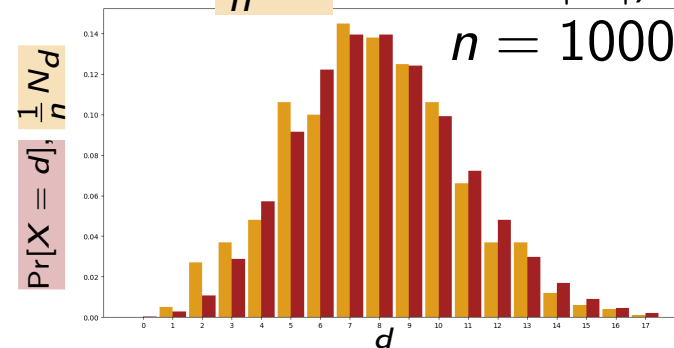
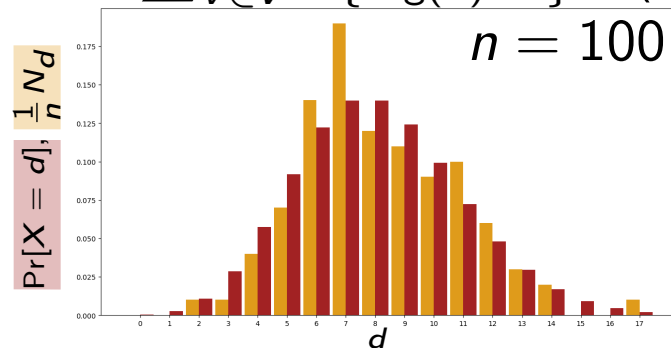
Application: ER – Degree Distribution

Recap

- $G(n, p)$: Start with n nodes, connect any two with fixed probability p , independently
- Probability distribution of the degree of a *single* node v : $\deg(v) \sim \text{Bin}(n - 1, p)$
- For $p = c/n$ with $c \in \Theta(1)$ the degree of a vertex is approximately Poisson-distributed
 - Total variation distance of X, Y taking values in a set S :

$$d_{TV}(X, Y) = \frac{1}{2} \sum_{x \in S} |\Pr[X = x] - \Pr[Y = x]|$$
 - For $\lambda = -n \log(1 - p) = c + O(1/n)$ and $X \sim \text{Pois}(\lambda)$ we have $d_{TV}(\deg(v), X) = o(1)$
- Empirical distribution of the degrees of *all* vertices in a graph $G = (V, E)$

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Theorem: Consider a $G(n, p)$ with $p = c/n$ for constant $c > 0$. For $\lambda = -n \log(1 - p)$, let $X \sim \text{Pois}(\lambda)$. Then for all $d > 0$ and every $\varepsilon > 0$ we have

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Proof

- Step 1: $\Pr[X = d]$ is close to the expectation of $\frac{1}{n} N_d$ $\lim_{n \rightarrow \infty} \left| \Pr[X = d] - \mathbb{E} \left[\frac{1}{n} N_d \right] \right| = 0$

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Chebychev: X finite variance, $b > 0$
 $\Pr[|X - \mathbb{E}[X]| \geq b] \leq \text{Var}[X]/b^2$

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$$N_d \in \{0, \dots, n\}$$

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$$\Pr \left[\left| \mathbb{E} \left[\frac{1}{n}N_d \right] - \frac{1}{n}N_d \right| \geq \varepsilon \right] \leq \underbrace{\text{Var} \left[\frac{1}{n}N_d \right]}_{\downarrow} / \varepsilon^2$$

$$\begin{aligned} \text{Var} \left[\frac{1}{n}N_d \right] &= \mathbb{E} \left[\left(\frac{1}{n}N_d \right)^2 \right] - \mathbb{E} \left[\frac{1}{n}N_d \right]^2 \\ &= \frac{1}{n^2} \left(\mathbb{E} \left[(N_d)^2 \right] - \mathbb{E} \left[N_d \right]^2 \right) \end{aligned}$$

$$\lim_{n \rightarrow \infty} \Pr \left[\left| \mathbb{E} \left[\frac{1}{n}N_d \right] - \frac{1}{n}N_d \right| \geq \varepsilon \right] = 0$$

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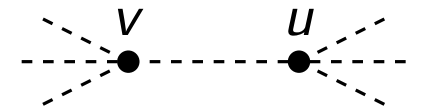
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Couplings

- Consider $\text{deg}(u)$ and $\text{deg}(v)$



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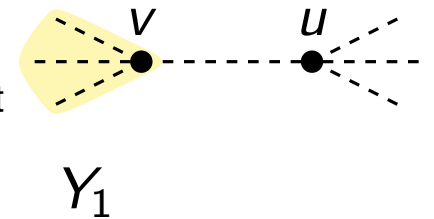
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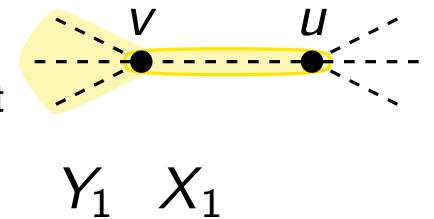
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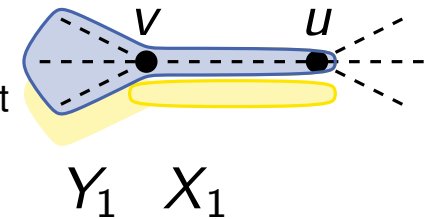
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$$\text{deg}(v)$$

$$\simeq$$

$$X_1 + Y_1$$

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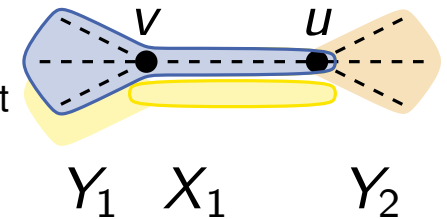
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$$\text{deg}(v)$$

$$\supseteq$$

$$X_1 + Y_1$$

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Step 2: Concentration of $\frac{1}{n}N_d$

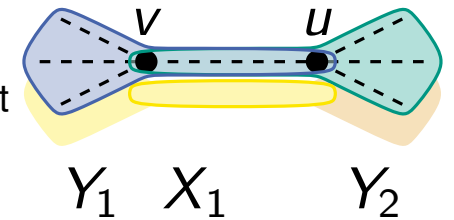
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$\text{deg}(v)$

$\text{deg}(u)$

\simeq

\parallel

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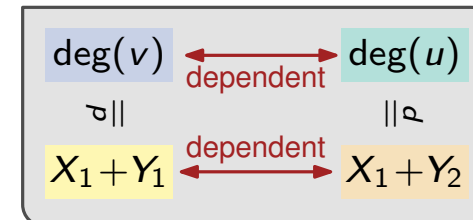
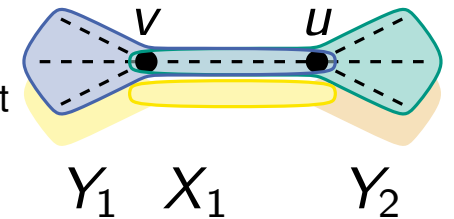
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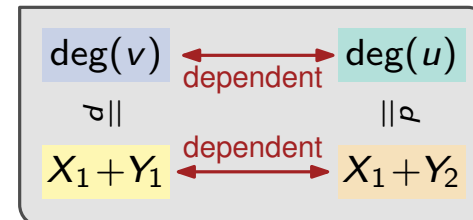
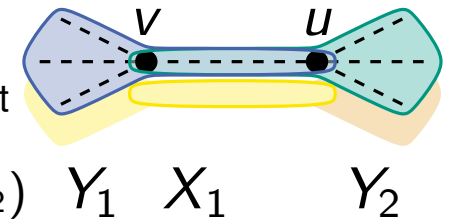
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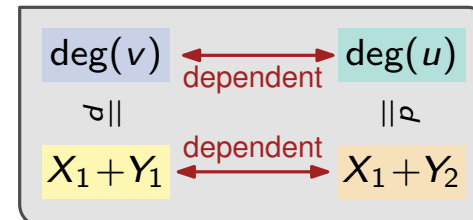
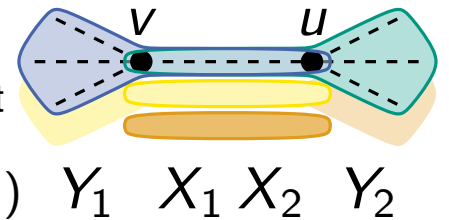
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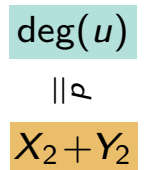
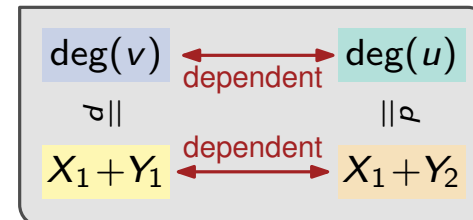
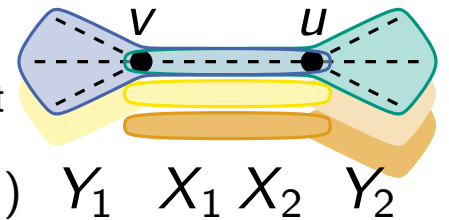
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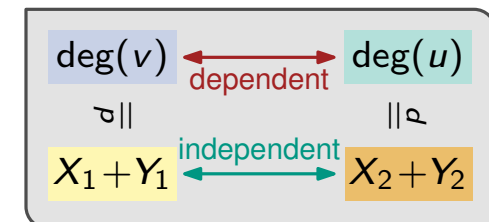
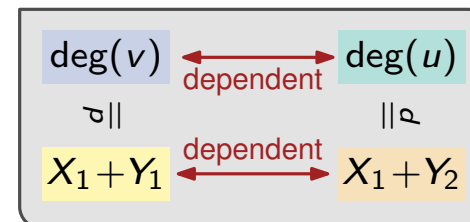
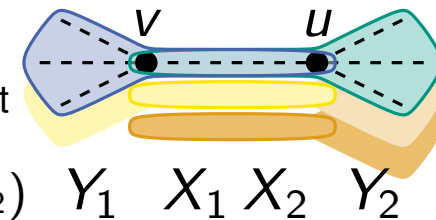
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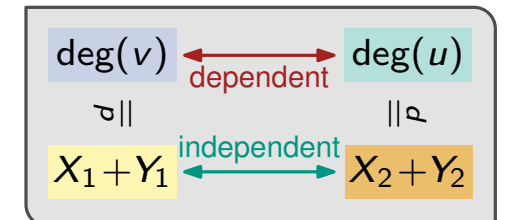
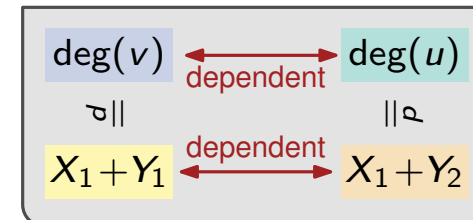
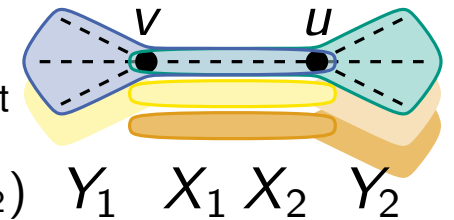
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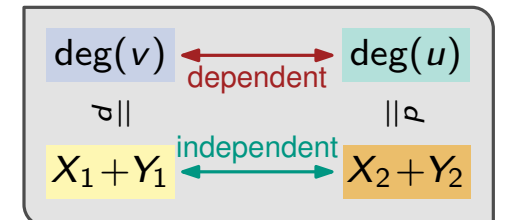
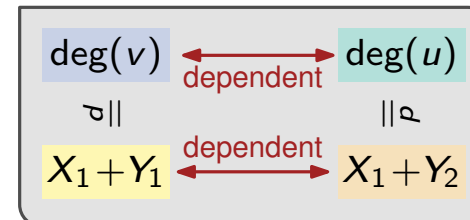
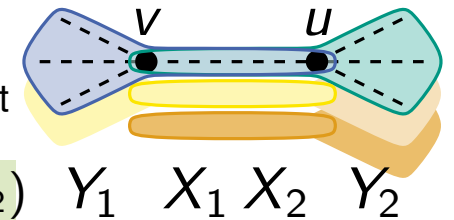
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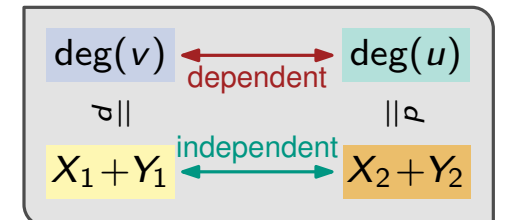
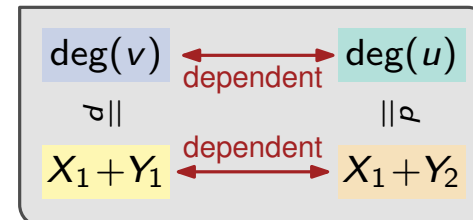
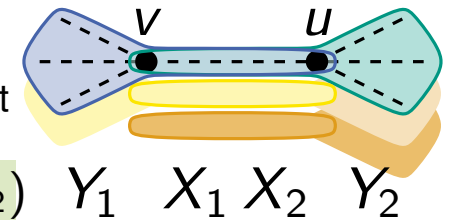
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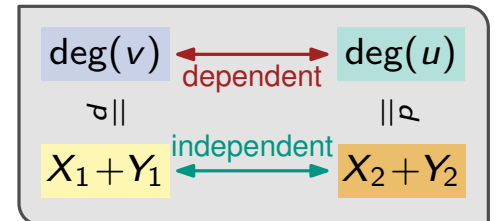
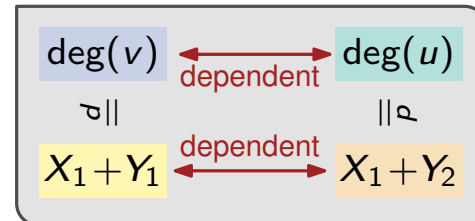
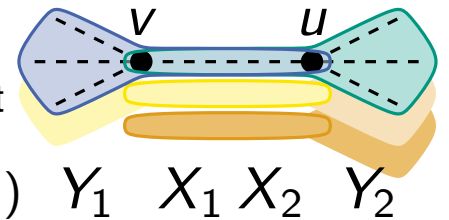
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Step 2: Concentration of $\frac{1}{n}N_d$

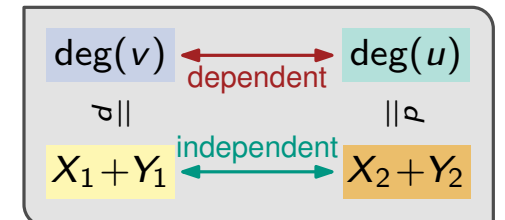
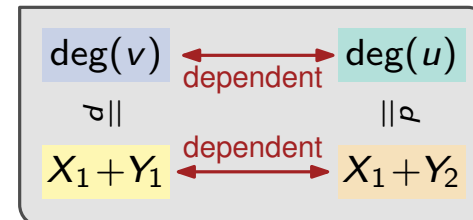
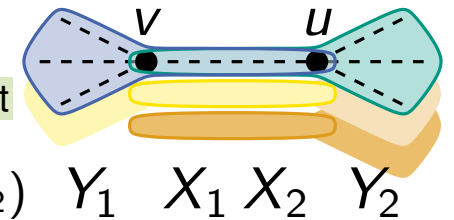
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$$\text{deg}(v) \stackrel{d}{=} \text{deg}(u)$$

Couplings

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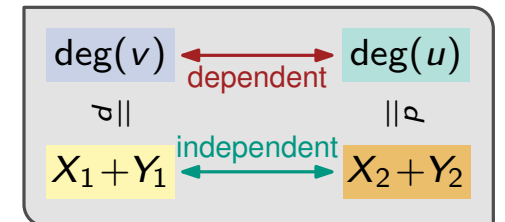
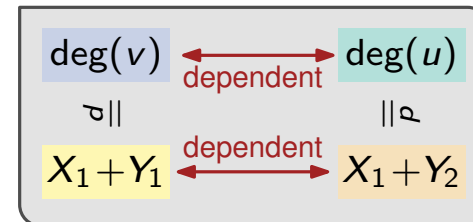
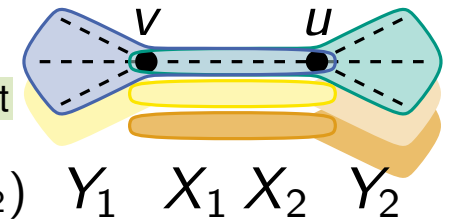
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Step 2: Concentration of $\frac{1}{n}N_d$

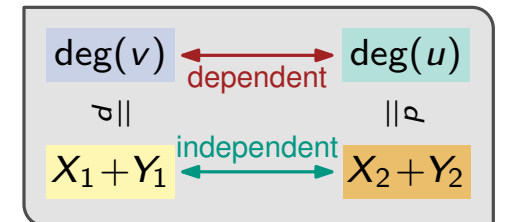
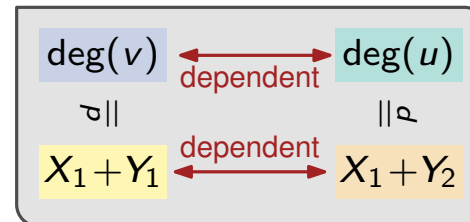
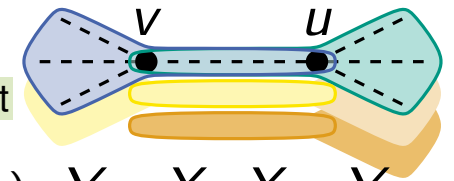
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Which excludes this from happening

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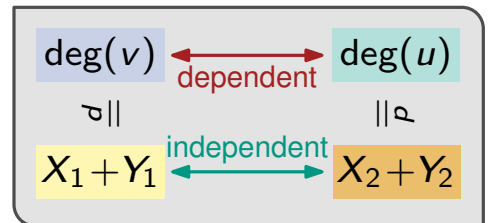
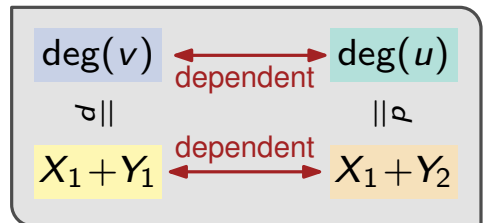
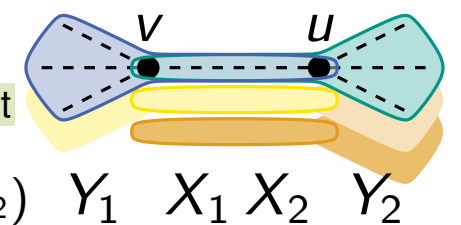
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Law of total probability

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Law of total probability

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$$\text{Var} \left[\frac{1}{n}N_d \right] = \mathbb{E} \left[\left(\frac{1}{n}N_d \right)^2 \right] - \mathbb{E} \left[\frac{1}{n}N_d \right]^2 \leq \frac{1}{n} + 2p = \frac{1}{n} + 2\frac{c}{n} \xrightarrow{n \rightarrow \infty} 0$$

$$= \frac{1}{n^2} \left(\mathbb{E} \left[(N_d)^2 \right] - \mathbb{E} \left[N_d \right]^2 \right)$$

$$\leq \frac{1}{n} + \Pr[\text{deg}(v) = d \wedge \text{deg}(u) = d]$$

$$- \Pr[\text{deg}(v) = d] \Pr[\text{deg}(u) = d] \quad \text{deg}(v) \stackrel{d}{=} \text{deg}(u)$$

$$\leq \frac{1}{n} + \Pr[X_1 + Y_1 = d \wedge X_1 + Y_2 = d \wedge X_2 + Y_2 \neq d]$$

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$$+ \Pr[X_1 + Y_1 = d \wedge X_1 + Y_2 = d \wedge X_2 + Y_2 \neq d | X_1 = 1] \Pr[X_1 = 1]$$

$$\leq \frac{1}{n} + \Pr[Y_1 = d \wedge Y_2 = d \wedge X_2 + Y_2 \neq d | X_1 = 0] \leq 1$$

$$+ \Pr[X_1 = 1]$$

$$\Rightarrow X_2 = 1$$

independent

$$= \frac{1}{n} + \Pr[Y_1 = d \wedge Y_2 = d \wedge X_2 = 1 | X_1 = 0] + \Pr[X_1 = 1] \leq \frac{1}{n} + \Pr[X_2 = 1] + \Pr[X_1 = 1]$$

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Fréchet: $\Pr[A] - \Pr[B] \leq \Pr[A \wedge \bar{B}]$

≤ 1

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$\left. \begin{array}{l} \blacksquare Y_1, Y_2 \sim \text{Bin}(n-2, p) \\ \blacksquare X_1, X_2 \sim \text{Ber}(p) \end{array} \right\} \text{independent}$

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$$\begin{aligned} \text{Var} \left[\frac{1}{n} N_d \right] &= \mathbb{E} \left[\left(\frac{1}{n} N_d \right)^2 \right] - \mathbb{E} \left[\frac{1}{n} N_d \right]^2 \leq \frac{1}{n} + 2p = \frac{1}{n} + 2 \frac{c}{n} \xrightarrow{n \rightarrow \infty} 0 \\ &= \frac{1}{n^2} \left(\mathbb{E} \left[(N_d)^2 \right] - \mathbb{E} \left[N_d \right]^2 \right) \\ &\leq \frac{1}{n} + \Pr[\text{deg}(v) = d \wedge \text{deg}(u) = d] \\ &\quad - \Pr[\text{deg}(v) = d] \Pr[\text{deg}(u) = d] \quad \text{deg}(v) \stackrel{d}{=} \text{deg}(u) \\ &\leq \frac{1}{n} + \Pr[X_1 + Y_1 = d \wedge X_1 + Y_2 = d \wedge X_2 + Y_2 \neq d] \\ &= \frac{1}{n} + \Pr[X_1 + Y_1 = d \wedge X_1 + Y_2 = d \wedge X_2 + Y_2 \neq d | X_1 = 0] \underbrace{\Pr[X_1 = 0]}_{\leq 1} \quad \text{Law of total probability} \\ &\quad + \Pr[X_1 + Y_1 = d \wedge X_1 + Y_2 = d \wedge X_2 + Y_2 \neq d | X_1 = 1] \Pr[X_1 = 1] \\ &\leq \frac{1}{n} + \Pr[Y_1 = d \wedge Y_2 = d \wedge X_2 + Y_2 \neq d | X_1 = 0] \leq 1 \quad \left. \begin{array}{l} \blacksquare Y_1, Y_2 \sim \text{Bin}(n-2, p) \\ \blacksquare X_1, X_2 \sim \text{Ber}(p) \end{array} \right\} \text{independent} \\ &\quad + \Pr[X_1 = 1] \quad \Rightarrow X_2 = 1 \\ &= \frac{1}{n} + \Pr[Y_1 = d \wedge Y_2 = d \wedge X_2 = 1 | X_1 = 0] + \Pr[X_1 = 1] \leq \frac{1}{n} + \Pr[X_2 = 1] + \Pr[X_1 = 1] \end{aligned}$$

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Application: ER – Degree Distribution

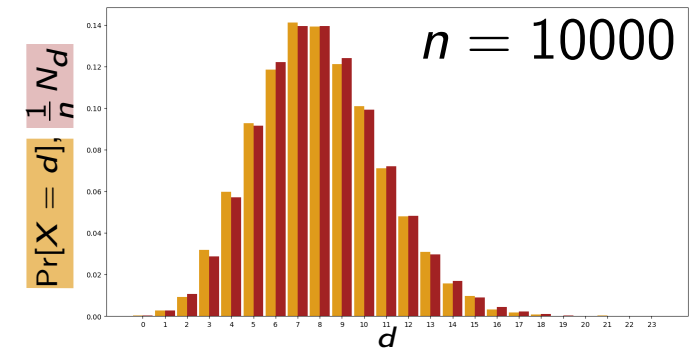
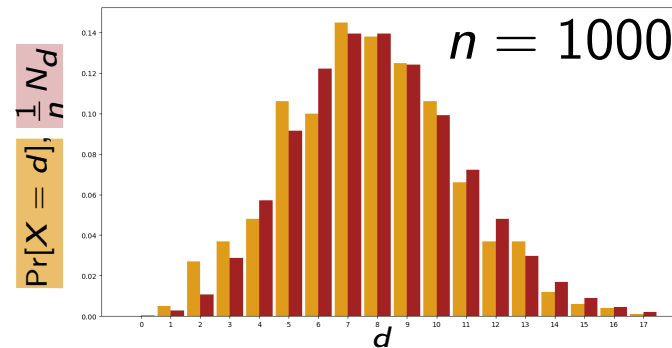
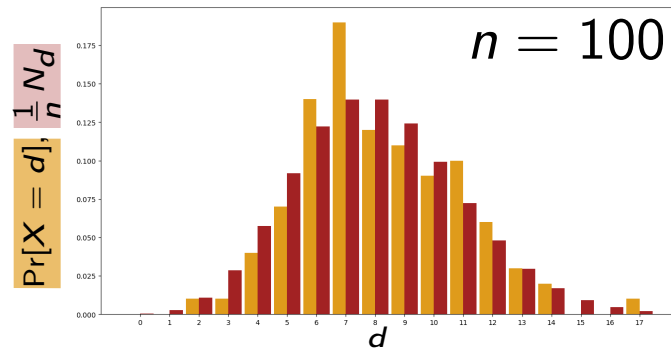
Theorem: Consider a $G(n, p)$ with $p = c/n$ for constant $c > 0$. For $\lambda = -n \log(1 - p)$, let $X \sim \text{Pois}(\lambda)$. Then for all $d > 0$ and every $\varepsilon > 0$ we have

$$\lim_{n \rightarrow \infty} \Pr \left[\left| \Pr[X = d] - \frac{1}{n} N_d \right| \geq \varepsilon \right] = 0.$$

$$\lambda = c + O(1/n) \rightarrow c \text{ for } n \rightarrow \infty$$

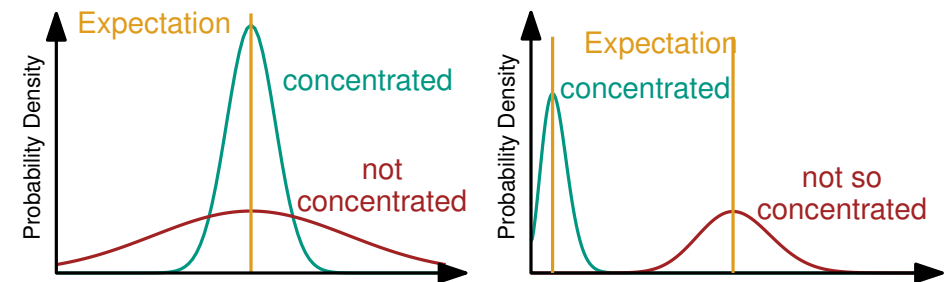
Proof

- Step 1: $\Pr[X = d]$ is close to the expectation of $\frac{1}{n} N_d$ $\lim_{n \rightarrow \infty} \left| \Pr[X = d] - \mathbb{E} \left[\frac{1}{n} N_d \right] \right| = 0 \checkmark$
- Step 2: $\frac{1}{n} N_d$ is concentrated (via Chebychev) $\lim_{n \rightarrow \infty} \Pr \left[\left| \mathbb{E} \left[\frac{1}{n} N_d \right] - \frac{1}{n} N_d \right| \geq \varepsilon \right] = 0 \checkmark$



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Definition: A **concentration inequality** bounds the probability of a random variable to deviate from a given value (typically its expectation) by a certain amount.



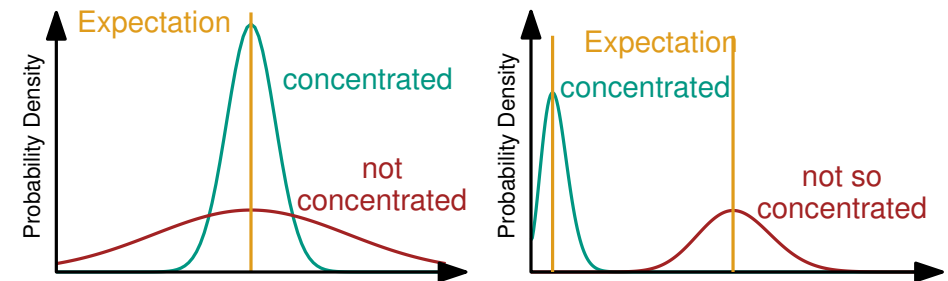
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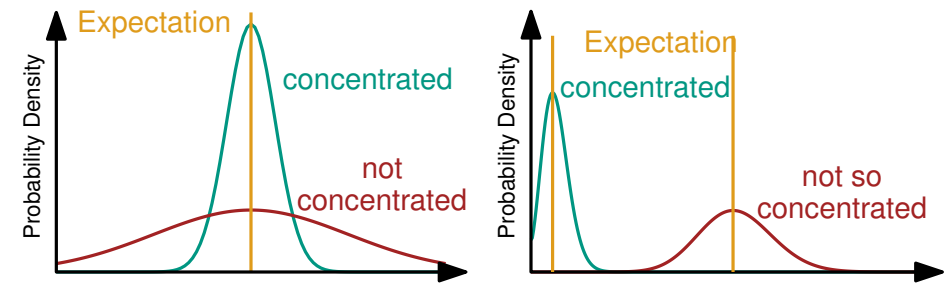
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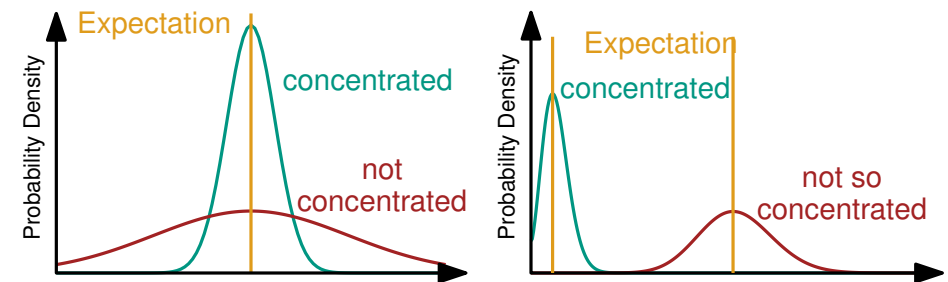
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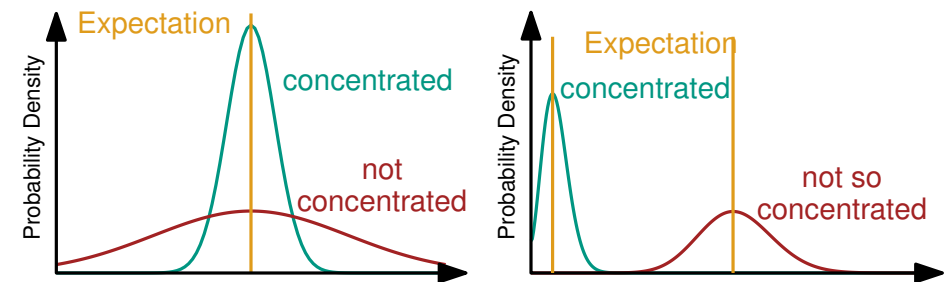
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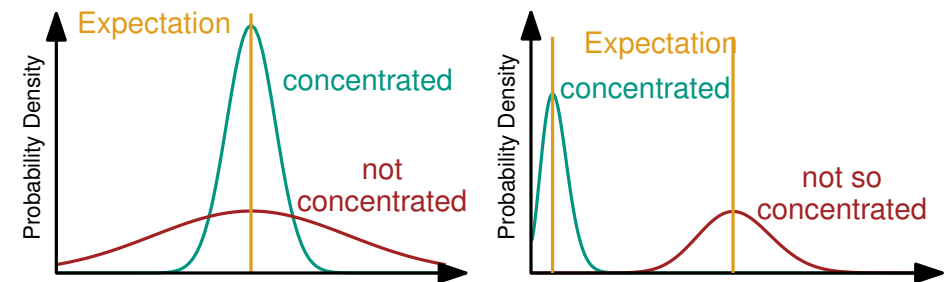
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Can we utilize higher-order moments for even stronger bounds?

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Theorem (Chernoff Bounds): Let X be a random variable and $a > 0$.
Then, $\Pr[X \geq a] \leq \min_{t>0} \mathbb{E}[e^{tX}] / e^{ta}$ and $\Pr[X \leq a] \leq \min_{t<0} \mathbb{E}[e^{tX}] / e^{ta}$.

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Another Moment Please

- The n -th raw moment of a random variable X is $\mathbb{E}[X^n]$
- We can capture *all* moments of X using a single function

Looks scary, but is again just $\mathbb{E}[f(X)]$ for $f(X) = e^{tX}$

Definition: For a random variable X the **moment generating function** is $M_X(t) = \mathbb{E}[e^{tX}]$

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(assuming the function exists in a neighborhood around 0)

Theorem: For independent random variables X, Y : $M_{X+Y}(t) = M_X(t) \cdot M_Y(t)$.

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Get bounds for specific random variables by finding a good t !

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Markov: ≤ 0.25

Chebychev: ≈ 0.022

Actual: ≈ 0.0000000138

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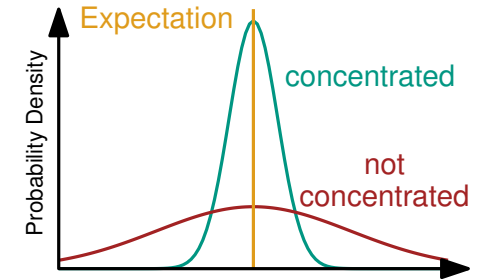
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- In fact, these also work when the X_i are Bernoulli random variables with different success probabilities

Conclusion

Concentration

- Is a random variable likely to yield values close to the expectation?
- Concentration inequalities bound the probability for a random variable to deviate from its expectation



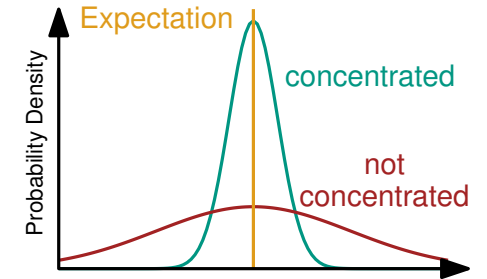
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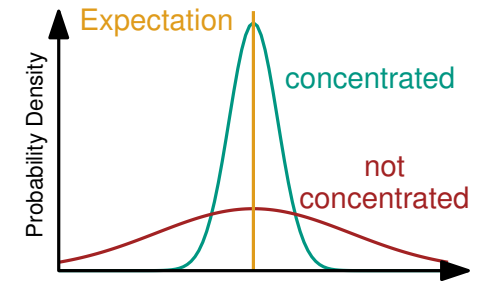
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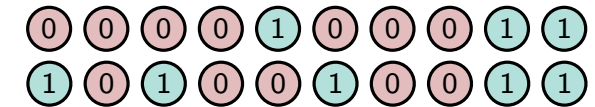
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Concentration Inequalities

- Markov: Based on first moment
- Chebychev: Squaring within Markov (utilizing second moment)
- Chernoff: Exponentiating within Markov (utilizing moment generating functions)
- Examples: Sum of coin flips, empirical degree distribution of ER graphs

