

Probability & Computing

Bounded Differences & Geometric Inhomogeneous Random Graphs





Concentration Inequalities

Bound the probability for a random variable to deviate from its expectation

Karlsruhe Institute of Technology

Recall: Concentration

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- Markov: generally applicable, but not very strong



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a ball does not fall into bin i



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$$\approx n \cdot e^{-k/n}$$
$$n \to \infty$$

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Institute of Theoretical Informatics, Algorithm Engineering & Scalable Algorithms



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2 Maximilian Katzmann, Stefan Walzer – Probability & Computing

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k balls distributed uniformly at random over *n* bins
Random variable X counts empty bins
Let X_i = 1_{Bin i is empty} for i ∈ [n] ⇒ X = ∑_{i=1}ⁿ X_i
Concentration: Pr[X ≥ E[X] + 5√k]
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Chebychev: tedious... ×



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 - Chernoff: X (our Bernoulli random variables are not independent)

 $X_1 = 0$ $X_2 = 1$ $X_3 = 0$ $X_4 = 1$ $X_5 = 0$ $X_6 = 0$

 $n \rightarrow \infty$

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 - $\Rightarrow X = f(Y_1, ..., Y_k) = \sum_{i \in [n]} \mathbb{1}_{\{ \nexists j: Y_j = i\}}$ (summands not independent, but the Y_j are)

$$\Pr[X \ge a] \le \mathbb{E}[X]/a.$$

Markov: X non-negative. a > 0:

 $V_2 = 3$

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 $= \sum_{i \in [n]} \max_{j \in [k]} \{2 - |\{Y_j, i\}|\}$ ("not" a sum Bernoulli random variables)

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 $V_{\cdot} = 1$

Can we show concentration for some arbitrary function of independent random variables? ... under certain conditions!

 $Y_2 - 3$



Aka ... Bounded differences inequality, McDiarmid's inequality, Azuma-Hoeffding inequality



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Definition: A function $f: S^n \to \mathbb{R}$ satisfies the **bounded differences condition** ("Lipschitz condition") with parameters Δ_i , if $|f(X_1, ..., X_i, ..., X_n) - f(X_1, ..., X'_i, ..., X_n)| \le \Delta_i$ for all $i \in [n]$ and $X_i, X'_i \in S$.





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Theorem: Let $X_1, ..., X_n$ be independent random variables taking values in a set *S*. Let $f: S^n \to \mathbb{R}$ satisfy the bounded differences condition with parameters Δ_i . Then, for $\Delta = \sum_{i \in [n]} \Delta_i^2$: $\Pr[|f - \mathbb{E}[f]| \ge t] \le 2e^{-2t^2/\Delta}$. (write *f* for $f(X_1, ..., X_n)$)





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Lemma: $\Pr[f \geq \mathbb{E}[f] + t] \leq e^{-2t^2/\Delta}$.



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Lemma: $\Pr[f \geq \mathbb{E}[f] + t] \leq e^{-2t^2/\Delta}$.

also for $\Pr[f \leq \mathbb{E}[f] - t]$



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Lemma: $\Pr[f \geq \mathbb{E}[f] + t] \leq e^{-2t^2/\Delta}$.

Cor.
$$\mathbb{E}[f] \leq g(n)$$
: $\Pr[f \geq cg(n)] \leq e^{-2((c-1)g(n))^2/\Delta}$.

also for $\Pr[f \leq \mathbb{E}[f] - t]$



- k balls distributed uniformly at random over n bins
- Random variable X counts empty bins
- $Y_1 = 1$ $Y_2 = 3$ $Y_6 = 4$ • Let independent $Y_j \sim \mathcal{U}([n])$ for $j \in [k]$ denote the bin of the j-th ball, and $X = f(Y_1, ..., Y_k)$



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Bounded differences condition

Intuition: How much can the number of empty bins change if we move a ball from one bin to another?

$$= 1 Y_2 = 3 Y_6 = 4$$

e *j*-th ball, and $X = f(Y_1, ..., Y_k)$
$$= |f(..., Y_i, ...) - f(..., Y'_i, ...)| \le \Delta_i$$

for all *i* and Y_i, Y'_i



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- Intuition: How much can the number of empty bins change if we move a ball from one bin to another?
 - A ball is moved from an almost empty bin to...



$$Y_2 = 3$$
 $Y_6 = 4$
j-th ball, and $X = f(Y_1, ..., Y_k)$
 $|f(..., Y_i, ...) - f(..., Y'_i, ...)| \le \Delta_i$
for all *i* and Y_i, Y'_i

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Application: Balls into Bins

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Bounded differences condition

- Intuition: How much can the number of empty bins change if we move a ball from one bin to another?
 - A ball is moved from an almost empty bin to...
 - ... an empty bin



 $|f(\ldots, Y_i, \ldots) - f(\ldots, Y'_i, \ldots)| \le \Delta_i$ for all *i* and Y_i, Y'_i

- k balls distributed uniformly at random over n bins
- Random variable X counts empty bins
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 - A ball is moved from an almost empty bin to...
 - ... an empty bin $\Rightarrow +1 1$



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 - A ball is moved from an almost empty bin to...
 - ... an empty bin $\Rightarrow +1 1 \Rightarrow \Delta_i = 0$



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$$Y_{1} = 1$$

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 $||f(..., Y_i, ...) - f(..., Y'_i, ...)| \leq \Delta_i$ for all *i* and Y_i, Y'_i

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 - A ball is moved from a not almost empty bin to...
 - ... an empty bin $\Rightarrow -1$

$$= 1 \qquad \stackrel{\uparrow}{Y_2} = 3 \qquad \stackrel{\uparrow}{Y_6} = 4$$

e *j*-th ball, and $X = f(Y_1, ..., Y_k)$
$$|f(..., Y_i, ...) - f(..., Y'_i, ...)| \le \Delta_i$$

for all *i* and Y: Y'

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ne *j*-th ball, and $X = f(Y_{1}, ..., Y_{k})$
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 - A ball is moved from a not almost empty bin to... $\Delta_i \leq 1$
 - ... an empty bin $\Rightarrow -1 \Rightarrow \Delta_i = 1$
 - ... a non-empty bin $\Rightarrow \Delta_i = 0$



$$|f(..., Y_i, ...) - f(..., Y'_i, ...)| \le \Delta_i$$

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Concentration via bounded differences

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$$\begin{array}{c} f_{1} = 1 \\ he \\ j-th \\ ball, and \\ X = f(Y_{1}, ..., Y_{k}) \\ \hline for all \\ i \\ and \\ Y_{i}, Y_{i}' \\ \end{array}$$

Function
$$f(Y_1, ..., Y_k)$$
:
 $Y_1, ..., Y_k$ independent
bounded differences Δ_i
 $\Delta = \sum_{i=1}^k \Delta_i^2$
Then $\Pr[f \ge \mathbb{E}[f] + t] \le e^{-2t^2/\Delta}$

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Concentration via bounded differences

 $\Delta = \sum_{i=1}^{k} \Delta_i^2 \leq \sum_{i=1}^{k} 1^2 = k$



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Concentration via bounded differences

 $\Delta = \sum_{i=1}^{k} \Delta_i^2 \le \sum_{i=1}^{k} 1^2 = k \quad \Rightarrow \Pr[f \ge \mathbb{E}[f] + 5\sqrt{k}] \le e^{-2(5\sqrt{k})^2/k}$



 V_{i}

$$|f(...,Y_{i},...) - f(...,Y'_{i},...)| \le \Delta_{i}$$

for all *i* and Y_{i},Y'_{i}

$$\leq 1 \begin{vmatrix} \mathsf{Function} \ f(Y_1, \dots, Y_k) :\\ \bullet \ Y_1, \dots, Y_k \text{ independent} \\ \bullet \text{ bounded differences } \Delta_i \\ \bullet \ \Delta = \sum_{i=1}^k \Delta_i^2 \\ \mathsf{Then} \ \mathsf{Pr}[f \geq \mathbb{E}[f] + t] \leq e^{-2t^2/\Delta} \end{vmatrix}$$

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 - A ball is moved from a not almost empty bin to... Δ_i
 - ... an empty bin $\Rightarrow -1 \Rightarrow \Delta_i = 1$
 - ... a non-empty bin $\Rightarrow \Delta_i = 0$

Concentration via bounded differences

 $\Delta = \sum_{i=1}^{k} \Delta_i^2 \le \sum_{i=1}^{k} 1^2 = k \implies \Pr[f \ge \mathbb{E}[f] + 5\sqrt{k}] \le e^{-2(5\sqrt{k})^2/k} = e^{-50}$



$$|f(..., Y_i, ...) - f(..., Y'_i, ...)| \le \Delta_i$$

for all *i* and Y_i, Y'_i

$$\leq 1 \quad \begin{array}{l} \text{Function } f(Y_1, ..., Y_k):\\ \bullet \ Y_1, ..., Y_k \text{ independent} \\ \bullet \text{ bounded differences } \Delta_i \\ \bullet \ \Delta = \sum_{i=1}^k \Delta_i^2 \\ \text{Then } \Pr[f \geq \mathbb{E}[f] + t] \leq e^{-2t^2/\Delta} \end{array}$$

Much better than Markov's
$$\rightarrow 1$$



Products are distributed uniformly at random over boxes on a conveyor belt





Products are distributed uniformly at random over boxes on a conveyor belt

$$m = n/k$$
 boxes $k = \log \log(n)$



Products are distributed uniformly at random over boxes on a conveyor belt

n products m = n/k boxes $k = \log \log(n)$







k+1



n products m = n/k boxes $k = \log \log(n)$

• A camera scans k+1 consecutive boxes simultaneously

• Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned

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Application: The Factory



• Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned

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k+1

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k+1

Application: The Factory



n products m = n/k boxes $k = \log \log(n)$

- A camera scans k+1 consecutive boxes simultaneously
- Problem: Empty box in view ⇒ reflection blinds camera, products remain unscanned
 Question: How many products avoid quality assurance?

k + 1

Application: The Factory



n products m = n/k boxes $k = \log \log(n)$ • A camera scans k+1 consecutive boxes simultaneously • Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned

• Question: How many products avoid quality assurance? Show: o(n) with prob. $1 - O(\frac{1}{n})$
k + 1

Application: The Factory



- A camera scans k+1 consecutive boxes simultaneously
- Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned
- Question: How many products avoid quality assurance? Show: o(n) with prob. $1-O(\frac{1}{n})$ Formalize
- chain: consecutive sequence of non-empty boxes

 $\dot{k+1}$

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

- A camera scans k+1 consecutive boxes simultaneously
- Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned
- Question: How many products avoid quality assurance? Show: o(n) with prob. $1 O(\frac{1}{n})$ Formalize
- chain: consecutive sequence of non-empty boxes
- short chain: incl. max. chain of length $\leq k \Rightarrow$ exactly products in short chains unscanned

 $\dot{k+1}$

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

- A camera scans k+1 consecutive boxes simultaneously
- Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned
- Question: How many products avoid quality assurance? Show: o(n) with prob. $1 O(\frac{1}{n})$ Formalize
- chain: consecutive sequence of non-empty boxes
- *short chain*: incl. max. chain of length $\leq k \Rightarrow$ *exactly* products in short chains unscanned
- X_i = number of products in box *i*, Y_i = indicator whether box *i* is in a short chain

k + 1

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

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• Then $X = \sum_{i=1}^{m} X_i \cdot Y_i$ is the number of unscanned products

k + 1

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

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- Problem: Dependencies (between X_i 's, between X_i and Y_i)

k + 1

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

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- chain: consecutive sequence of non-empty boxes
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- X_i = number of products in box *i*, Y_i = indicator whether box *i* is in a short chain
- Then $X = \sum_{i=1}^{m} X_i \cdot Y_i$ is the number of unscanned products
- Problem: Dependencies (between X_i 's, between X_i and Y_i)
- Solution: Relax dependencies and compute upper bound instead

 $\dot{k+1}$

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

n products m = n/k boxes $k = \log \log(n)$

- A camera scans k+1 consecutive boxes simultaneously
- Problem: Empty box in view \Rightarrow reflection blinds camera, products remain unscanned
- Question: How many products avoid quality assurance? Show: o(n) with prob. $1 O(\frac{1}{n})$ **Relax and bound**
- X_i = number of products in box *i*, Y_i = indicator whether box *i* is in a short chain
- Then $X = \sum_{i=1}^{m} X_i \cdot Y_i$ is the number of unscanned products

k + 1

Application: The Factory

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- A camera scans k+1 consecutive boxes simultaneously
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- Question: How many products avoid quality assurance? Show: o(n) with prob. $1 O(\frac{1}{n})$ **Relax and bound**
- X_i = number of products in box *i*, Y_i = indicator whether box *i* is in a short chain
- Then $X = \sum_{i=1}^{m} X_i \cdot Y_i$ is the number of unscanned products
- $E_k(i)$ = number of empty boxes in box *i* and *k* closest (assuming *k* even)

 $k \stackrel{.}{+} 1$

Application: The Factory

Products are distributed uniformly at random over boxes on a conveyor belt

n products m = n/k boxes $k = \log \log(n)$

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 $E_{k}(i) = 1$



 $\dot{k+1}$

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► Box *i* in short chain $\Rightarrow E_k(i) > 0$

• Y'_i = indicator whether $E_k(i) > 0 \Rightarrow Y_i \le Y'_i$

 $\dot{k+1}$

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- $E_k(i) =$ number of empty boxes in box *i* and *k* closest (assuming *k* even)
 - Box *i* in short chain $\Rightarrow E_k(i) > 0$

•
$$Y'_i$$
 = indicator whether $E_k(i) > 0 \Rightarrow Y_i \le Y'_i$

 $L = \sum_{i=1}^{m} X_i \cdot Y_i \leq \sum_{i=1}^{m} X_i \cdot Y'_i =: X'$







$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$















n products *m*=*n*/*k* boxes, *k*=log log(*n*) *X'* = $\sum X_i \cdot Y'_i$ *X_i*, products in box *i E_k(i)*, number empty boxes in box *i* and *k* closest *Y'_i*, indicator *E_k(i)* > 0

1. 1

















knowing that exactly ℓ boxes are empty









• Else: *n* products distributed u.a.r. over $m' = m - \ell$ boxes









$$m' \ge \frac{n}{\log \log(n)} - \log \log(n)$$















 $\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$ (law of total expectation) $\stackrel{}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$ $= \sum_{\ell=0}^{k} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$ $= \sum_{\ell=1}^{k} \mathbb{E}[X_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$ $\leq \sum_{\ell=1}^{k} 2\log\log(n) \cdot \Pr[E_k(i) = \ell]$ $= 2\log\log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$



$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
(law of total expectation) $\stackrel{k=1}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$

$$= \sum_{\ell=0}^{k} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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n products *m*=n/k boxes, k=log log(n)
X' = ∑ X_i ⋅ Y'_i
X_i, products in box *i*E_k(*i*), number empty boxes in box *i* and k closest

•
$$Y'_i$$
, indicator $E_k(i) > 0$

 \leq Pr["Exists an empty box among k + 1"]



$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y_i']$$
(law of total expectation)
$$\stackrel{k}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y_i' \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$= 2 \log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$

$$\leq \Pr[\text{``Exists an empty box among } k + 1'']$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{``A given box is empty''}]$$

i and k closest



 $\mathbb{E}[X'] =$

$$E[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
(law of total expectation)
$$E[X_i \cdot Y'_i | E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

$$= \sum_{\ell=0}^{k} \mathbb{E}[X_i \cdot Y'_i | E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

$$= \sum_{\ell=1}^{k} \mathbb{E}[X_i | E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$\leq \Pr[\text{"Exists an empty box among } k + 1"]$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{"A given box is empty"}]$$

$$\leq 2k \left(1 - \frac{1}{m}\right)^{n}$$



$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
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$$= \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$\leq \sum_{\ell=1}^{k} 2 \log \log(n) \cdot \Pr[E_k(i) = \ell]$$

$$= 2 \log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$

$$\leq \Pr[\text{``Exists an empty box among } k + 1'']$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{``A given box is empty''}]$$

$$\leq 2k \left(1 - \frac{1}{m}\right)^n$$
a product hits a given box

 $k \mid 1$



$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
(law of total expectation) $\stackrel{k}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$

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$$= \sum_{\ell=1}^{k} \mathbb{E}[X_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

$$= \sum_{\ell=1}^{k} \mathbb{E}[X_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$= 2\log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$

$$\leq \Pr[\text{"Exists an empty box among } k + 1"]$$
(union bound) $\leq (k+1) \cdot \Pr[\text{"A given box is empty"}]$

$$\leq 2k \left(1 - \frac{1}{m}\right)^n$$
a product does not hit a given box



in box *i*

 $E_{k}(i) > 0$

$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y_i']$$
(law of total expectation) $\stackrel{}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y_i' \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$

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$$\leq \sum_{\ell=1}^{k} 2 \log \log(n) \cdot \Pr[E_k(i) = \ell]$$

$$= 2 \log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$

$$\leq \Pr[\text{"Exists an empty box among } k + 1"]$$
(union bound) $\leq (k+1) \cdot \Pr[\text{"A given box is empty"}]$

$$\leq 2k \left(1 - \frac{1}{m}\right)^{n}$$
none of the *n* products hit a given box


$$\mathbb{E}[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
(law of total expectation)
$$\stackrel{\bullet}{=} \sum_{\ell=0}^{k+1} \mathbb{E}[X_i \cdot Y'_i \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$\leq \Pr[\text{"Exists an empty box among } k + 1"]$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{"A given box is empty"}]$$

$$\leq 2k \left(1 - \frac{1}{m}\right)^n$$

$$= 2k \left(1 - \frac{k}{n}\right)^n$$

 $k \mid 1$



 $\mathbb{E}[X'] =$

$$E[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y_i']$$
(law of total expectation)
$$\sum_{k=0}^{k+1} \mathbb{E}[X_i \cdot Y_i' \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$\leq \sum_{\ell=1}^{k} 2 \log \log(n) \cdot \Pr[E_k(i) = \ell]$$

$$= 2 \log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{``A given box is empty'']}$$

$$\leq 2k \left(1 - \frac{1}{n}\right)^n \sum_{\ell=1}^{n} (1 + x \le e^x)$$

 $k \mid 1$



 $\mathbb{E}[X'] =$

$$E[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y_i']$$
(law of total expectation)
$$\sum_{k=0}^{k+1} \mathbb{E}[X_i + Y_i' \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

$$= \sum_{\ell=0}^{k} \mathbb{E}[X_i \cdot Y_i' \mid E_k(i) = \ell] \cdot \Pr[E_k(i) = \ell]$$

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$$\leq \sum_{\ell=1}^{k} 2 \log \log(n) \cdot \Pr[E_k(i) = \ell]$$

$$= 2 \log \log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$
(union bound)
$$\leq (k+1) \cdot \Pr[\text{``A given box is empty'']}$$

$$\leq 2k \left(1 - \frac{1}{n}\right)^n \leq 2k \cdot e^{-k} = 2 \frac{\log \log(n)}{\log(n)}$$

 $k \mid 1$



 $\mathbb E$

$$[X'] = \sum_{i=1}^{m} \mathbb{E}[X_i \cdot Y'_i]$$
Haw of total expectation)
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$$= 2\log\log(n) \sum_{\ell=1}^{k} \Pr[E_k(i) = \ell]$$

$$\leq 2\log\log(n) \cdot 2\frac{\log\log(n)}{\log(n)}$$

$$= 4\frac{\log\log(n)^2}{\log(n)}$$

n products

 $X' = \sum X_i \cdot Y'_i$

• X_i , products in box *i*

box *i* and *k* closest Y'_i , indicator $E_k(i) > 0$

• m = n/k boxes, $k = \log \log(n)$

• $E_k(i)$, number empty boxes in







 $\blacksquare E_k(i)$, number empty boxes in

box *i* and *k* closest





















n products *m*=*n*/*k* boxes, *k*=log log(*n*) *X* = $\sum X_i \cdot Y_i$ *X_i*, products in box *i Y_i*, indicator *i* in short chain $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$

Institute of Theoretical Informatics, Algorithm Engineering & Scalable Algorithms





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• View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var. where Z_j for $j \in [n]$ denotes the box of the *j*-th product







• View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var. where Z_j for $j \in [n]$ denotes the box of the *j*-th product

Bounded differences condition:







Bounded Differences

• View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var.

where Z_j for $j \in [n]$ denotes the box of the *j*-th product

- Bounded differences condition:
 - Worst change in number of products in short chains when moving a single product from one box to another

n products *m* = *n*/*k* boxes, *k* = log log(*n*) *X* = ∑ *X_i* · *Y_i X_i*, products in box *i Y_i*, indicator *i* in short chain
E[X] ≤ E[X'] ≤ 4n \frac{\log \log(n)}{\log(n)} *f*(..., *Z_j*, ...) - *f*(..., *Z'_i*, ...)| ≤ Δ_j

for all *j* and Z_i, Z'_i



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 - Consider chain of 2k + 1 boxes containing *all n* products and one box contains only one of them



n products *m*=*n*/*k* boxes, *k*=log log(*n*) *X* = $\sum X_i \cdot Y_i$ *X_i*, products in box *i Y_i*, indicator *i* in short chain $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$ *f*(..., *Z_i*, ...) - *f*(..., *Z'_i*, ...)| < Δ_i

$$|f(..., Z_j, ...) - f(..., Z_j, ...)| \leq \Delta_j$$

for all *j* and Z_j, Z'_j



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Image of the i

$$|f(..., \mathbb{Z}_j, ...) - f(..., \mathbb{Z}_j, ...)| \leq \Delta_j$$

for all *j* and Z_j, Z'_j



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n products $m = n/k \text{ boxes, } k = \log \log(n)$ $X = \sum X_i \cdot Y_i$ $X_i, \text{ products in box } i$ $Y_i, \text{ indicator } i \text{ in short chain}$ $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$

for all j and
$$Z_j, Z'_j$$
, Z'_j, Z'_j



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 $\Rightarrow X = 0$, since no short chain and, thus, no products in short chains

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$$|f(..., Z_j, ...) - f(..., Z'_j, ...)| \le \Delta_j$$

for all j and Z_j, Z'_j



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E[X] ≤ E[X'] ≤ 4n \frac{\log \log(n)}{\log(n)}

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for all j and Z_j, Z'_j

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Move product to next box



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 $\Rightarrow X = 0$, since no short chain and, thus, no products in short chains

Move product to next box



Bounded Differences

• View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var.

where Z_j for $j \in [n]$ denotes the box of the *j*-th product

- Bounded differences condition:
 - Worst change in number of products in short chains when moving a single product from one box to another
 - Consider chain of 2k + 1 boxes containing all n products and one box contains only one of them



 $\Rightarrow X = 0$, since no short chain and, thus, no products in short chains

- Move product to next box
 - $\Rightarrow X = n$, since all products in short chains now

n products
m=*n*/*k* boxes, *k*=log log(*n*)
X =
$$\sum X_i \cdot Y_i$$
X_i, products in box *i Y_i*, indicator *i* in short chain
 $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$
f(..., *Z_i*, ...) - *f*(..., *Z'_i*, ...)| < Δ_i

$$|f(..., Z_j, ...) - f(..., Z'_j, ...)| \le \Delta_j$$

for all j and Z_j, Z'_j



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 $\Delta_j \leq n$



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n products • m = n/k boxes, $k = \log \log(n)$ $X = \sum X_i \cdot Y_i$ \mathbf{X}_i , products in box *i* • Y_i , indicator *i* in short chain $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$ $|f(\ldots, Z_j, \ldots) - f(\ldots, Z'_j, \ldots)| \leq \Delta_j$ for all *j* and Z_i, Z'_i Function $f(Z_1, ..., Z_n)$: \blacksquare Z₁, ..., Z_n independent • bounded differences Δ_i



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7



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n products

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$$\frac{1}{n^3}$$
This bound is useless, since worst-case changes are too big
$$= \exp\left(-\Theta\left(\frac{\log \log(n)^2}{n \log(n)^2}\right)\right) \xrightarrow{n \to \infty} 1$$

n products

 $X = \sum X_i \cdot Y_i$

for all j and Z_i, Z'_i

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 $\Pr[f > cg(n)] \le e^{-2((c-1)g(n))^2/\Delta}$


Concentration of X (for *n* large enough)

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$$Function f(Z_1, \dots, Z_n):$$

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$$= C_1 + C_2 +$$

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Method of Typical Bounded Differences



Definition: A function $f: S^n \to \mathbb{R}$ satisfies the **typical bounded differences condition** with respect to • an event $A \subset S^n$ and

■ parameters
$$\Delta_i^A \leq \Delta_i$$
 for $i \in [n]$,
if $|f(X_1, ..., X_i, ..., X_n) - f(X_1, ..., X'_i, ..., X_n)| \leq \begin{cases} \Delta_i^A, \text{ if } (X_1, ..., X_i, ..., X_n) \in A, \\ \Delta_i, \text{ otherwise} \end{cases}$
for all $i \in [n]$ and $X_i, X'_i \in S$.

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• Δ_i^A is worst-case change, assuming A held before the change



Definition: A function $f: S^n \to \mathbb{R}$ satisfies the **typical bounded differences condition** with respect to

- an event $A \subset S^n$ and
- parameters $\Delta_i^A \leq \Delta_i$ for $i \in [n]$,
- $|f(X_1, ..., X_i, ..., X_n) f(X_1, ..., X'_i, ..., X_n)| \le \begin{cases} \Delta_i^A, \text{ if } (X_1, ..., X_i, ..., X_n) \in A, \\ \Delta_i, \text{ otherwise} \end{cases}$ if for all $i \in [n]$ and $X_i, X'_i \in S$.

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Theorem: Let $X_1, ..., X_n$ be independent random variables taking values in a set S, let $A \subseteq S^n$ be an event, and let $f: S^n \to \mathbb{R}$ satisfy the typical bounded differences condition w.r.t. A and parameters $\Delta_i^A \leq \Delta_i$. Then, for $g(n) \geq \mathbb{E}[f]$, for all $\varepsilon_i \in (0, 1]$ and $\Delta = \sum_{i \in [n]} (\Delta_i^A + \varepsilon_i (\Delta_i - \Delta_i^A))^2 \cdot \Pr[f \ge cg(n)] \le e^{-((c-1)g(n))^2/(2\Delta)} + \Pr[\neg A] \sum_{i \in [n]} \frac{1}{\varepsilon_i}$

Corollary of "On the Method of Typical Bounded Differences", Warnke, Comb. Probab. Comput. 2015





Function of independent random variables as before



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A is the good, typical event that should be very likely to occur



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• A is the good, typical event that should be very likely to occur

 $\square \Delta$ is sum of squared worst-case changes as before



- Function of independent random variables as before
- A is the good, typical event that should be very likely to occur
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- But we have to pay for the mitigation!
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The more we need to mitigate, the higher the price! Not too bad if *A* is very likely to occur!



- View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var. where Z_j for $j \in [n]$ denotes the box of the *j*-th product
- Bounded differences condition: $\Delta_j \leq n$





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 $\mathbb{E}[B_i] = \frac{n}{m} = \frac{n}{\frac{n}{\log\log(n)}} = \log\log(n)$

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$$\mathbb{E}[B_i] = \frac{n}{m} = \frac{n}{\frac{n}{\log\log(n)}} = \log\log(n)$$

• And, thus, expected number in sequence of 2k + 1 boxes

$$\mathbb{E}[S] = \sum_{i=1}^{2k+1} \mathbb{E}[B_i] = O(\log \log(n)^2)$$

Image: A state of the state



- t Note: n products $m = n/k \text{ boxes, } k = \log \log(n)$ $X = \sum X_i \cdot Y_i$ $X_i, \text{ products in box } i$ $Y_i, \text{ indicator } i \text{ in short chain}$ $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$
- View X as a function $f(Z_1, ..., Z_n)$ of independent rand. var. where Z_j for $j \in [n]$ denotes the box of the *j*-th product
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 - When all *n* products fall into $2k + 1 = O(\log \log(n))$ boxes
 - But expected number of products in a single box i:

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• And, thus, expected number in sequence of 2k + 1 boxes

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 - Union bound over $\leq n$ sequences: $\Pr[\neg A] \leq n^{-\delta \varepsilon^2/3+1} \leq n^{-\lambda}$ (for arbitrarily large λ)





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Moving one product empties at most one box







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• Moving one product empties at most one box \Rightarrow at most two new short chains



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contain $O(\log(n))$ products

- Moving one product empties at most one box \Rightarrow at most two new short chains
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Function $f(Z_1, ..., Z_n)$:
• *Z*₁, ..., *Z_n* independent
• typical event *A*
• bounded differences $\Delta_j^A \leq \Delta_j$
• $\Delta = \sum_{j=1}^n (\Delta_j^A + \varepsilon_j (\Delta_j - \Delta_j^A))^2$
• $g(n) \geq \mathbb{E}[f]$
 $\Pr[f \geq cg(n)] \leq e^{-((c-1)g(n))^2/(2\Delta)} + \Pr[\neg A] \sum_{j=1}^n \frac{1}{\varepsilon_j}$



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$$= \sum_{j=1}^{n} (O(\log(n)) + 1)^{2}$$

$$= O(n \log(n)^{2}) \qquad \text{Much better than } n^{3} \text{ from before!}$$

n products • m = n/k boxes, $k = \log \log(n)$ $X = \sum X_i \cdot Y_i$ \bullet X_i, products in box i • Y_i , indicator *i* in short chain $\mathbb{E}[X] \leq \mathbb{E}[X'] \leq 4n \frac{\log \log(n)}{\log(n)}$ Function $f(Z_1, ..., Z_n)$: \square $Z_1, ..., Z_n$ independent typical event A • bounded differences $\Delta_i^A \leq \Delta_j$ $\Delta = \sum_{j=1}^{n} (\Delta_{j}^{A} + \varepsilon_{j} (\Delta_{j} - \Delta_{j}^{A}))^{2}$ $g(n) \geq \mathbb{E}[f]$ $\Pr[f \ge cg(n)] \le e^{-((c-1)g(n))^2/(2\Delta)}$ $+ \Pr[\neg A] \sum_{i=1}^{n} \frac{1}{\varepsilon_i}$

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Average-case analysis: analyze models that represent the real world

- Models seen so far
 - Erdős-Rényi random graphs: simple but no locality



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konect.cc/plot/degree.a.youtube-links.full.png



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Idea

Add Pareto distribution to RGGs







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$$x \leq \lambda \frac{w_v \cdot y}{n} \Leftrightarrow \mathbf{y} \geq \frac{n}{\lambda w_v} \mathbf{x}$$





Definition

Consider n vertices

- For each vertex *v* independently:
 - Draw a *position* x_v uniformly on \mathbb{T}^d
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"Power-Law Exponent"

$$\underbrace{\operatorname{dist}(x_u, x_v)}_{L_{\infty}-\operatorname{norm}} \leq \left(\lambda \frac{w_u \cdot w_v}{n} \right)^{1/d}$$

const. controls the avg. degree
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 W_{V}

0

Connect *u* and *v* with an edge, iff

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Institute of Theoretical Informatics, Algorithm Engineering & Scalable Algorithms



Definition

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L_∞-norm const. controls the avg. degree For d = 1, linear relation between distance and weight $y = w_{\mu}, x = \text{dist}(x_{\mu}, x_{\nu})$

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The lower w_v , the steeper the wedge \downarrow The lower the degree







Consider n vertices

- For each vertex *v* independently:
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"Power-Law Exponent"

Connect *u* and *v* with an edge, iff

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• Consider vertex v with weight w_v



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Expected Degree (d = 1)

Consider vertex v with weight wv
We want to compute E[deg(v) | wv





Consider vertex v with weight wv
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 This is a random variable



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- Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$



$$\begin{array}{l} \textbf{GIRG} \\ \bullet n \text{ independent vertices} \\ \bullet x_v \sim \mathcal{U}([0,1]) \\ \bullet w_v \sim \operatorname{Par}(\tau-1,1) \text{ for } \tau \in (2,3) \\ f_{w_v}(w) = (\tau-1)w^{-\tau} \\ \bullet u, v \text{ adjacent iff} \\ \operatorname{dist}(x_u, x_v) \leq \lambda \frac{w_u \cdot w_v}{n} \end{array}$$

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GIRG
n independent vertices

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$$= \Theta(n \Pr[\{u, v\} \in E \mid w_v])$$
$$u \in N(v)$$





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N(v)











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Expected Degree (d = 1)

• Consider vertex v with weight w_v

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$$= \Theta(n \int_1^\infty \Pr[u \in N(v) \mid w_u = w, w_v] f_{w_u}(w) dw)$$

$$= \Theta\left(n \left(\int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw + \int_{\frac{n}{2\lambda w_v}}^\infty 1 \cdot f_{w_u}(w) dw\right)\right)$$

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$$= \Pr[w_u \ge \frac{n}{2\lambda w_v}]$$

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$$\begin{aligned} \deg(v) &= \sum_{u \in V \setminus \{v\}} X_u \\ \mathbb{E}[\deg(v) \mid w_v] &= \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v] \\ &= \Theta(n \operatorname{Pr}[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \\ &= \Theta(n \int_1^\infty \operatorname{Pr}[u \in N(v) \mid w_u = w, w_v] f_{w_u}(w) dw] \\ &= \Theta\left(n \left(\int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw + \int_{\frac{n}{2\lambda w_v}}^{\infty} 1 \cdot f_{w_u}(w) dw\right)\right) \\ &= \operatorname{Pr}[w_u \ge \frac{n}{2\lambda w_v}] \end{aligned}$$

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$$\begin{aligned} \deg(v) &= \sum_{u \in V \setminus \{v\}} X_u \\ \mathbb{E}[\deg(v) \mid w_v] &= \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v] \\ &= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \\ &= \Theta(n \int_1^\infty \Pr[u \in N(v) \mid w_u = w, w_v] f_{w_u}(w) dw) \\ &= \Theta\left(n \left(\int_1^{\frac{2}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_v}(w) dw + \int_{\frac{n}{2\lambda w_v}}^\infty 1 \cdot f_{w_u}(w) dw\right)\right) \\ &= \Pr[w_u \ge \frac{n}{2\lambda w_v}] \\ &= \Pr[w_u \ge \frac{n}{2\lambda w_v}] \\ &= \Pr[w_u \ge 1] = 1 \end{aligned}$$

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Consider vertex v with weight
$$w_v$$
We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \left\{ \Theta(n), \text{ if } w_v \geq \frac{n}{2\lambda} \right\}$
Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$
 $\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$
 $\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$
 $= \Theta(n \Pr[\{u, v\} \in E \mid w_v])$
W.l.o.g $x_v = \frac{1}{2}$
If $w_v < \frac{n}{2\lambda}$
 $= \Theta(n \int_1^{\infty} \Pr[u \in N(v) \mid w_u = w, w_v] f_{w_u}(w) dw)$
 $= \Theta\left(n \left(\int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw + \Pr[w_u \geq \frac{n}{2\lambda w_v}]\right)\right)$

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Consider vertex v with weight w_v We want to compute $\mathbb{E}[\deg(v) | w_v] = \begin{cases} \Theta(n), \text{ if } w_v \ge \frac{n}{2\lambda} \\ \end{bmatrix}$ • Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$ $\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$ $\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$ $= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$ If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) | w_u = w, w_v] f_{w_u}(w) dw$ $= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right)$ (via CDF of Par) = $\left(\frac{n}{2\lambda w_{v}}\right)^{-(\tau-1)}$

GIRG
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$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$
If $\frac{w_v < \frac{n}{2\lambda}}{w_v < \frac{n}{2\lambda}} = \Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw$
$= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right)$
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< 1

GIRG
n independent vertices

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(via CDF of Par) = $\left(\frac{n}{2\lambda w_v}\right)^{-(\tau-1)}$
$=\left(\underbrace{\frac{2\lambda w_{v}}{n}}_{<1}\right)^{\tau-1}$
$=O(\frac{w_v}{n})$

GIRG
• *n* independent vertices
•
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$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$
$= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right)$
$= \Theta\left(n \int_{c}^{\frac{n}{2\lambda_{W_{v}}}} \frac{w \cdot w_{v}}{w \cdot w_{v}} f_{w}\left(w\right) dw\right) + O(w_{v})$

$$\begin{array}{l} \textbf{GIRG} \\ \bullet n \text{ independent vertices} \\ \bullet x_v \sim \mathcal{U}([0,1]) \\ \bullet w_v \sim \operatorname{Par}(\tau-1,1) \text{ for } \tau \in (2,3) \\ f_{w_v}(w) = (\tau-1)w^{-\tau} \\ \bullet u, v \text{ adjacent iff} \\ \operatorname{dist}(x_u, x_v) \leq \lambda \frac{w_u \cdot w_v}{n} \end{array}$$



• Consider vertex v with weight $w_v \qquad \int \Theta(n)$, if $w_v \geq \frac{n}{2\lambda}$
• We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \{$
• Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$
$\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$
$\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$
$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$
$= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right)$
$= \Theta\left(n\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) \mathrm{d}w\right) + O(w_{v})$
$= \Theta\left(n\frac{w_{v}}{n}\int_{1}^{\frac{n}{2\lambda w_{v}}} w \cdot (\tau - 1)w^{-\tau} dw\right) + O(w_{v})$

$$\begin{array}{l} \textbf{GIRG} \\ \bullet \ n \ \text{independent vertices} \\ \bullet \ x_v \sim \mathcal{U}([0,1]) \\ \bullet \ w_v \sim \operatorname{Par}(\tau-1,1) \ \text{for} \ \tau \in (2,3) \\ f_{w_v}(w) = (\tau-1)w^{-\tau} \\ \bullet \ u, \ v \ \text{adjacent iff} \\ \operatorname{dist}(x_u, x_v) \leq \lambda \frac{w_u \cdot w_v}{n} \end{array}$$



• Consider vertex v with weight $w_v \qquad \int \Theta(n)$, if $w_v \geq \frac{n}{2\lambda}$	(
• We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \{$	
Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$	
$\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$	
$\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$	
$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \qquad \text{w.l.o.g } x_v = \frac{1}{2}$	
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$	
$= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right)$	
$= \Theta\left(n\int_{1}^{\frac{n}{2\lambda_{w_{v}}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) \mathrm{d}w\right) + O(w_{v})$	
$= \Theta\left(n \frac{w_v}{n} \int_1^{\frac{n}{2\lambda w_v}} w \cdot (\tau - 1) w^{-\tau} d w\right) + O(w_v)$	
$=\Theta\left(w_{v}\int_{1}^{\frac{n}{2\lambda w_{v}}}w^{-(\tau-1)}dw\right)+O(w_{v})$	

GIRG
n independent vertices

$$x_v \sim \mathcal{U}([0, 1])$$

 $w_v \sim \operatorname{Par}(\tau - 1, 1)$ for $\tau \in (2, 3)$
 $f_{w_v}(w) = (\tau - 1)w^{-\tau}$
u, *v* adjacent iff
 $\operatorname{dist}(x_u, x_v) \leq \lambda \frac{w_u \cdot w_v}{n}$

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• Consider vertex v with weight
$$w_v$$

• We want to compute $\mathbb{E}[\deg(v) | w_v] = \begin{cases} \Theta(n), \text{ if } w_v \ge \frac{n}{2\lambda} \\ 0 \le 2\lambda \end{cases}$
• Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$
 $\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$
 $\mathbb{E}[\deg(v) | w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u | w_v]$
 $= \Theta(n \Pr[\{u, v\} \in E | w_v])$ w.i.o.g $x_v = \frac{1}{2}$
If $w_v < \frac{n}{2\lambda}$ $= \Theta(n \int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw + \Pr[w_u \ge \frac{n}{2\lambda w_v}])$
 $= \Theta\left(n \left(\int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw + \Pr[w_u \ge \frac{n}{2\lambda w_v}]\right)\right)$ $\Rightarrow \Theta\left(n \int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw\right) + O(w_v)$
 $= \Theta\left(n \int_1^{\frac{n}{2\lambda w_v}} \frac{w \cdot w_v}{n} f_{w_u}(w) dw\right) + O(w_v)$
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 $= \Theta\left(w_v \int_1^{\frac{n}{2\lambda w_v}} w \cdot (\tau - 1) w^{-\tau} dw\right) + O(w_v)$

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• Consider vertex v with weight $w_v \qquad \int \Theta(n)$, if $w_v \geq \frac{n}{2\lambda}$	GIRG
• We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \{$	n independent vertices
• Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$	
$\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$	$f_{w_v}(w) = (\tau - 1)w^{-\tau}$
$\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$	• u, v adjacent iff dist $(x_u, x_v) < \lambda \frac{w_u \cdot w_v}{w_u \cdot w_v}$
$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$	(u, v) = n
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$	n
$=\Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}}\frac{w \cdot w_{v}}{n}f_{w_{u}}(w)dw + \Pr[w_{u} \geq \frac{n}{2\lambda w_{v}}]\right)\right) = \Theta\left(w_{u} \otimes \frac{w}{2\lambda w_{v}}\right)$	$V_{v}\left[\frac{1}{-(\tau-2)}W^{-(\tau-2)}\right]_{1}^{\frac{1}{2\lambda w_{v}}}+O(w_{v})$
$=\Theta\left(n\int_{1}^{\frac{n}{2\lambda w_{v}}}\frac{w \cdot w_{v}}{n}f_{w_{u}}(w)dw\right)+O(w_{v}) \qquad \qquad =\Theta\left(w_{v}\right)$	$\left[w^{-(\tau-2)}\right]_{\frac{n}{2\lambda w_{v}}}^{1} + O(w_{v})$
$= \Theta\left(n \frac{w_{v}}{n} \int_{1}^{\frac{n}{2\lambda w_{v}}} w \cdot (\tau - 1) w^{-\tau} \mathrm{d}w\right) + O(w_{v})$	
$=\Theta\left(w_{v}\int_{1}^{\frac{n}{2\lambda w_{v}}}w^{-(\tau-1)}\mathrm{d}w\right)+O(w_{v})$	



• Consider vertex v with weight $w_v \qquad \int \Theta(n)$, if $w_v \geq \frac{n}{2\lambda}$	GIRG
• We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \begin{cases} 2\pi \\ 2\pi$	n independent vertices
• Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$	$ \begin{array}{ } \bullet x_{v} \sim \mathcal{U}([0,1]) \\ \bullet w_{v} \sim Par(\tau-1,1) \text{ for } \tau \in (2,3) \end{array} $
$\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$	$f_{w_v}(w) = (\tau - 1)w^{-\tau}$
$\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$	• u, v adjacent iff dist $(x_u, x_v) \le \lambda \frac{w_u \cdot w_v}{n}$
$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \qquad \text{w.l.o.g } x_v = \frac{1}{2}$	
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$	$n \rightarrow$
$=\Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}}\frac{w \cdot w_{v}}{n}f_{w_{u}}(w)dw + \Pr[w_{u} \geq \frac{n}{2\lambda w_{v}}]\right)\right) = \Theta\left(w_{u}\right)$	$ \sqrt{\left[\frac{1}{-(\tau-2)}w^{-(\tau-2)}\right]_{1}^{\frac{1}{2\lambda_{w_{v}}}}} + O(w_{v}) $
$=\Theta\left(n\int_{1}^{\frac{n}{2\lambda_{w_{v}}}}\frac{w\cdot w_{v}}{n}f_{w_{u}}(w)dw\right)+O(w_{v}) \qquad \qquad =\Theta\left(w_{v}\right)$	$\left[w^{-(\tau-2)}\right]^{1}_{\frac{n}{2\lambda w_{v}}} + O(w_{v})$
$=\Theta\left(n\frac{w_{v}}{n}\int_{1}^{\frac{n}{2\lambda w_{v}}}w\cdot(\tau-1)w^{-\tau}dw\right)+O(w_{v}) \qquad =\Theta\left(w_{v}\right)$	$\left(1-\left(\frac{n}{2\lambda w}\right)^{-(\tau-2)}\right)+O(w_{\nu})$
$=\Theta\left(w_{v}\int_{1}^{\frac{n}{2\lambda w_{v}}}w^{-(\tau-1)}\mathrm{d}w\right)+O(w_{v})$	



• Consider vertex v with weight $w_v \qquad \int \Theta(n)$, if $w_v \geq \frac{n}{2\lambda}$	GIRG
• We want to compute $\mathbb{E}[\deg(v) \mid w_v] = \begin{cases} 1 \\ 1 \\ 1 \\ 2 \\ 2$	n independent vertices
• Consider X_u for $u \in V \setminus \{v\}$ indicating whether $\{u, v\} \in E$	$ \mathbf{x}_{v} \sim \mathcal{U}([0, 1]) $ $ \mathbf{w}_{v} \sim \operatorname{Par}(\tau - 1, 1) \text{ for } \tau \in (2, 3) $
$\deg(v) = \sum_{u \in V \setminus \{v\}} X_u$	$f_{w_v}(w) = (au - 1)w^{- au}$
$\mathbb{E}[\deg(v) \mid w_v] = \sum_{u \in V \setminus \{v\}} \mathbb{E}[X_u \mid w_v]$	• u, v adjacent iff dist $(x_u, x_v) \leq \lambda \frac{w_u \cdot w_v}{w_u \cdot w_v}$
$= \Theta(n \Pr[\{u, v\} \in E \mid w_v]) \qquad \text{w.l.o.g } x_v = \frac{1}{2}$	
If $w_v < \frac{n}{2\lambda}$ = $\Theta(n \int_1^\infty \Pr[u \in N(v) w_u = w, w_v] f_{w_u}(w) dw)$	n
$= \Theta\left(n\left(\int_{1}^{\frac{n}{2\lambda w_{v}}} \frac{w \cdot w_{v}}{n} f_{w_{u}}(w) dw + \Pr[w_{u} \ge \frac{n}{2\lambda w_{v}}]\right)\right) = \Theta\left(\frac{1}{2\lambda w_{v}}\right)$	$W_{v}\left[\frac{1}{-(\tau-2)}W^{-(\tau-2)}\right]_{1}^{\frac{n}{2\lambda_{W_{v}}}}+O(W_{v})$
$=\Theta\left(n\int_{1}^{\frac{n}{2\lambda_{w_{v}}}}\frac{w\cdot w_{v}}{n}f_{w_{u}}(w)dw\right)+O(w_{v}) \qquad \qquad =\Theta\left(w_{v}\right)$	$W_{v}\left[W^{-(\tau-2)}\right]_{\frac{n}{2\lambda W_{v}}}^{1} + O(W_{v})$
$=\Theta\left(n\frac{w_{v}}{n}\int_{1}^{\frac{n}{2\lambda w_{v}}}w\cdot(\tau-1)w^{-\tau}dw\right)+O(w_{v})\qquad \qquad =\Theta\left(w_{v}\right)$	$v_{v}\left(1-\left(\frac{n}{2\lambda w_{v}}\right)^{-(\tau-2)}\right)+O(w_{v})$
$=\Theta\left(w_{v}\int_{1}^{2\lambda w_{v}}w^{-(\tau-1)}\mathrm{d}w\right)+O(w_{v})$	< 1 and O(1)

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Structural Properties

• Heterogeneity: deg(v) $\approx w_v$, $w_v \sim Par(\tau - 1, 1) \rightarrow power-law$ degree distribution \checkmark



Structural Properties

• Heterogeneity: deg(v) $\approx w_v$, $w_v \sim Par(\tau - 1, 1) \rightarrow power-law degree distribution \checkmark$ (also works with other weight distributions)



Structural Properties

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• Heterogeneity: deg(v) $\approx w_v$, $w_v \sim Par(\tau - 1, 1) \rightarrow power-law$ degree distribution \checkmark (also works with other weight distributions)

Locality (not seen here)

Maximilian Katzmann, Stefan Walzer - Probability & Computing



• Heterogeneity: $\deg(v) \approx w_v$, $w_v \sim \mathsf{Par}(\tau - 1, 1) \rightsquigarrow \mathsf{power-law}$ degree distribution \checkmark

Locality (not seen here)

Algorithmic Properties

"On the External Validity of Average-Case Analyses of Graph Algorithms", Bläsius, Fischbeck, ACM Trans. Algorithms 2023



Are GIRGs Realistic?

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Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)



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"On the External Validity of Average-Case Analyses of Graph Algorithms", Bläsius, Fischbeck, ACM Trans. Algorithms 2023

• Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)



- What we considered just now

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Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)



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Algorithmic Properties

"On the External Validity of Average-Case Analyses of Graph Algorithms", Bläsius, Fischbeck, ACM Trans. Algorithms 2023

- Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)
- Measure algorithmic properties on GIRGs and real graphs



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- Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)
- Measure algorithmic properties on GIRGs and real graphs
 - Bidirectional breadth-first-search




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- Setup: GIRGs with varying degrees of heterogeneity and locality (each dot is a graph)
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 - Bidirectional breadth-first-search
 - Diameter computation via BFS



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 - Diameter computation via BFS
 - Vertex cover kernel size





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 - Vertex cover kernel size
 - Louvain clustering algorithm





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 rather structural property
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Use GIRGs for average-case analysis!





Vertex Cover
Given undirected graph G = (V, E)





Vertex Cover

• Given undirected graph G = (V, E) (induced subgraph)

• Find a smallest $S \subseteq V$ such that $\overline{G[V \setminus S]}$ is edgeless





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Vertex Cover Approximation

Find a small vertex cover S' fast





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- Find a small vertex cover S' fast
- Approximation ratio: r = |S'|/|S|





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Vertex Cover Approximation

- Find a small vertex cover S' fast
- Approximation ratio: r = |S'|/|S|
- NP-hard to approximate with $r < \sqrt{2}$





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14

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- Believed to be NP-hard for $r < 2 \varepsilon$ for const. ε





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Practice





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Practice





Vertex Cover

- Given undirected graph G = (V, E) (induced subgraph)
- Find a smallest $S \subseteq V$ such that $\overline{G[V \setminus S]}$ is edgeless
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Practice

- Simple approximation algorithm repeatedly takes/deletes vertex of largest degree
- Close to optimal ratios on real graphs

"Vertex Cover on Complex Networks", Da Silva, Gimenez-Lugo, Da Silva, IJMPC 2013





(based on)

"Efficiently Approximating Vertex Cover on Scale-Free Networks with Underlying Hyperbolic Geometry", Bläsius, Friedrich, K., Algorithmica 2023



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Analsysis on GIRGs

(based on) "Efficiently Approximating Vertex Cover on Scale-Free Networks with Underlying Hyperbolic Geometry", Bläsius, Friedrich, K., Algorithmica 2023

Keep it simple

Consider vertices in order of decreasing degree in original graph





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Learn from the Model

Once high-degree vertices are taken/removed, remaining vertices have roughly equal weight/degree





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- Improve quality by solving small separated components exactly
 log log(n)





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- Two variants
 - Search and solve small components after each greedily taken vertex





Stefan Walzer – Probability & Computing

Analsysis on GIRGs

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This variant yields an upper bound on the

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Theorem: Let *G* be GIRG with *n* vertices and *m* edges. Then, an approximate vertex cover *S'* of *G* can be computed in time $O(m \log(n))$ such that the approximation ratio is (1 + o(1)) asymptotically almost surely.

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Proof Approximation Ratio

Differentiate greedily taken vertices S'_g from ones in exactly solved components S'_e





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• $|S| = \Omega(n)$ with prob 1 - o(1)

"Greed is Good for Deterministic Scale-Free Networks", Chauhan et al. FSTTCS 2016

Remains to show: $|S'_g| = o(n)$







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• Consider random variable $X_v = \mathbb{1}_{\{w_v \ge t\}}$





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Since there is a $g(n) \in o(n) \cap \Omega(\log(n))$ with $g(n) \ge \mathbb{E}[N_{w \ge t}]$, Chernoff gives concentration

GIRG *n* independent vertices $w_v \sim Par(\tau - 1, 1)$ for $\tau \in (2, 3)$





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Discretize ground space into cells such that edges cannot span empty cells



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- Regard chains of non-empty cells as one component





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- Count all vertices that are in chains containing $> \log \log(n)$ vertices (also potentially counting small components)

When does a chain contain too many vertices?



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 - Total number of cells in long chains does not change much ($\leq 2k + 1$) when one cell moves from empty to non-empty (or vice versa) > k





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- Use Poissonization to get rid of dependencies $t = \frac{W_V}{V}$





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 \rightsquigarrow w.h.p., o(n) vertices in large components \checkmark



Conclusion



Method of Bounded Differences

Concentration for function of independent random variables

Conclusion



- Concentration for function of independent random variables
- Bounded differences ("Lipschitz") condition
 - What is the worst that can happen when changing one input?





Conclusion

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What is the worst that can happen when changing one input?

Conclusion

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Method of Bounded Differences

Method of Typical Bounded Differences

Bounded differences ("Lipschitz") condition

- Define typical event, distinguish worst changes depending on whether event occurred
- Use mitigators to weaken impact of general worst changes
- Pay with probability that typical event does not occur, multiplied with inverse mitigators

Concentration for function of independent random variables





Bounded differences ("Lipschitz") condition What is the worst that can happen when changing one input?

Chernoff-like bound, weakened by sum of squared worst changes

Concentration for function of independent random variables

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Geometric Inhomogeneous Random Graphs

- Pretty realistic graph model (heterogeneity, locality)
- Not too hard to analyze
- Used for average-case analysis (e.g. vertex cover approximation)

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(not discussed in lecture)





